

CHAPTER 6

FORECASTING WITH MOVING AVERAGE (MA) MODELS

6.1 A Model with No Dependence: White Noise

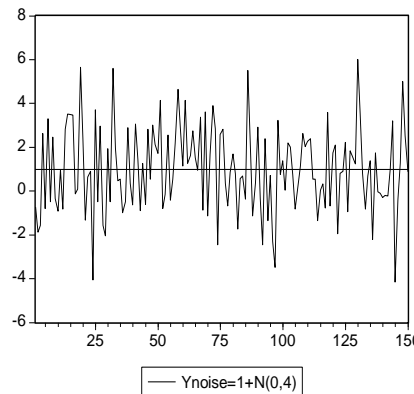
Definition: (White noise)

Let ε_t be a stochastic process.

If $E[\varepsilon_t] = 0$, $Var(\varepsilon_t) = \sigma_\varepsilon^2$ (constant)

and $Cov(\varepsilon_t, \varepsilon_{t-k}) = 0, k \neq 0$

Then we say that the process ε_t is white noise.



Sample: 1 1000 Included observations: 1000			
Autocorrelation	Partial Correlation	AC	PAC
1	-0.020	-0.020	
2	-0.013	-0.014	
3	-0.066	-0.066	
4	-0.027	-0.030	
5	-0.004	-0.007	
6	-0.004	-0.010	
7	0.056	0.052	
8	-0.001	0.000	
9	0.026	0.027	
10	0.018	0.026	
11	-0.030	-0.026	
12	-0.013	-0.009	
13	0.001	0.004	
14	0.040	0.035	
15	0.001	0.000	
16	0.043	0.042	

The white noise process is stationary and it does not exhibit any linear dependence.

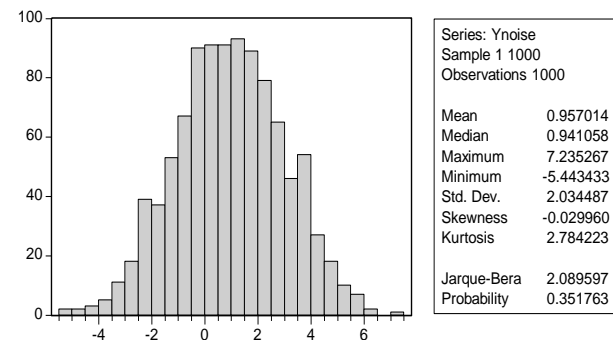
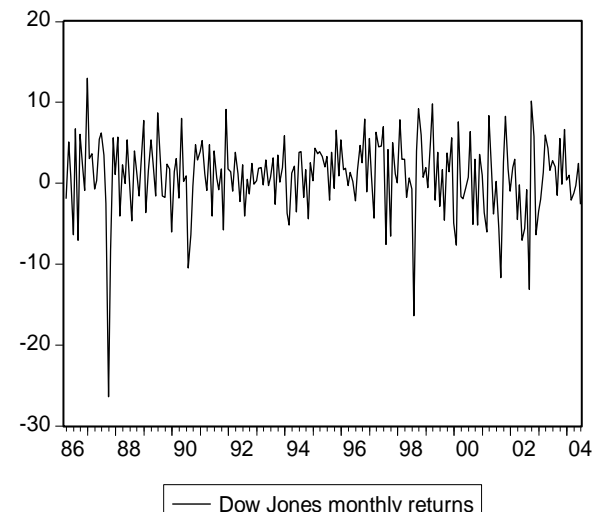
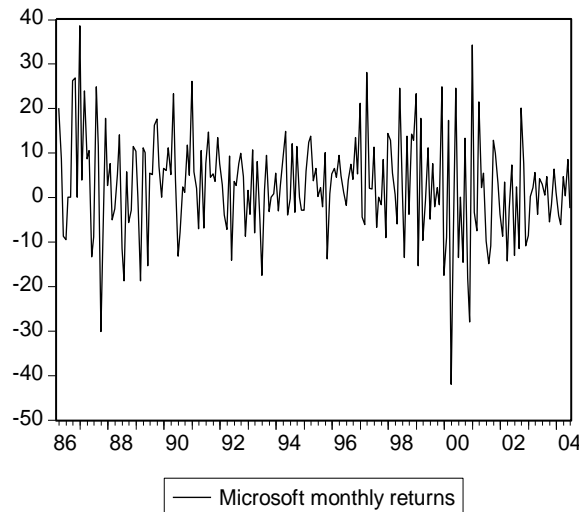


