

# 2

## STOCHASTIC INTEREST RATE MODELS

## 2.1. CONTINUOUS TIME FINANCE RECAP

Note: Please see Hull (2018), Chap.14.

- **Stochastic process** – any variable whose value changes over time in an uncertain way => different random trajectories for the variable.
- **Discrete vs continuous time stochastic processes:**
  - Discrete – the variable value can change only at certain fixed points in time
  - Continuous – changes can take place at any time
- **Continuous vs discrete variables:**
  - Discrete – only certain values are possible
  - Continuous – can take any value within a certain range
- Continuous-variable, continuous-time – variables can assume any value and changes can occur at any time.

# STOCHASTIC PROCESSES

- Continuous-variable, continuous-time stochastic processes are key to understanding the pricing of options and other derivatives.
- However, in practice, most asset prices do not follow continuous-variable, continuous-time stochastic processes.
- For instance, stock prices are restricted to discrete values (e.g. multiples of a cent) and changes can be observed only when the markets are open.
- Nonetheless, continuous-variable, continuous-time stochastic processes are useful for many valuation purposes.
- **Markov Stochastic Process** – stochastic process where **only the current value of a variable is relevant for predicting the future** => **all past information is irrelevant**, as it is already incorporated into today's stock price (**market efficiency**).



- The probability distribution at any particular future time is independent from the path followed by the variable in the past.
- If market efficiency didn't hold, market participants could make above-average returns by interpreting the past behavior of asset prices.

# STOCHASTIC PROCESSES

- Assuming a Markov process  $X(t)$ , the 1-year change  $\sim \mathcal{N}(0,1)$ .
- 2-year change =  $\mathcal{N}(0,1) + \mathcal{N}(0,1) = \mathcal{N}(0,2)$ , as both distributions are independent - given that this is a Markov process, the 2<sup>nd</sup> distribution does not depend on the 1<sup>st</sup>.

$\Delta t$  (very small period of time) change  $\sim \mathcal{N}(0, \Delta t)$

The expected value of any future outcome is equal to the current value (Martingale):  $z=25 \Rightarrow$  1 year after,  $z \sim \mathcal{N}(25,1)$ ; 5 years after,  $z \sim \mathcal{N}(25,5)$

# WIENER PROCESS

Wiener Process – a particular type of Markov process with:

- a mean change = 0
- a variance rate (per year) = 1



A stochastic process  $z$  follows a Wiener process (or the continuous random walk) if it has the following properties:

Property 1. *The change  $\Delta z$  during a small period of time  $\Delta t$  is*

$$\Delta z = \epsilon \sqrt{\Delta t} \quad (14.1)$$

*where  $\epsilon$  has a standard normal distribution  $\phi(0, 1)$ .*

Property 2. *The values of  $\Delta z$  for any two different short intervals of time,  $\Delta t$ , are independent.*

It follows from the first property that  $\Delta z$  itself has a normal distribution with

mean of  $\Delta z = 0$

standard deviation of  $\Delta z = \sqrt{\Delta t}$

variance of  $\Delta z = \Delta t$

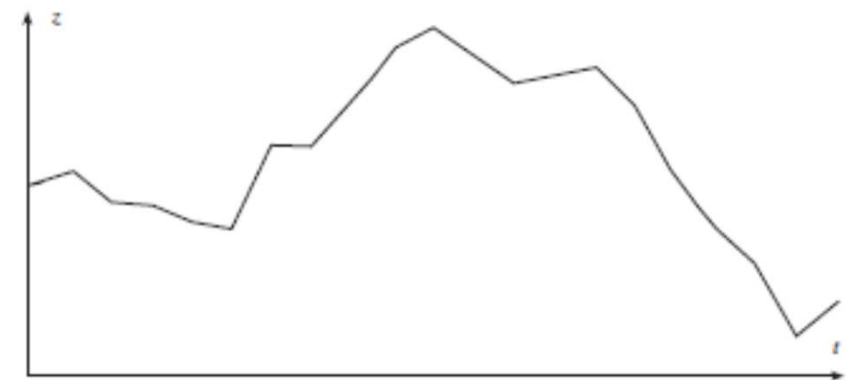
The second property implies that  $z$  follows a Markov process

Hull, John (2018), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 10<sup>th</sup> Edition

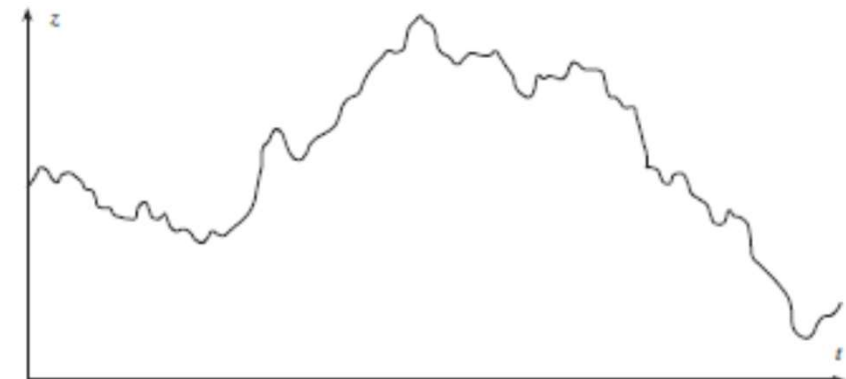
# WIENER PROCESS

Wiener processes for different magnitudes of change in time:

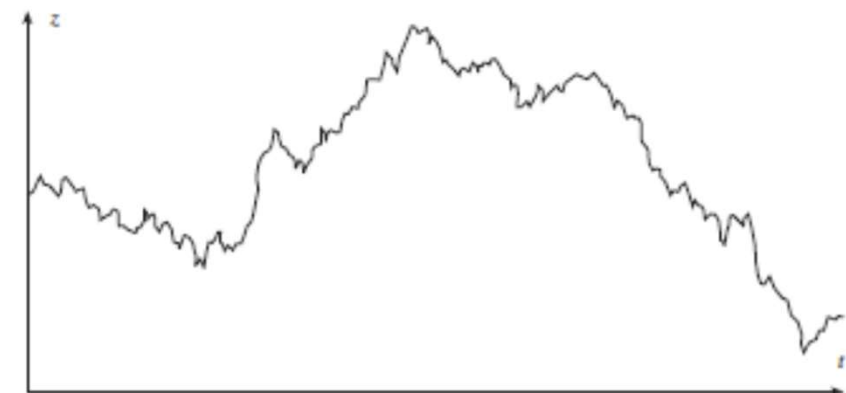
When  $\Delta t \rightarrow 0$ , the path becomes much more irregular, as the size of the movement in the variable in time  $\Delta t$  is proportional to the  $\sqrt{\Delta t}$ . When  $\Delta t$  is small,  $\sqrt{\Delta t}$  is much larger than  $\Delta t \Rightarrow$  **the changes in  $z$  will be much larger than  $\Delta t$ , as  $\Delta z = \epsilon \sqrt{\Delta t}$**



Relatively large value of  $\Delta t$



Smaller value of  $\Delta t$



The true process obtained as  $\Delta t \rightarrow 0$

Source: Hull, John (2018), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 10<sup>th</sup> Edition

