

Illustration 1

Aim:

- Explaining annual health care expenditures (*med*) and the annual number of medical consultations (*mdu*) as a function of health insurance characteristics (*lcoins*), socio-economic factors (*female*, *age*, *lfam*, *child*) and health status variables (*ndisease*)

Details:

- Cameron and Trivedi (2005), ch. 16.6

Illustration 1 – Question 1

```
. summarize med
```

Variable	Obs	Mean	Std. Dev.	Min	Max
med	20186	171.5892	698.2689	0	39182.02

```
. summarize med if med>0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
med	15733	220.1551	784.1543	.4276146	39182.02

- Non-negative outcome
- High proportion of zeros: 4453 observations (22.1% of the sample units)

Illustration 1 – Question 2.1

```
. drop if year!=1  
(14548 observations deleted)
```

```
. poisson med lcoins ndisease female age lfam child, robust  
(...)
```

		Robust				
med		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
lcoins		-.0478877	.0245805	-1.95	0.051	-.0960646 .0002892
ndisease		.0281757	.0060746	4.64	0.000	.0162696 .0400817
female		.2263111	.1192456	1.90	0.058	-.007406 .4600282
age		.0097846	.003701	2.64	0.008	.0025307 .0170384
lfam		-.0928784	.0882497	-1.05	0.293	-.2658446 .0800879
child		-.4112285	.2444742	-1.68	0.093	-.8903892 .0679321
_cons		4.631177	.2039135	22.71	0.000	4.231514 5.03084

```
. estimates store poisson
```

Illustration 1 – Question 2.2

```
. poisson med lcoins ndisease female age lfam child if med>0, robust
note: you are responsible for interpretation of noncount dep. variable
(...)
```

```
Poisson regression                                Number of obs   =          4451
                                                    Wald chi2(6)    =          66.48
                                                    Prob > chi2     =          0.0000
Log pseudolikelihood = -1087511.2                Pseudo R2       =          0.0657
```

		Robust				
med		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+						
lcoins		-.0200287	.0246889	-0.81	0.417	-.068418 .0283606
ndisease		.0210867	.0062785	3.36	0.001	.0087812 .0333923
female		.1474704	.1177851	1.25	0.211	-.0833842 .3783249
age		.0087803	.0036967	2.38	0.018	.0015349 .0160258
lfam		-.0823594	.0902764	-0.91	0.362	-.2592978 .0945791
child		-.3678983	.2481063	-1.48	0.138	-.8541777 .1183811
_cons		4.91825	.2098069	23.44	0.000	4.507036 5.329464

```
. estimates store poisson0
```

Illustration 1 – Question 2.3

```
. gen lmed=log(med)
(1187 missing values generated)
```

```
. regress lmed lcoins ndisease female age lfam child if med>0, robust
(...)
```

		Robust				
lmed		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
lcoins		-.054901	.0099633	-5.51	0.000	-.0744341 -.0353679
ndisease		.030717	.0030927	9.93	0.000	.0246537 .0367803
female		.1199091	.0417646	2.87	0.004	.0380298 .2017885
age		.0072817	.0020791	3.50	0.000	.0032055 .0113578
lfam		-.1357759	.0442937	-3.07	0.002	-.2226136 -.0489382
child		-.4972463	.0737896	-6.74	0.000	-.6419107 -.3525819
_cons		3.954487	.1043663	37.89	0.000	3.749877 4.159096

```
. estimates store log0
```

Illustration 1 – Question 2.4

```
. gen lmed1=log(med+1)
```

```
. regress lmed1 lcoins ndisease female age lfam child, robust
```

```
(...)
```

		Robust				
lmed1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
lcoins		-.1485372	.0126915	-11.70	0.000	-.1734174 -.1236569
ndisease		.0584591	.0039132	14.94	0.000	.0507878 .0661304
female		.3064252	.0535977	5.72	0.000	.201353 .4114974
age		.0089512	.0027704	3.23	0.001	.0035202 .0143821
lfam		-.1806985	.0556686	-3.25	0.001	-.2898305 -.0715666
child		-.4759264	.0946324	-5.03	0.000	-.6614425 -.2904103
_cons		3.000465	.1346348	22.29	0.000	2.736529 3.264401

```
. estimates store log1
```

Illustration 1 – Question 2 (summary)