Figure 6.1 Time Series and Autocorrelation Functions of a Simulated White Noise Process

Figure 6.2 Autocorrelation Functions of Monthly Returns to Microsoft and the Dow Jones Index

Difficult to predict



Sample: 1986:03 2004:07
Included observations: 220

Autocorrelation	Partial Correlation	AC	PAC
█	█	1 -0.081	-0.081
█	█	2 -0.094	-0.101
█	█	3 0.132	0.117
█	█	4 -0.017	-0.006
█	█	5 0.008	0.030
█	█	6 -0.013	-0.029
█	█	7 0.106	0.113
█	█	8 0.015	0.023
█	█	9 -0.006	0.024
█	█	10 0.131	0.112
█	█	11 0.013	0.035
█	█	12 -0.016	0.005
█	█	13 -0.020	-0.045
█	█	14 0.030	0.013
█	█	15 -0.075	-0.091
█	█	16 0.064	0.068
█	█	17 0.085	0.051
█	█	18 -0.094	-0.064
█	█	19 -0.049	-0.079
█	█	20 0.046	0.005

Sample: 1986:03 2004:07
Included observations: 220

Autocorrelation	Partial Correlation	AC	PAC
█	█	1 -0.021	-0.021
█	█	2 -0.044	-0.044
█	█	3 -0.056	-0.058
█	█	4 -0.126	-0.131
█	█	5 0.048	0.037
█	█	6 -0.031	-0.045
█	█	7 0.092	0.081
█	█	8 -0.044	-0.057
█	█	9 -0.043	-0.030
█	█	10 0.043	0.035
█	█	11 -0.012	0.006
█	█	12 0.015	-0.006
█	█	13 -0.003	0.003
█	█	14 -0.034	-0.034
█	█	15 -0.059	-0.060
█	█	16 0.039	0.042
█	█	17 0.031	0.012
█	█	18 0.079	0.076
█	█	19 -0.026	-0.028
█	█	20 -0.016	0.005

The Wold Decomposition Theorem

If $\{Y_t\}$ is a covariance stationary process and $\{\varepsilon_t\}$ is a white noise process then there exists a unique linear representation as:

$$Y_t = V_t + \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \dots = V_t + \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j}$$

where V_t is a deterministic component and $\sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j}$ is the stochastic component with $\psi_0 = 1, \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j} < \infty$.

The Wold decomposition theorem guarantees that any (purely nondeterministic) covariance stationary stochastic process can be expressed as a linear combination of past shocks

Finite Representation of the Wold Decomposition Theorem

$$Y_t = \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \dots = (1 + \psi_1 L + \psi_2 L^2 + \dots) \varepsilon_t = \Psi(L) \varepsilon_t$$

The infinite polynomial can be approximated by the ratio of two finite polynomials :

$$\Psi(L) \approx \frac{\Theta_q(L)}{\Phi_p(L)}$$

where $\Theta_q(L) = 1 + \theta_1 L + \theta_2 L^2 + \dots + \theta_q L^q$

and $\Phi_p(L) = 1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p$

and the Wold decomposition can be approximated (or written) as:

$$Y_t = \Psi(L) \varepsilon_t \approx \frac{\Theta_q(L)}{\Phi_p(L)} \varepsilon_t \Leftrightarrow \Phi_p(L) Y_t = \Theta_q(L) \varepsilon_t \Leftrightarrow$$

$$\Leftrightarrow Y_t - \underbrace{\phi_1 Y_{t-1} - \phi_2 Y_{t-2} - \dots - \phi_p Y_{t-p}}_{\text{AR}(p)} = \underbrace{\varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \dots + \theta_q \varepsilon_{t-q}}_{\text{MA}(q)}$$

A moving average process of order $q \geq 0$, referred as MA(q), has the following functional form:

$$Y_t = \mu + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} \cdots + \theta_q \varepsilon_{t-q} + \varepsilon_t$$

where ε_t is a zero-mean white noise process.