# GENERALIZED WIENER PROCESS

- Instead of a drift = 0 and a variance rate = 1 as in the Wiener process ( $dz$ ), we may have a stochastic process **where the drift can assume any value  $a$  and the variance rate can be  $b^2$**   
=> **Generalized Wiener Process.**

$$\Delta x = a \Delta t + b\epsilon\sqrt{\Delta t} \quad \text{where } a \text{ and } b \text{ are constants.}$$

- In the Wiener Process,  $a=0$  and  $b=1$ :  $\Delta z = \epsilon\sqrt{\Delta t}$

- For very small time changes  $\Delta t$ :

$$\Delta x = a \Delta t + b\epsilon\sqrt{\Delta t}$$

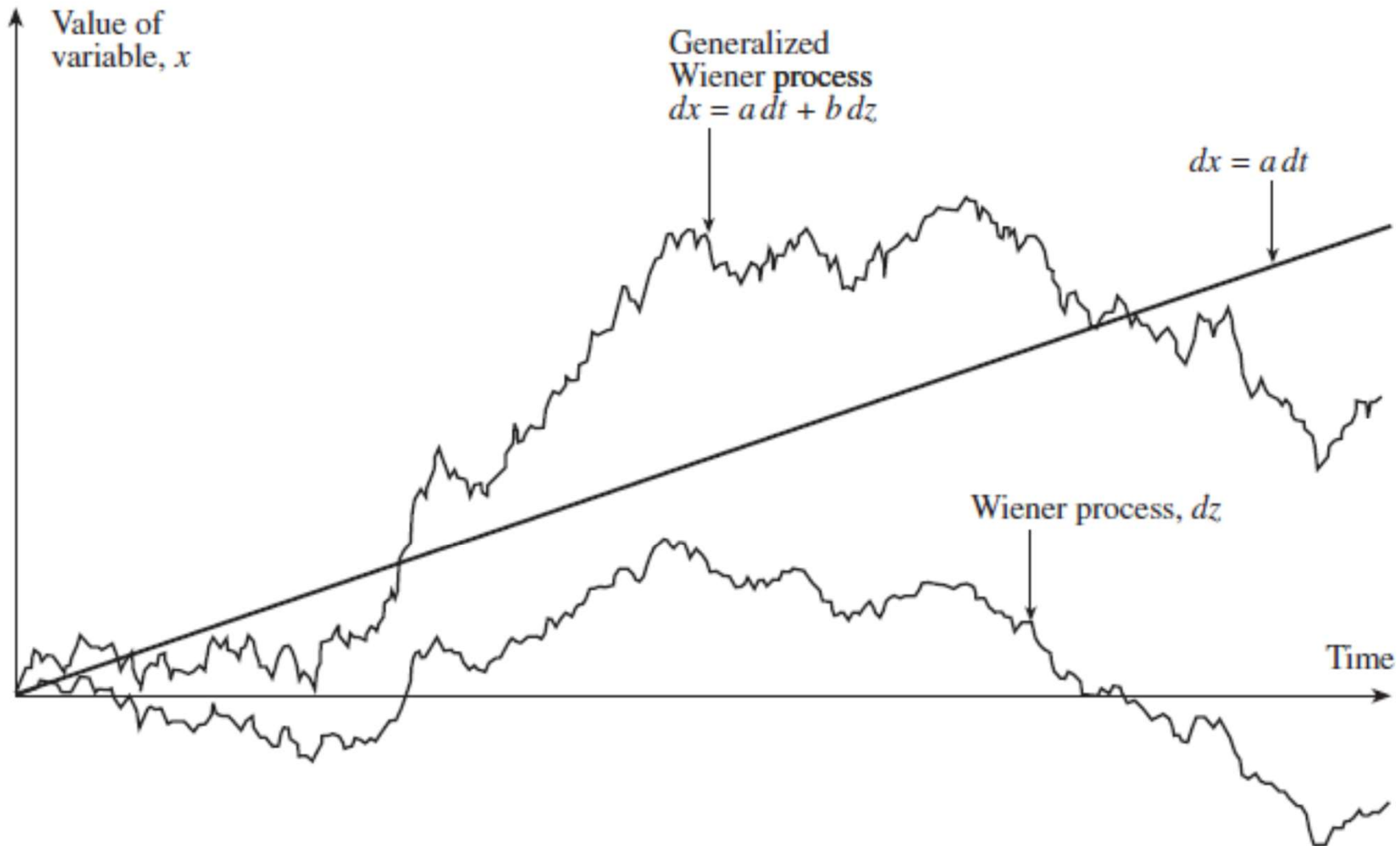
  
 $\Delta x \sim N$ , with

mean of  $\Delta x = a \Delta t$   
standard deviation of  $\Delta x = b\sqrt{\Delta t}$   
variance of  $\Delta x = b^2 \Delta t$

- The average increases in  $x$  are proportional to time (if there is no drift, i.e.  $a = 0$ , the mean of  $x$  doesn't change, i.e.  $\Delta x = 0$ ).

# GENERALIZED WIENER PROCESS

Figure 14.2 Generalized Wiener process with  $a = 0.3$  and  $b = 1.5$ .



Source: Hull, John (2018), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 10<sup>th</sup> Edition

# ITÔ PROCESS

- Definition: Generalized Wiener process with average and standard-deviation as a function of the underlying variable and time (instead of being only a function of time):

$$dx = a(x, t) dt + b(x, t) dz$$


$$\Delta x = a \Delta t + b \epsilon \sqrt{\Delta t}$$

- For small time intervals, we may assume that the average and the standard-deviation don't change (we're assuming that the drift and the variance rate don't change between  $t$  and  $t+\Delta t$ ):  $\Delta x = a(x, t)\Delta t + b(x, t)\epsilon\sqrt{\Delta t}$
- This is still a Markov process, as  $a$  and  $b$  only depend on the current value of  $x$ , not on previous values.
- It may be tempting to assume that a stock price follows a Generalized Wiener process (constant drift and variance).
- However, this assumption is not valid, having in mind that investors require or expect a given level of returns (as a % variation) regardless the price level, i.e. for higher prices, expected changes will also be higher.
- One can acknowledge this fact by replacing the assumption of constant expected drift by the assumption of constant expected returns (i.e. constant expected drift divided by the stock price  $\Leftrightarrow$  variable drift along time).

## ITÔ PROCESS



- Expected drift rate in  $S = a(x, t) = \mu S$

(being  $S$  the stock price at time  $t$  and  $\mu$  the expected rate of return of the stock, expressed in decimal form, for a time unit = 1).