```
. estimates table poisson poisson0, b star(0.1 0.05 0.01)
```

Variable	poisson	poisson0
lcoins	-.04788772*	-.0200287
ndisease	.02817566***	.02108674***
female	.22631112*	.14747037
age	.00978456***	.00878032**
lfam	-.09287836	-.08235938
child	-.41122852*	-.3678983
_cons	4.6311773***	4.9182498***

legend: * p<.1; ** p<.05; *** p<.01

```
. estimates table log0 log1, b star(0.1 0.05 0.01)
```

Variable	log0	log1
lcoins	-.05490096***	-.14853716***
ndisease	.03071698***	.0584591***
female	.11990912***	.3064252***
age	.00728165***	.00895117***
lfam	-.13577591***	-.18069853***
child	-.4972463***	-.47592639***
_cons	3.9544865***	3.0004653***

Illustration 1 – Question 3.1

```
. summarize mdu
```

Variable	Obs	Mean	Std. Dev.	Min	Max
-----+-----					
mdu	5638	2.877971	4.332918	0	69

```
. tabulate mdu
```

number			
face-to-fac			
e md visits	Freq.	Percent	Cum.
-----+-----			
0	1,729	30.67	30.67
1	1,047	18.57	49.24
2	814	14.44	63.68
3	511	9.06	72.74
4	385	6.83	79.57
5	271	4.81	84.37
6	188	3.33	87.71
7	156	2.77	90.48
8	134	2.38	92.85
9	80	1.42	94.27
10	56	0.99	95.26
11	56	0.99	96.26
12	32	0.57	96.83

(...)

Illustration 1 – Question 3.2.1

```
. poisson mdu lcoins ndisease female age lfam child
```

```
Poisson regression                                Number of obs   =          5638
                                                    LR chi2(6)       =       2296.50
                                                    Prob > chi2      =          0.0000
Log likelihood = -17173.058                      Pseudo R2       =          0.0627
```

mdu	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
lcoins	-.0682563	.0038214	-17.86	0.000	-.0757462	-.0607665
ndisease	.03538	.0010286	34.40	0.000	.0333639	.037396
female	.1276798	.0164078	7.78	0.000	.0955211	.1598386
age	.0041167	.0007841	5.25	0.000	.0025798	.0056536
lfam	-.141855	.0154716	-9.17	0.000	-.1721789	-.1115312
child	.056673	.0278998	2.03	0.042	.0019903	.1113557
_cons	.7463066	.0385653	19.35	0.000	.67072	.8218932

```
. estimates store cpoissonml
```

Illustration 1 – Question 3.2.2

```
. poisson mdu lcoins ndisease female age lfam child, robust
```

```
Poisson regression                                Number of obs   =          5638
                                                    Wald chi2(6)    =          388.20
                                                    Prob > chi2     =          0.0000
Log pseudolikelihood = -17173.058                Pseudo R2       =          0.0627
```

		Robust				
mdu		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
lcoins		-.0682563	.0094538	-7.22	0.000	-.0867854 -.0497273
ndisease		.03538	.002696	13.12	0.000	.0300959 .0406641
female		.1276798	.04043	3.16	0.002	.0484385 .2069211
age		.0041167	.0019155	2.15	0.032	.0003624 .007871
lfam		-.141855	.0336317	-4.22	0.000	-.207772 -.0759381
child		.056673	.0625248	0.91	0.365	-.0658734 .1792194
_cons		.7463066	.0861806	8.66	0.000	.5773956 .9152175

```
. estimates store cpoissonqml
```

Illustration 1 – Question 3.2.3

```
. nbreg mdu lcoins ndisease female age lfam child, dispersion(constant)
```

```
Negative binomial regression          Number of obs   =       5638
                                      LR chi2(6)         =       529.02
Dispersion      = constant            Prob > chi2       =       0.0000
Log likelihood = -12128.041            Pseudo R2        =       0.0213
```

	mdu	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
	lcoins	-.076432	.0069326	-11.02	0.000	-.0900197 -.0628442
	ndisease	.0309127	.0019493	15.86	0.000	.0270922 .0347333
	female	.1447002	.029679	4.88	0.000	.0865305 .2028699
	age	.0027029	.0014494	1.86	0.062	-.0001378 .0055436
	lfam	-.1235042	.0276393	-4.47	0.000	-.1776761 -.0693323
	child	-.0049686	.0499784	-0.10	0.921	-.1029245 .0929873
	_cons	.8537349	.0697031	12.25	0.000	.7171195 .9903504
-----+-----						
	/lndelta	1.278115	.0319841			1.215427 1.340802
-----+-----						
	delta	3.589865	.1148187			3.371733 3.822109