A process is **invertible** if it can be written as a linear function of past observations (up to an unpredictable shock):

$$X_t = \varepsilon_t + \pi_1 X_{t-1} + \pi_2 X_{t-2} + \pi_3 X_{t-3} + \dots$$

This happens *iff* all the roots ξ_i of the $\pi(L)$ polynomial are outside the unit circle:

$$|\xi_i| > 1,$$

I.e., *iff* the modules of the inverse roots are smaller than 1:

$$|1/\xi_i| < 1$$

(if $1/\xi = a + bi$, where $i = \sqrt{-1}$, $\sqrt{(a^2 + b^2)} < 1$)

6.3.1 MA(1) Process

$$Y_t = \mu + \theta \varepsilon_{t-1} + \varepsilon_t$$

$$E[Y_t] = \mu$$

$$\sigma_Y^2 = \text{Var}(Y_t) = (1 + \theta^2)\sigma_\varepsilon^2$$

$$\rho_1 = \frac{\theta}{1 + \theta^2}$$

$$\rho_k = 0, k \geq 2$$

r_k decays to zero

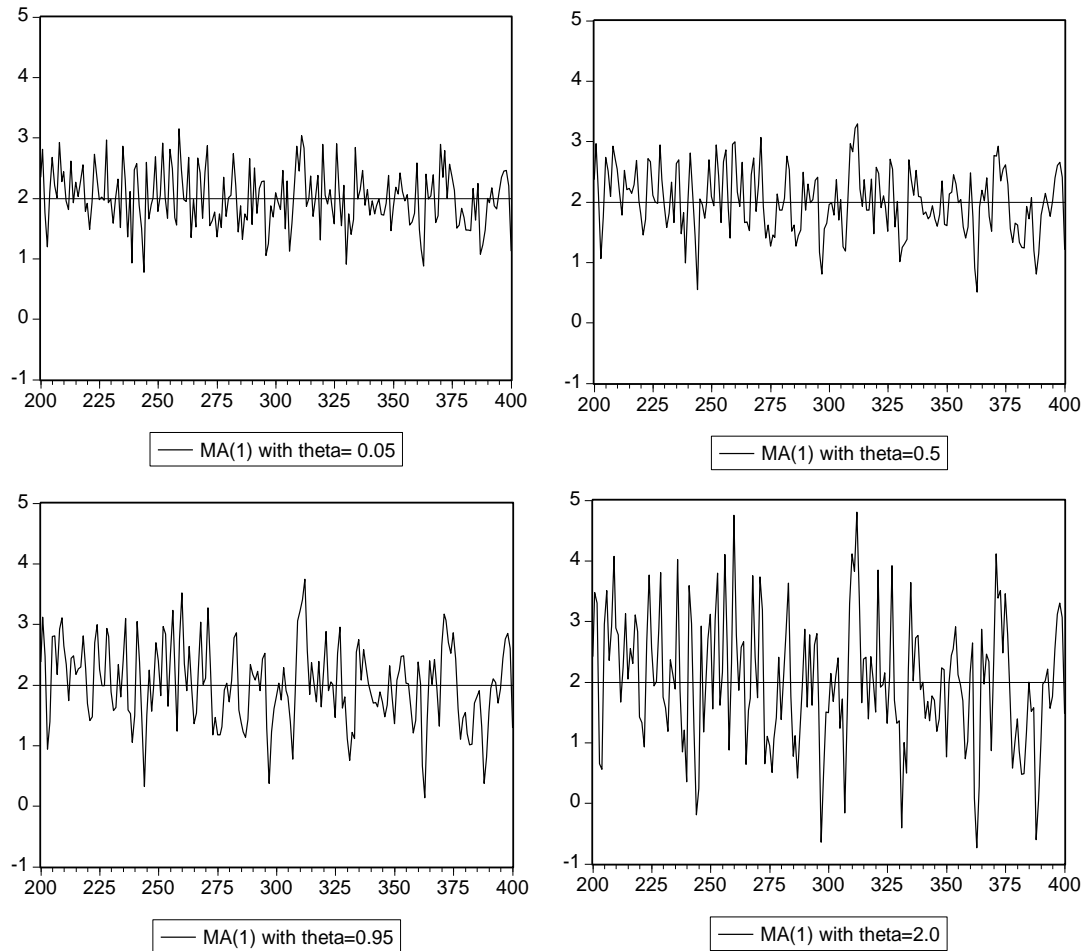


Figure 6.3 Simulated MA(1) Processes

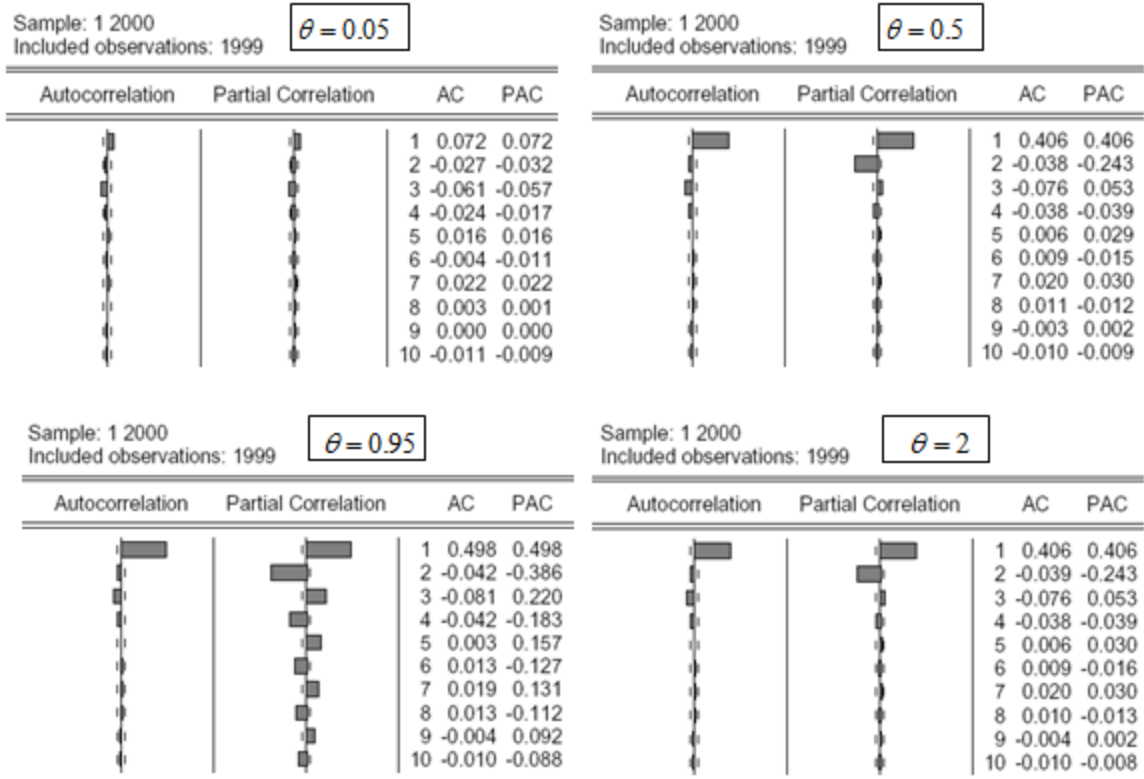
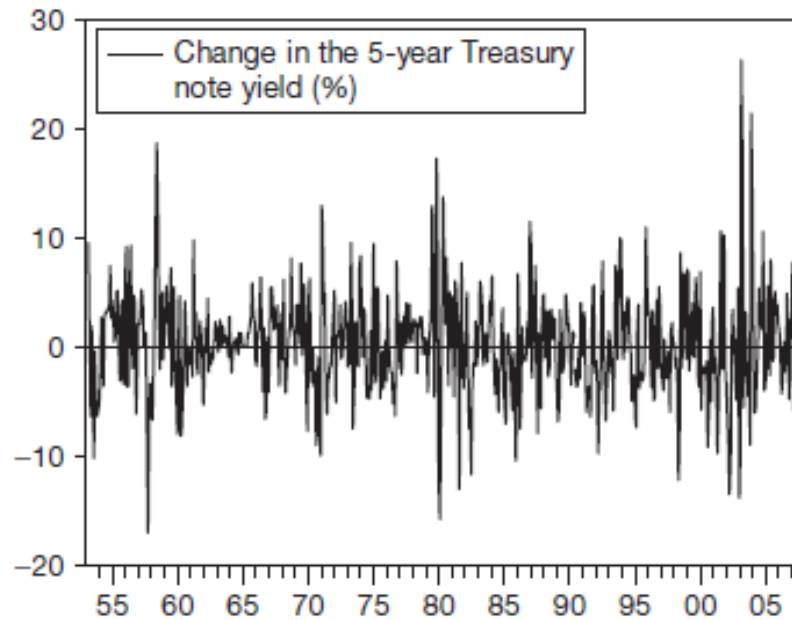