- For a short interval of time  $\Delta t$ , the expected increase in  $S$  is  $\mu S \Delta t$ , i.e. the expected rate of return on the stock, times the stock price, times the time interval:

$$\Delta S = \mu S \Delta t$$

$$\Delta x = a(x, t) \Delta t + b(x, t) \epsilon \sqrt{\Delta t}$$

- If  $\Delta t \rightarrow 0 \Rightarrow dS = \mu S dt \Leftrightarrow \frac{dS}{S} = \mu dt$

- This corresponds to the **price of an asset following a continuously compounding process** (under no uncertainty, being  $\mu$  = risk-free rate in a risk-neutral world)



$$S_T = S_0 e^{\mu T}$$

# GEOMETRIC BROWNIAN MOTION

- Given that in practice there is uncertainty, a reasonable assumption is that the variability of the percentage return ( $\sigma$ ) in a short period of time  $\Delta t$  is the same regardless the stock price.



- An investor is as uncertain about his return when the stock price is high or low.
- Accordingly, the standard deviation of the change in a short period of time must also be proportional to the stock price, as the standard deviation for the percentage change is constant – **Geometric Brownian Motion**:

$$dS = \mu S dt + \sigma S dz \Leftrightarrow$$

$$\frac{dS}{S} = \mu dt + \sigma dz$$

## GEOMETRIC BROWNIAN MOTION

- Example:

Consider a stock that pays no dividends, has a volatility of 30% per annum, and provides an expected return of 15% per annum with continuous compounding. In this case,  $\mu = 0.15$  and  $\sigma = 0.30$ . The process for the stock price is

$$\frac{dS}{S} = 0.15 dt + 0.30 dz$$

If  $S$  is the stock price at a particular time and  $\Delta S$  is the increase in the stock price in the next small interval of time, the discrete approximation to the process is

$$\frac{\Delta S}{S} = 0.15\Delta t + 0.30\epsilon\sqrt{\Delta t}$$

where  $\epsilon$  has a standard normal distribution. Consider a time interval of 1 week, or 0.0192 year, so that  $\Delta t = 0.0192$ . Then the approximation gives

$$\frac{\Delta S}{S} = 0.15 \times 0.0192 + 0.30 \times \sqrt{0.0192} \epsilon$$

or

$$\Delta S = 0.00288S + 0.0416S\epsilon$$

Source: Hull, John (2018), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 10<sup>th</sup> Edition

# GEOMETRIC BROWNIAN MOTION

- Monte Carlo simulation:

A path for the stock price over 10 weeks can be simulated by sampling repeatedly for  $\epsilon$  from  $\phi(0, 1)$  and substituting into equation (14.10). The expression `=RAND()` in Excel produces a random sample between 0 and 1. The inverse cumulative normal distribution is `NORMSINV`. The instruction to produce a random sample from a standard normal distribution in Excel is therefore `=NORMSINV(RAND())`. Table 14.1 shows one path for a stock price that was sampled in this way. The initial stock price is assumed to be \$100. For the first period,  $\epsilon$  is sampled as 0.52. From equation (14.10), the change during the first time period is

$$\frac{\Delta S}{S} = 0.15 \times 0.0192 + 0.30 \times \sqrt{0.0192} \epsilon$$

$$\Delta S = 0.00288S + 0.0416S\epsilon$$

**Table 14.1** Simulation of stock price when  $\mu = 0.15$  and  $\sigma = 0.30$  during 1-week periods.

| <i>Stock price<br/>at start of period</i> | <i>Random sample<br/>for <math>\epsilon</math></i> | <i>Change in stock price<br/>during period</i> |
|-------------------------------------------|----------------------------------------------------|------------------------------------------------|
| 100.00                                    | 0.52                                               | 2.45                                           |
| 102.45                                    | 1.44                                               | 6.43                                           |
| 108.88                                    | -0.86                                              | -3.58                                          |
| 105.30                                    | 1.46                                               | 6.70                                           |
| 112.00                                    | -0.69                                              | -2.89                                          |
| 109.11                                    | -0.74                                              | -3.04                                          |
| 106.06                                    | 0.21                                               | 1.23                                           |
| 107.30                                    | -1.10                                              | -4.60                                          |
| 102.69                                    | 0.73                                               | 3.41                                           |
| 106.11                                    | 1.16                                               | 5.43                                           |
| 111.54                                    | 2.56                                               | 12.20                                          |

Source: Hull, John (2018), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 10<sup>th</sup> Edition

## ITÔ'S LEMMA

- An option price ( $G$ ) is a function of the underlying asset's price and time.



- It is key to understand the **behavior of functions of stochastic variables**.
- An important result was discovered by K. Itô in 1951 and is known as **Itô's lemma**.

<sup>6</sup> See K. Itô, "On Stochastic Differential Equations," *Memoirs of the American Mathematical Society*, 4 (1951): 1–51.

- Assuming that a **variable  $x$  follows an Itô process**:

$$dx = a(x, t) dt + b(x, t) dz$$

where  $dz$  is a Wiener process and  $a$  and  $b$  are functions of  $x$  and  $t$ . The variable  $x$  has a drift rate of  $a$  and a variance rate of  $b^2$ . Itô's lemma shows that a function  $G$  of  $x$  and  $t$  follows the process

$$dG = \left( \frac{\partial G}{\partial x} a + \frac{\partial G}{\partial t} + \frac{1}{2} \frac{\partial^2 G}{\partial x^2} b^2 \right) dt + \frac{\partial G}{\partial x} b dz$$

# ITÔ'S LEMMA

- Thus,  $G$  also follows an Itô process with a drift rate =  $\frac{\partial G}{\partial x}a + \frac{\partial G}{\partial t} + \frac{1}{2}\frac{\partial^2 G}{\partial x^2}b^2$  and a variance rate of  $\left(\frac{\partial G}{\partial x}\right)^2 b^2$ 

$$dx = a(x, t)dt + b(x, t)dz$$

- Assuming that the stock price follows a Geometric Brownian Motion, with constant  $\mu$  and  $\sigma$ .  $dS = \mu Sdt + \sigma Sdz$

- Applying the Ito's Lemma to the previous equation:

$$dG = \left( \frac{\partial G}{\partial S} \mu S + \frac{\partial G}{\partial t} + \frac{1}{2} \frac{\partial^2 G}{\partial S^2} \sigma^2 S^2 \right) dt + \frac{\partial G}{\partial S} \sigma S dz$$

$$dG = \left( \frac{\partial G}{\partial x} a + \frac{\partial G}{\partial t} + \frac{1}{2} \frac{\partial^2 G}{\partial x^2} b^2 \right) dt + \frac{\partial G}{\partial x} b dz$$




- Therefore, **both  $S$  and  $G$  are affected by the same volatility source –  $dz$ .**
- In the Black-Scholes option pricing formula,  $G$  (the option price) is determined by the instantaneous volatility of the returns of the underlying asset price.

# APPLICATION TO FORWARD CONTRACTS

- Forward:  $F_0 = S_0 e^{rT}$

- Forward at  $t$ :  $F = S e^{r(T-t)}$

- Ito's Lemma => 
$$dG = \left( \frac{\partial G}{\partial S} \mu S + \frac{\partial G}{\partial t} + \frac{1}{2} \frac{\partial^2 G}{\partial S^2} \sigma^2 S^2 \right) dt + \frac{\partial G}{\partial S} \sigma S dz$$



- The stochastic process of  $F$  can be defined calculating the derivatives of  $F$  in order to  $S$  and  $t$  (i.e.  $F$  now corresponds to  $G$ ):

$$\frac{\partial F}{\partial S} = e^{r(T-t)}, \quad \frac{\partial^2 F}{\partial S^2} = 0, \quad \frac{\partial F}{\partial t} = -r S e^{r(T-t)}$$



$$dF = [e^{r(T-t)} \mu S - r S e^{r(T-t)}] dt + e^{r(T-t)} \sigma S dz \Rightarrow dF = (\mu - r) F dt + \sigma F dz$$

Substituting  $F$  for  $S e^{r(T-t)}$

- Like  $S$ ,  $F$  follows a GMB, with the same volatility and a trend of  $\mu - r$  (instead of  $\mu$ ).