Figure 6.4 Autocorrelation Functions of Simulated MA(1) Processes

We say that an MA(1) process is **invertible** if $|\theta| < 1$

Figure 6.5 Percentage Changes in the 5-Year Treasury Note Yield



Sample: 1953M04 2008M04
Included observations: 660

Autocorrelation	Partial Correlation	AC	PAC	
		1	0.339	0.339
		2	-0.073	-0.213
		3	0.007	0.129
		4	0.014	-0.063
		5	-0.043	-0.017
		6	-0.073	-0.060
		7	-0.069	-0.035

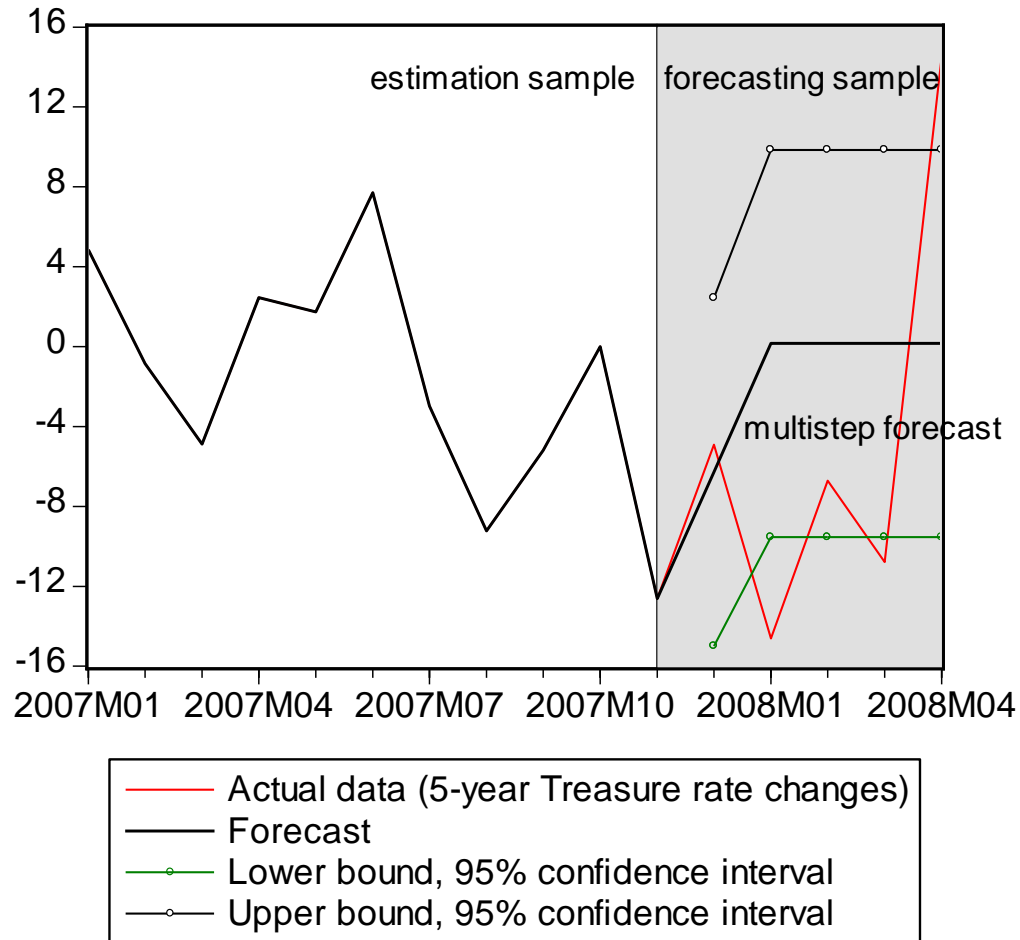
Table 6.1 Estimation Output: 5-Year Treasury Yield (Monthly Percentage Changes)

Dependent Variable: DY				
Method: Least Squares				
Sample (adjusted): 1953M05 2007M11				
Included observations: 655 after adjustments				
Convergence achieved after 7 iterations				
Backcast: 1953M04				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.160159	0.258095	0.620544	0.5351
MA(1)	0.485011	0.034468	14.07130	0.0000
R-squared	0.165370	Mean dependent var	0.168613	
Adjusted R-squared	0.164092	S.D. dependent var	4.866609	
S.E. of regression	4.449443	Akaike info criterion	5.826484	
Sum squared resid	12927.79	Schwarz criterion	5.840177	
Log likelihood	-1906.173	F-statistic	129.3829	
Durbin-Watson stat	2.055799	Prob(F-statistic)	0.000000	
Inverted MA Roots	-.49			

$h = 1$ 12/2007	$f_{t,1} = \hat{\mu} + \hat{\theta}\varepsilon_t =$ $= 0.160 + 0.485\hat{\varepsilon}_t$ $= -6.276\%$	$\sigma_{t+1 t}^2 = \hat{\sigma}_\varepsilon^2 = 4.449^2$	$f(Y_{t+1} I_t) \rightarrow N(\mu_{t+1 t}, \sigma_{t+1 t}^2)$ $= N(-6.276, 4.449^2)$
$h = 2$ 1/2008	$f_{t,2} = \hat{\mu} = 0.160\%$	$\sigma_{t+2 t}^2 = \hat{\sigma}_\varepsilon^2(1 + \hat{\theta}^2)$ $= 4.449^2(1 + 0.485^2)$ $= 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+2} I_t) \rightarrow N(0.16, 23.683)$
$h = 3$ 2/2008	$f_{t,3} = \hat{\mu} = 0.160\%$	$\sigma_{t+3 t}^2 = 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+3} I_t) \rightarrow N(0.16, 23.683)$
$h = 4$ 3/2008	$f_{t,4} = \hat{\mu} = 0.160\%$	$\sigma_{t+4 t}^2 = 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+4} I_t) \rightarrow N(0.16, 23.683)$
$h = 5$ 4/2008	$f_{t,5} = \hat{\mu} = 0.160\%$	$\sigma_{t+5 t}^2 = 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+5} I_t) \rightarrow N(0.16, 23.683)$

Table 6.2 December 2007-April 2008 Forecasts of 5-year Treasury Yield Changes

Figure 6.6 Multistep Forecast of Monthly Changes of 5-year Treasury Yield

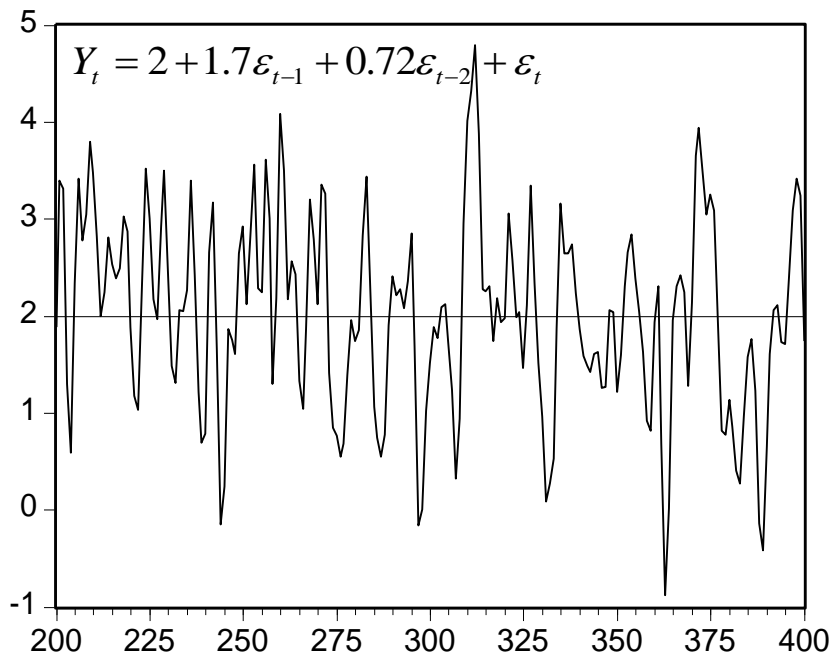


MA(2) Process

$$Y_t = \mu + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \varepsilon_t$$