# PROBABILITY DISTRIBUTION

- From the stochastic process of the rate of returns,

$$\frac{dS}{S} = \mu dt + \sigma dz$$

- Its distribution gets

$$\frac{\Delta S}{S} \sim \phi(\mu \Delta t, \sigma^2 \Delta t)$$

- Assuming  $G = \ln S$ , since  $\frac{\partial G}{\partial S} = \frac{1}{S}$ ,  $\frac{\partial^2 G}{\partial S^2} = -\frac{1}{S^2}$ ,  $\frac{\partial G}{\partial t} = 0$ , it follows from the Itô's lemma that


$$dG = \left( \mu - \frac{\sigma^2}{2} \right) dt + \sigma dz$$


$$dG = \left( \frac{\partial G}{\partial S} \mu S + \frac{\partial G}{\partial t} + \frac{1}{2} \frac{\partial^2 G}{\partial S^2} \sigma^2 S^2 \right) dt + \frac{\partial G}{\partial S} \sigma S dz$$

# PROBABILITY DISTRIBUTION

Since  $\mu$  and  $\sigma$  are constant, this equation indicates that  $G = \ln S$  follows a generalized Wiener process. It has constant drift rate  $\mu - \sigma^2/2$  and constant variance rate  $\sigma^2$ . The change in  $\ln S$  between time 0 and some future time  $T$  is therefore normally distributed, with mean  $(\mu - \sigma^2/2)T$  and variance  $\sigma^2 T$ . This means that

$$\ln S_T - \ln S_0 \sim \phi \left[ \left( \mu - \frac{\sigma^2}{2} \right) T, \sigma^2 T \right]$$

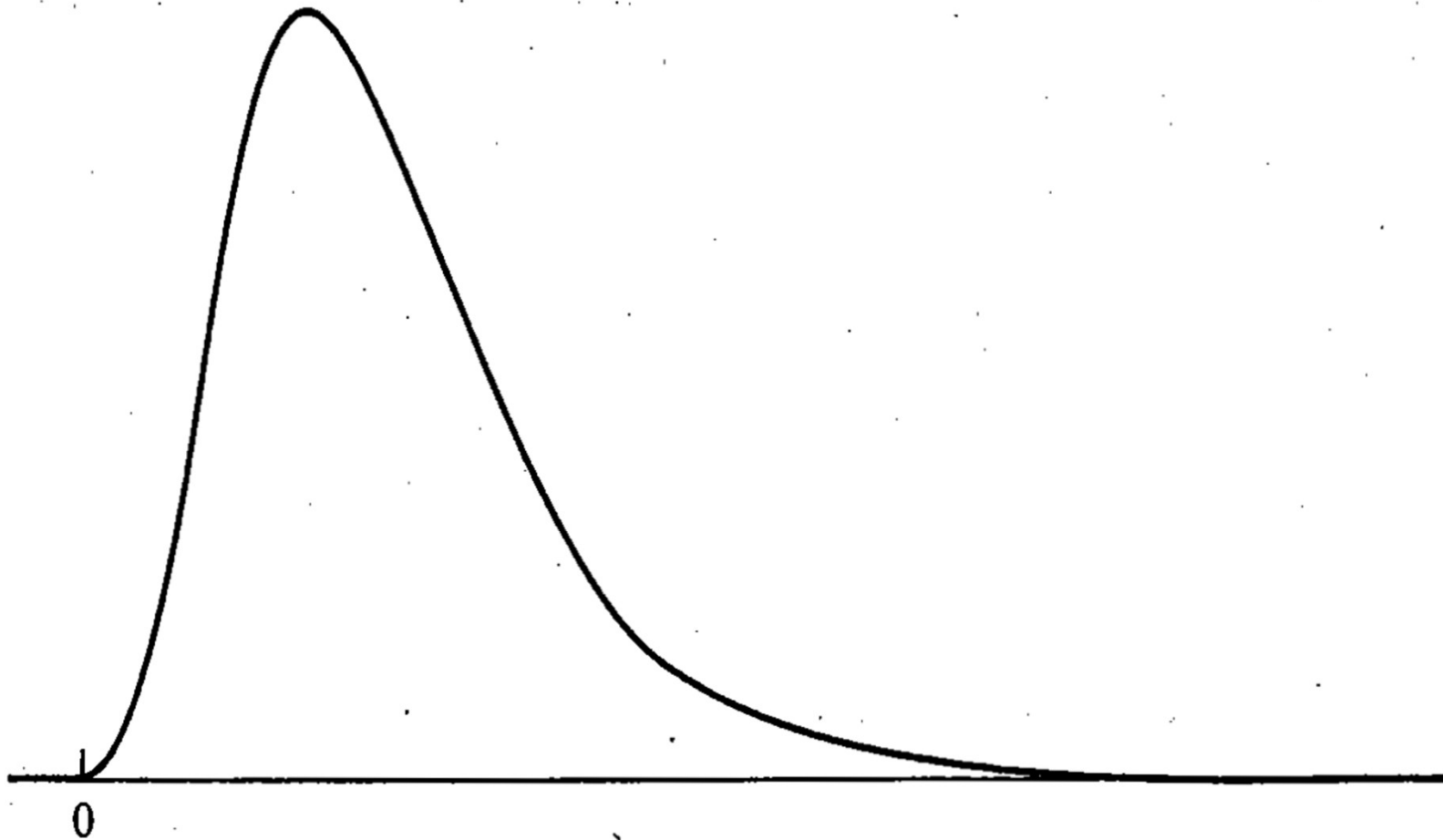
or

$$\ln S_T \sim \phi \left[ \ln S_0 + \left( \mu - \frac{\sigma^2}{2} \right) T, \sigma^2 T \right]$$

- 
- **$\ln S_T$  is normally distributed (and  $S_T$  has a log normal distribution),** with a standard deviation  $\sigma\sqrt{T}$  (proportional to the square root of time).
- 
- **The growth rate of the asset price is normally distributed => the asset price is lognormally distributed.**

# PROBABILITY DISTRIBUTION

**Figure 13.1** Lognormal distribution.



Source: Hull, John (2009), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 7<sup>th</sup> Edition

## **2.2. SHORT RATE MODELS**

- 2.2.1. Interest Rate Trees
- 2.2.2. Continuous-time Single-factor models
- 2.2.3. Continuous-time Multi-Factor models