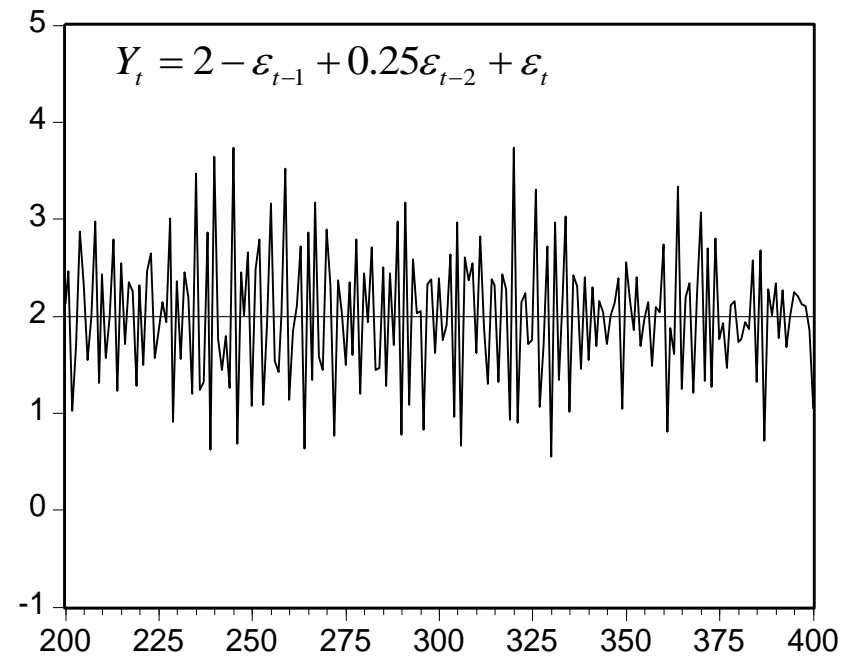
$$E[Y_t] = \mu$$

$$\sigma_Y^2 = \text{Var}(Y_t) = (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2$$



— MA(2) with theta_1=1.70 and theta_2=0.72

(a)



— MA(2) with theta_1=-1 and theta_2=0.25

(b)

Figure 6.7 Simulated MA(2) Processes

Figure 6.8 Autocorrelation Functions of Simulated MA(2) Processes

$$\rho_1 = \frac{\theta_1 + \theta_1\theta_2}{1 + \theta_1^2 + \theta_2^2}$$

$$\rho_2 = \frac{\theta_2}{1 + \theta_1^2 + \theta_2^2}$$

$$\rho_k = 0, k \geq 3$$

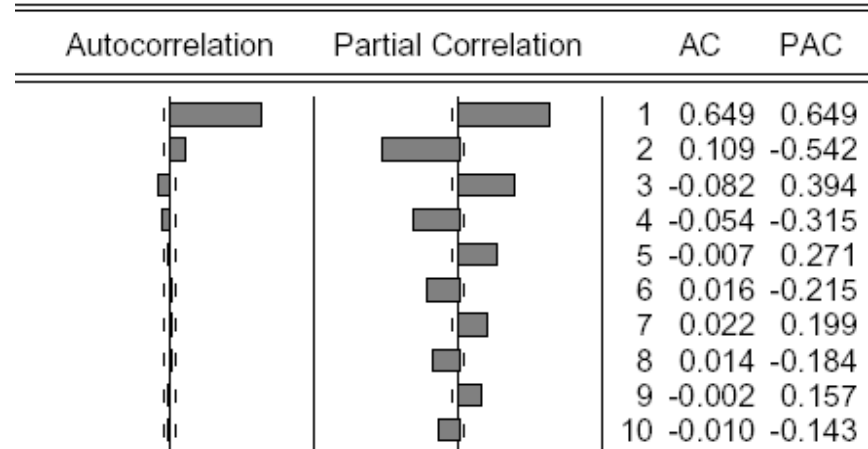
r_k decays to zero

The order q of an MA(q) process can be identified as the last k such that $\rho_k \neq 0$