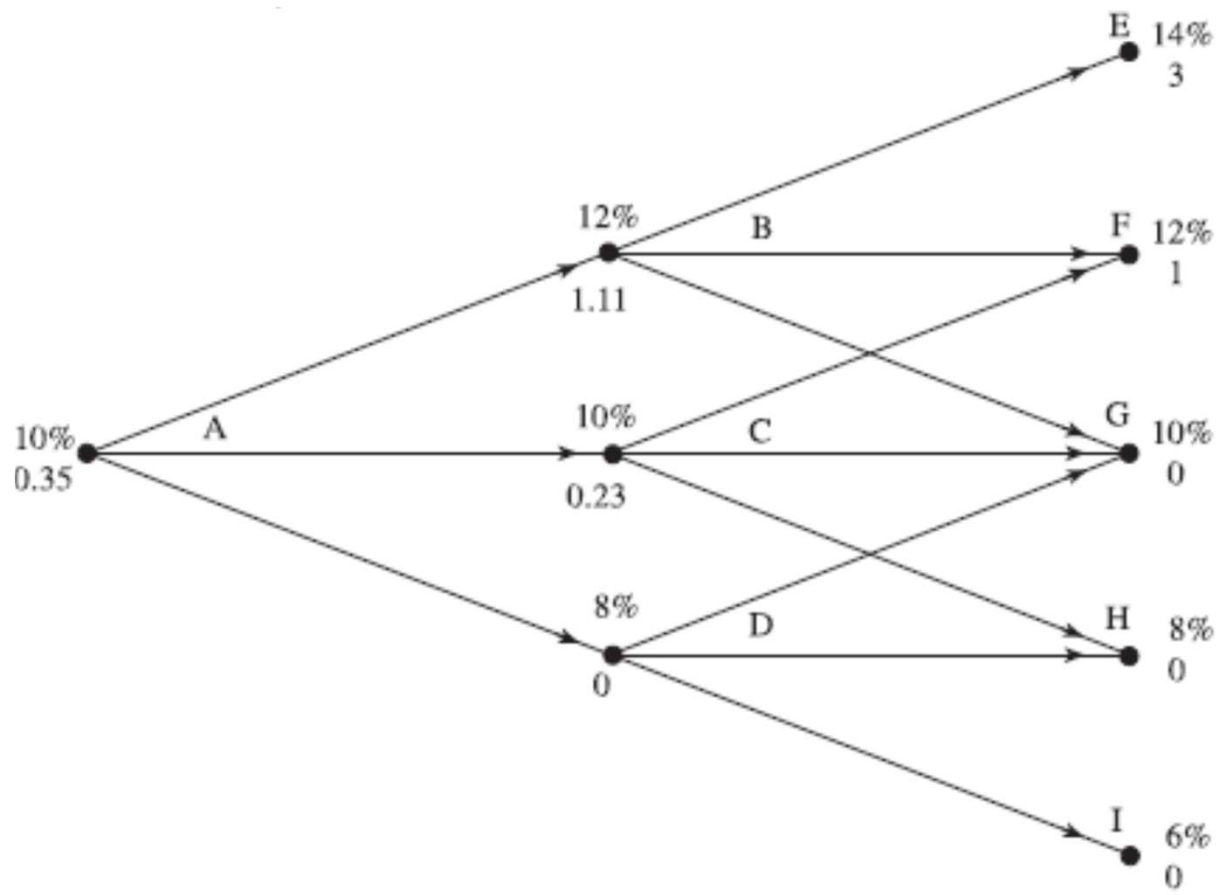
## 2.2.1. INTEREST RATE TREES

- Focus: How to model the TSIR by specifying the behavior of the short-term interest rate?
- Bond and interest rate derivative prices depend on the behavior of the risk-free short-term interest rate (or instantaneous short rate).



- The variable to be modeled by trees will be the instantaneous short rate.
- Why are trees used? a tree is a discrete-time representation of the stochastic process.
- Binomial trees are often used, even though trinomial trees are recommended to value interest rate derivatives.
- At the final nodes, the value of the derivative equals its pay-off.
- At previous nodes, the value of the derivative is calculated through a rollback procedure, calculating the expected value of the derivative according to the probabilities attached to the different scenarios and discounting this expected value using the interest rate at that node.

**Figure 32.4** Example of the use of trinomial interest rate trees. Upper number at each node is rate; lower number is value of instrument.



**Assumptions:**

Probabilities of up, middle and down are 0.25, 0.5 and 0.25, respectively.

**Derivative value at Node B:**

$$[0.25 \times 3 + 0.5 \times 1 + 0.25 \times 0]e^{-0.12 \times 1} = 1.11$$

**Derivative value at Node C:**

$$(0.25 \times 1 + 0.5 \times 0 + 0.25 \times 0)e^{-0.1 \times 1} = 0.23$$

**Derivative value at Node D = 0**

**Derivative value at Node A:**

$$(0.25 \times 1.11 + 0.5 \times 0.23 + 0.25 \times 0)e^{-0.1 \times 1} = 0.35$$

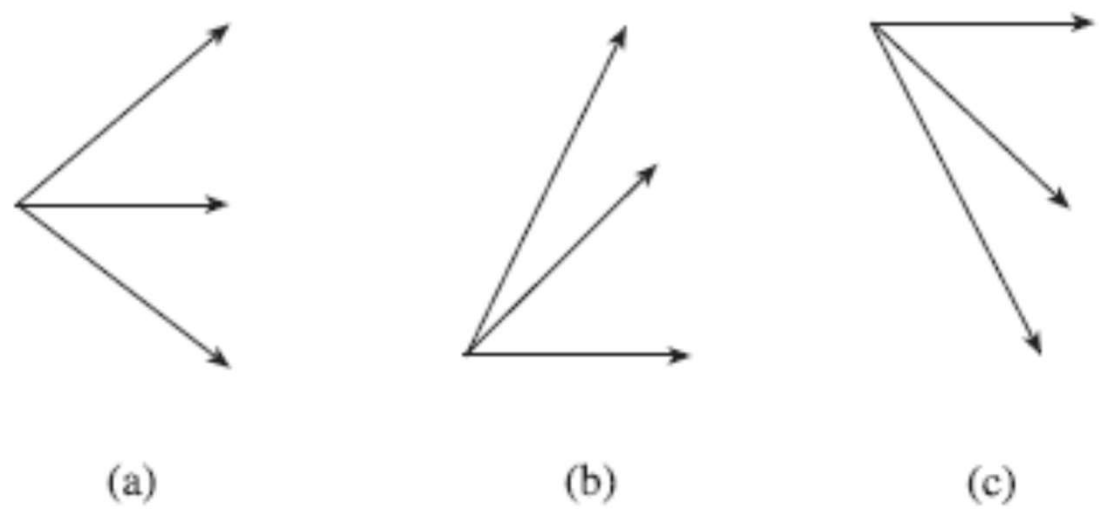
The tree is used to value a derivative that provides a payoff at the end of the second time step of

$$\max[100(R - 0.11), 0]$$

Source: Hull, John (2018), "Options, Futures and Other Derivatives", Pearson Prentice Hall, 10<sup>th</sup> Edition

➤ **Non-standard branching:**

**Figure 32.5** Alternative branching methods in a trinomial tree.



Source: Hull, John (2018), “Options, Futures and Other Derivatives”, Pearson Prentice Hall, 10<sup>th</sup> Edition

- (b) and (c) are useful to represent mean-reverting interest rates when interest rates are either very low (and are not supposed to move even lower) or very high (and are not supposed to move even higher), respectively.