Sample: 1 2000

Included observations: 1998

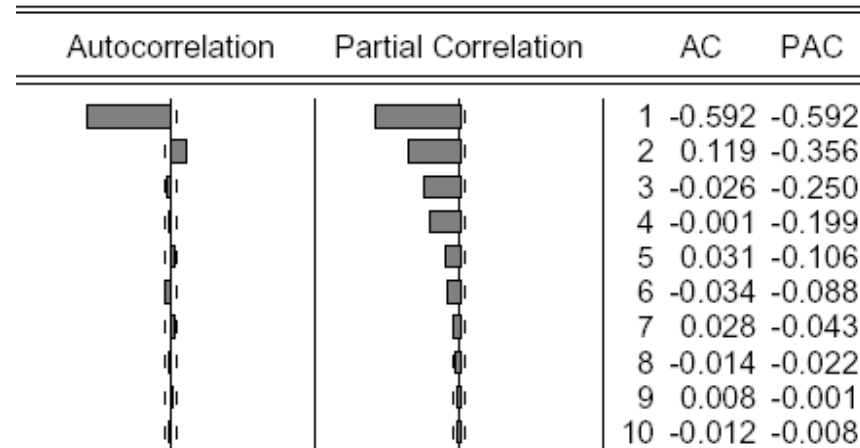
$$Y_t = 2 + 1.7\varepsilon_{t-1} + 0.72\varepsilon_{t-2} + \varepsilon_t$$



Sample: 1 2000

Included observations: 1998

$$Y_t = 2 - \varepsilon_{t-1} + 0.25\varepsilon_{t-2} + \varepsilon_t$$



The MA(2) process is invertible if the roots of the characteristic equation are, in absolute value, greater than one.

$$1 - \theta_1 x + \theta_2 x^2 = 0$$

Sample: 1 2000

Included observations: 1998

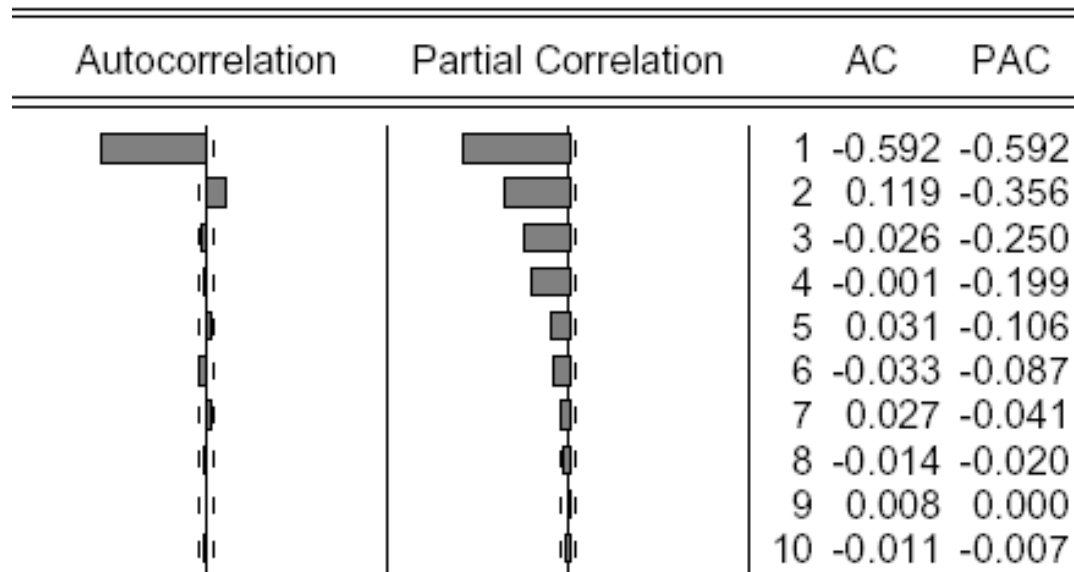


Figure 6.9 Autocorrelation Functions of MA Process $Y_t = 2 - 4\varepsilon_{t-1} + 4\varepsilon_{t-2} + \varepsilon_t$