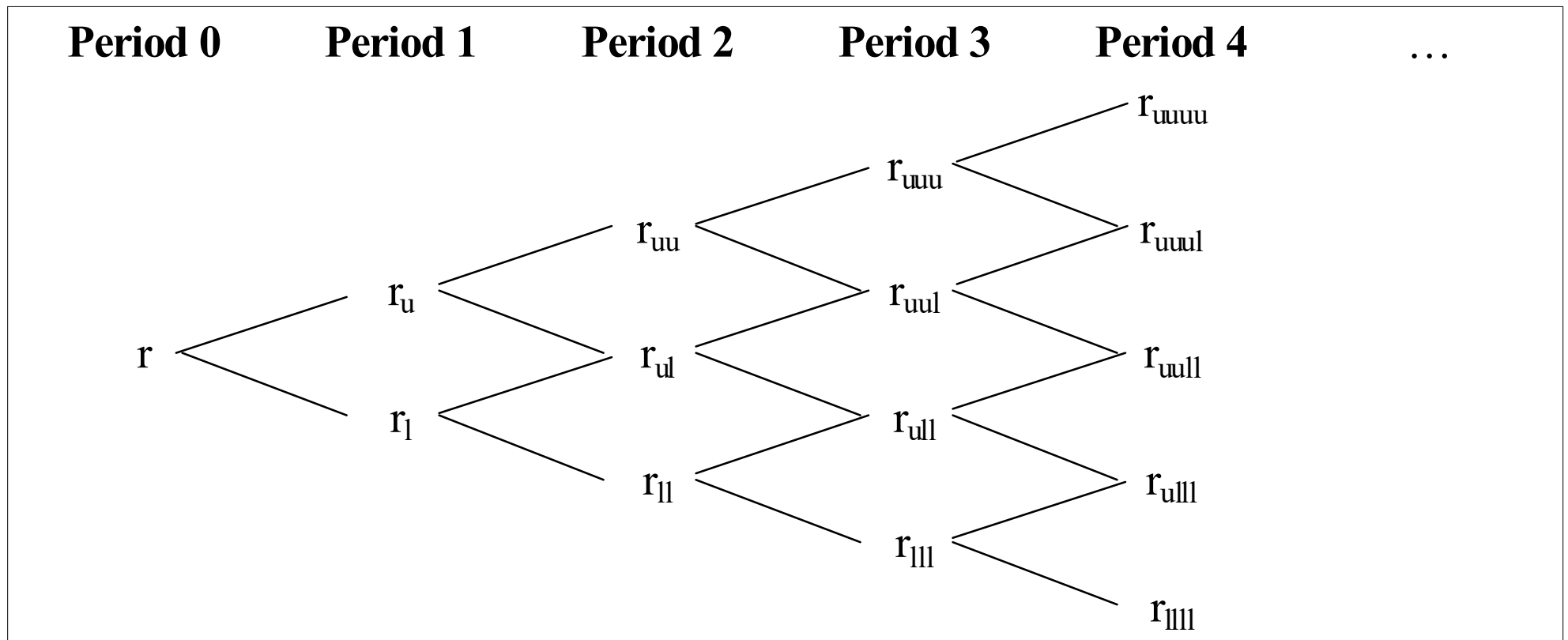
## ➤ General binomial model

- Given the current level of short-term rate  $r$ , the next-period short rate can take only two possible values: an upper value  $r_u$  and a lower value  $r_l$ , with equal probability 0.5.
- In the following period, the short-term interest rate can take four possible values:  $r_{uu}, r_{ul}, r_{lu}, r_{ll}$ .
- More generally, in period  $n$ , the short-term interest rate can take on  $2^n$  values => very time-consuming and computationally inefficient.

## ➤ Recombining trees

- Means that an upward-downward sequence leads to the same result as a downward-upward sequence (regardless being binomial or trinomial trees)
- For example,  $r_{ul} = r_{lu}$  => only  $(n+1)$  different values at period  $n$ .

# INTEREST RATE TREE - RECOMBINING



# INTEREST RATE TREE – ANALYTICAL

- We may write down the binomial process as:

$$\Delta r_t \equiv r_{t+1} - r_t = \sigma \varepsilon_t$$



Itô process

$$dx = a(x, t) dt + b(x, t) dz$$

$$\Delta r_t \equiv r_{t+\Delta t} - r_t = \mu(t, \Delta t, r_t) + \sigma(t, \Delta t, r_t) \varepsilon_t$$

- Specific case – assuming that the drift and the variance are proportional to the time increment:

$$\Delta r_t \equiv r_{t+\Delta t} - r_t = \mu \Delta t + \sigma \sqrt{\Delta t} \varepsilon_t$$

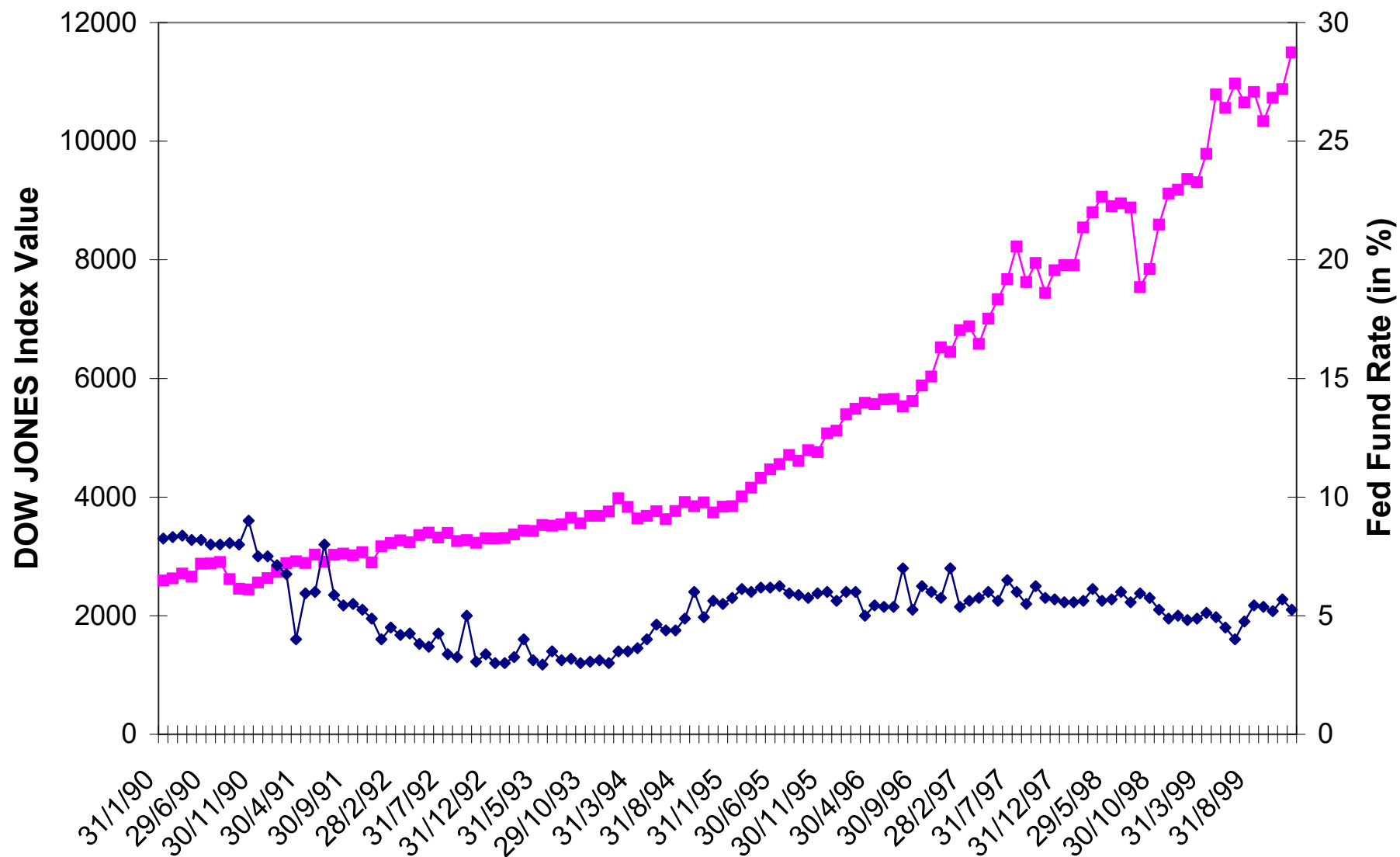
- Continuous-time limit (Merton (1973)):

$$dr_t \equiv r_{t+dt} - r_t = \mu dt + \sigma dW_t$$

## 2.2.2. CT SINGLE FACTOR MODELS

- In the initial term structure models, the short rate is the only driver of the yield curve, being assumed as a continuous and stochastic or random variable.
- Therefore, a single-factor continuous-time model specifies the dynamics of the short-term rate:  $dr_t = \mu(t, r_t)dt + \sigma(t, r_t)dW_t$
- The term  $W$  denotes a Brownian motion - process with independent normally distributed increments:
$$dW_t = \varepsilon_t \sqrt{dt}$$
  - $dW$  represents the instantaneous change;
  - It is stochastic (uncertain);
  - It is a stochastic variable with a normal distribution with zero mean and variance  $dt$ ;
- A good model is a model that is consistent with reality =>
  - Tractable
  - Parsimonious
- Stylized facts about the dynamics of the term structure:
  - Fact 1: (nominal) interest rates are (usually) positive
  - Fact 2: interest rates are mean-reverting
  - Fact 3: interest rates with different maturities are imperfectly correlated
  - Fact 4: the volatility of interest rates evolves (randomly) in time

## EMPIRICAL FACTS 1, 2 AND 4



**Note: Interest Rate in blue**

## EMPIRICAL FACT 3

|     | 1M    | 3M    | 6M    | 1Y    | 2Y    | 3Y    | 4Y    | 5Y    | 7Y    | 10Y |
|-----|-------|-------|-------|-------|-------|-------|-------|-------|-------|-----|
| 1M  | 1     |       |       |       |       |       |       |       |       |     |
| 3M  | 0.999 | 1     |       |       |       |       |       |       |       |     |
| 6M  | 0.908 | 0.914 | 1     |       |       |       |       |       |       |     |
| 1Y  | 0.546 | 0.539 | 0.672 | 1     |       |       |       |       |       |     |
| 2Y  | 0.235 | 0.224 | 0.31  | 0.88  | 1     |       |       |       |       |     |
| 3Y  | 0.246 | 0.239 | 0.384 | 0.808 | 0.929 | 1     |       |       |       |     |
| 4Y  | 0.209 | 0.202 | 0.337 | 0.742 | 0.881 | 0.981 | 1     |       |       |     |
| 5Y  | 0.163 | 0.154 | 0.255 | 0.7   | 0.859 | 0.936 | 0.981 | 1     |       |     |
| 7Y  | 0.107 | 0.097 | 0.182 | 0.617 | 0.792 | 0.867 | 0.927 | 0.97  | 1     |     |
| 10Y | 0.073 | 0.063 | 0.134 | 0.549 | 0.735 | 0.811 | 0.871 | 0.917 | 0.966 | 1   |

# EQUILIBRIUM VS NO-ARBITRAGE MODELS OF THE SHORT RATE

## Equilibrium models:

- (i) Start with assumptions about economic variables and derive a process for the short rate ( $r$ ).
- (ii) Accordingly, the initial yield curve is given by an analytical formula as a function of the short-term rate and the model parameters, assuming that the economy is in equilibrium.
- (iii) The process for  $r$  in a one-factor equilibrium model involves only one source of uncertainty - the (instantaneous) short-term rate itself => endogenous models.
- (iv) A one-factor model implies that all rates move in the same direction over any short time interval, but not that they all move by the same amount => the shape of the zero curve can change with the passage of time.
- (v) The process for the short rate is usually assumed to be stationary, in the sense that the parameters of the process are not functions of time.
- (vi) If the instantaneous short rate follows a Markov process, all rates can be calculated at all times as a function of the short rate.
- (vii) A shortcoming of these models is that they do not automatically fit today's term structure of interest rates.

# EQUILIBRIUM VS NO-ARBITRAGE MODELS OF THE SHORT RATE

## No-arbitrage models:

- (i) A no-arbitrage model is designed to be exactly consistent with today's TSIR.
  - (ii) Essential difference between an equilibrium and a no-arbitrage model – in an equilibrium model, today's TSIR is an output, while in a no-arbitrage model it is an input.
  - (iii) the drift is, in general, dependent on time, as the shape of the initial spot curve governs the average path taken by the short rate in the future – positively sloped zero curve => positive drift for the short rate.
- 
- (iv) Some equilibrium models can be transformed into no-arbitrage models by including a function of time in the drift of the short rate.

# EQUILIBRIUM ONE-FACTOR MODELS OF THE SHORT RATE

- Interest rate dynamics:

$$dr = m(r) dt + s(r) dz \longrightarrow \text{The drift is not a function of time, but of the interest rate itself.}$$

- Main type of models:

$$m(r) = \mu r; \quad s(r) = \sigma r \quad (\text{Rendleman and Bartter model})$$

$$m(r) = a(b - r); \quad s(r) = \sigma \quad (\text{Vasicek model})$$

$$m(r) = a(b - r); \quad s(r) = \sigma \sqrt{r} \quad (\text{Cox, Ingersoll, and Ross model})$$

# EQUILIBRIUM ONE-FACTOR MODELS OF THE SHORT RATE

## 1. Rendleman and Bartter

- The short-term interest rate follows a GMB:  $dr = \mu r dt + \sigma r dz$

Rendleman, R. and B. Bartter (1980). "The Pricing of Options on Debt Securities". *Journal of Financial and Quantitative Analysis*. **15**: 11–24).

### Pros:

- More tractable model, as it follows a GMB.

### Cons:

- Assumes that interest rates follow a stochastic process similar to stocks, while they usually exhibit a mean-reversion behavior.

# EQUILIBRIUM ONE-FACTOR CT MODELS OF THE SHORT RATE

## 2. Vasicek (1977)

$$dr = a(b - r) dt + \sigma dz$$

Vasicek, O., 1977, "An Equilibrium Characterization of the Term Structure," *Journal of Financial Economics*, 5, 177–188.

Also known as Hull and White (1990) model or an Ornstein–Uhlenbeck process.

Hull, J., and White, A., "Pricing Interest Rate Derivative Securities", *Review of Financial Studies*, 1990, pp. 573–592.

### Pros:

- More tractable model, due to constant volatility.
- Interest rates are mean-reverting (to  $b$ ), at a reversion rate (pace)  $a$ .

### Cons:

- The model assumes a constant volatility, while interest rate volatility is often variable, namely during periods of higher uncertainty, when the estimation of interest rates becomes more complex but also more useful.

# EQUILIBRIUM ONE-FACTOR CT MODELS OF THE SHORT RATE

## 3. Cox, Ingersoll and Ross (CIR)

$$dr = a(b - r) dt + \sigma \sqrt{r} dz$$

Stochastic volatility model => higher volatility with higher interest rates.

Cox, Ingersoll, and Ross. 1985, "A Theory of the Term Structure of Interest Rates", *Econometrica*, Vol 53, March.

### Pros:

- Model closer to reality, as interest rates have stochastic volatilities (higher volatilities with higher interest rates).

### Cons:

- Model becomes less tractable, as it requires the single factor to be  $>0$ , which is not a problem when the single factor is the short-term interest rate and this is positive, but becomes a problem if the short-term rate turns negative or the factor is a different variable.

# NO-ARBITRAGE SHORT RATE CT MODELS

## 1. Ho-Lee (1986)

Ho, T.S.Y., and S.-B. Lee, “Term Structure Movements and Pricing Interest Rate Contingent Claims,” *Journal of Finance*, 41 (December 1986): 1011–29.

$$dr = \theta(t) dt + \sigma dz$$

$\theta(t)$  defines the average direction that  $r$  moves at time  $t$ :

## 2. Hull-White One-Factor Model (1990)

Extended version of Vasicek, to provide an exact fit to the initial TSIR:

$$dr = [\theta(t) - ar] dt + \sigma dz \quad \text{or} \quad dr = a \left[ \frac{\theta(t)}{a} - r \right] dt + \sigma dz$$

Corresponds to the Ho-Lee model, with mean reversion at rate  $a$ .

Hull, J. C., and A. White, “Pricing Interest Rate Derivative Securities,” *The Review of Financial Studies*, 3, 4 (1990): 573–92.

# NO-ARBITRAGE SHORT RATE CT MODELS

## 3. Black-Derman-Toy (1990)

Black, F., E. Derman, and W. Toy, "A One-Factor Model of Interest Rates and Its Application to Treasury Bond Prices," *Financial Analysts Journal*, January/February 1990: 33–39.

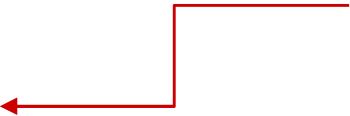
$$d \ln r = [\theta(t) - a(t) \ln r] dt + \sigma(t) dz$$

with  $a(t) = -\frac{\sigma'(t)}{\sigma(t)}$  and  $\sigma'(t)$  is the derivative of  $\sigma$  with respect to  $t$ .

- It is similar to Hull-White One-Factor Model, but in logs and with mean reversion rate  $a$  being time-dependent.
- It doesn't allow negative interest rates.

Constant volatility  $\Rightarrow \sigma'(t) = 0 \Rightarrow a(t) = 0 \Rightarrow$  BDT model:  $d \ln r = \theta(t) dt + \sigma dz$

Log-normal version of Ho-Lee model



# NO-ARBITRAGE SHORT RATE CT MODELS

## 4. Black-Karasinski (1991)

Black, F., and P. Karasinski, “Bond and Option Pricing When Short Rates Are Lognormal,”  
*Financial Analysts Journal*, July/August (1991): 52–59.

- Extended version of BDT (1990) model, where the reversion rate and volatility are determined independently of each other:

$$d \ln r = [\theta(t) - a(t) \ln r] dt + \sigma(t) dz, \text{ with } a(t) = -\frac{\sigma'(t)}{\sigma(t)} \quad - \text{BDT}$$

- The model is the same as BDT (1990), but with no relation between  $a(t)$  and  $\sigma(t)$ .
- As in practice  $a(t)$  and  $\sigma(t)$  are often assumed to be constant, the model becomes:

$$d \ln r = [\theta(t) - a \ln r] dt + \sigma dz$$

## 2.2.3. CT MULTI FACTOR MODELS

1. Fong and Vasicek (1991) model - short rate and its volatility ( $v$ ) as two-state variables

H. G. Fong and O. A. Vasicek: Fixed-income volatility management. Journal of Portfolio Management, 41-56, 1991.

$$dr = \alpha(\bar{r} - r)dt + \sqrt{v}dz_1$$

$$dv = \gamma(\bar{v} - v)dt + \xi\sqrt{v}dz_2$$

## 2. Longstaff and Schwartz (1992) model

Longstaff, F. A. and E. S. Schwartz, "Interest Rate Volatility and the Term Structure: A Two Factor General Equilibrium Model," *Journal of Finance*, 47, 4 (September 1992): 1259–82.

- Longstaff and Schwartz (1992) uses the same two-state variables (the short rate and its volatility), but with a different specification, as the drift is governed by two factors or state variables, while the variance is a function of only one of them:

$$\frac{dQ}{Q} = (\mu X + \theta Y) dt + \sigma \sqrt{Y} dZ_1$$

- With this specification, it is ensured that the drift and the variance are not perfectly correlated.
- The dynamics of the state variables are as follows:

$$dX = (a - bX) dt + c\sqrt{X} dZ_2$$

$$dY = (d - eY) dt + f\sqrt{Y} dZ_3.$$

### 3. Balduzzi et al. (1996) models

Balduzzi, P., S. R. Das, S. Foresi, and R. Sundaran, 1996, "A Simple Approach to Three-Factor Affine Term Structure Models," *The Journal of Fixed Income*, 6, 14–31.

- Balduzzi *et al.* (1996) suggest the use of a 3-factor model by adding the mean of the short-term rate ( $\theta$ ) to a 2-factor model.

$$dr = \mu_r(r, \theta, t)dt + \sigma_r(r, V, t)dz$$

$$d\theta = \mu_\theta(\theta, t)dt + \sigma_\theta(\theta, t)dw$$

$$dV = \mu_V(V, t)dt + \sigma_V(V, t)dy$$

$$dr = \kappa(\theta - r)dt + \sqrt{V}dz$$

$$d\theta = \alpha(\beta - \theta)dt + \eta dw$$

$$dV = a(b - V)dt + \phi\sqrt{V}dy$$