```
Likelihood-ratio test of delta=0:  chibar2(01) = 1.0e+04 Prob>=chibar2 = 0.000
```

```
. estimates store cnegbin1
```

Overdispersion test:
the hypothesis of
correct specification
of the Poisson model
is rejected

Illustration 1 – Question 3.2.4

```
. nbreg mdu lcoins ndisease female age lfam child, dispersion(mean)
```

```
Negative binomial regression      Number of obs   =       5638
                                LR chi2(6)             =       459.10
Dispersion      = mean           Prob > chi2       =       0.0000
Log likelihood = -12163.001      Pseudo R2        =       0.0185
```

	mdu	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
lcoins		-.0719827	.0083402	-8.63	0.000	-.0883292	-.0556362
ndisease		.0391488	.002618	14.95	0.000	.0340175	.0442801
female		.1229883	.0350213	3.51	0.000	.0543479	.1916288
age		.0035899	.0017692	2.03	0.042	.0001223	.0070574
lfam		-.1834739	.036481	-5.03	0.000	-.2549754	-.1119723
child		.0842864	.0621692	1.36	0.175	-.0375631	.2061359
_cons		.7638038	.087809	8.70	0.000	.5917012	.9359063
/lnalpha		.2193579	.0274111			.1656333	.2730826
alpha		1.245277	.0341344			1.18014	1.314009

```
Likelihood-ratio test of alpha=0:  chibar2(01) = 1.0e+04 Prob>=chibar2 = 0.000
```

```
. estimates store cnegbin2
```

Overdispersion test:
the hypothesis of
correct specification
of the Poisson model
is rejected

Illustration 1 – Question 3.2 (summary)

```
. estimates table cpoissonml cpoissonqml cnegbin1 cnegbin2, b star(0.1 0.05 0.01)
```

Variable	cpoissonml	cpoissonqml	cnegbin1	cnegbin2
mdu				
lcoins	-.06825635***	-.06825635***	-.07643198***	-.0719827***
ndisease	.03537998***	.03537998***	.03091273***	.03914881***
female	.12767984***	.12767984***	.1447002***	.12298834***
age	.0041167***	.0041167**	.0027029*	.00358987**
lfam	-.14185505***	-.14185505***	-.1235042***	-.18347387***
child	.05667301**	.05667301	-.00496864	.0842864
_cons	.74630658***	.74630658***	.85373493***	.76380377***
lndelta				
_cons			1.2781147***	
lnalpha				
_cons				.21935795***

legend: * p<.1; ** p<.05; *** p<.01

Illustration 1 – Question 3.3

```
. estimates restore cpoissonml  
(results cpoissonml are active now)
```

Formulas:

$$Pr(Y_i = y|x_i) = \frac{e^{-\lambda_i} \lambda_i^y}{y!}, \quad \lambda_i = E(Y|X) = \exp(x_i' \beta)$$

- $Pr(Y_i = 0|x_i) = e^{-\lambda_i}$
- $Pr(Y_i = 1|x_i) = e^{-\lambda_i} \lambda_i$
- $Pr(Y_i \geq 2|x_i) = 1 - Pr(Y_i = 0|x_i) - Pr(Y_i = 1|x_i)$

```
. scalar lambda0 = exp(_b[_cons] + _b[lcoins]*log(1) + _b[age]*50 +  
_b[lfam]*log(3))
```

```
. scalar lambda1 = exp(_b[_cons] + _b[lcoins]*log(51) + _b[age]*50 +  
_b[lfam]*log(3))
```

```
. scalar lambda2 = exp(_b[_cons] + _b[lcoins]*log(101) + _b[age]*50 +  
_b[lfam]*log(3))
```

Note that lcoins is defined in the dataset as $\ln(\text{coins}+1)$ to avoid discarding 0's

Illustration 1 – Question 3.3 (cont.)

```
. display lambda0  
2.2173168
```

```
. display lambda1  
1.695412
```

```
. display lambda2  
1.6181549
```

```
. display exp(-lambda0)  
.10890092
```

```
. display exp(-lambda1)  
.18352361
```

```
. display exp(-lambda2)  
.19826418
```

```
. display exp(-lambda0)*lambda0  
.24146784
```

```
. display exp(-lambda1)*lambda1  
.31114812
```

```
. display exp(-lambda2)*lambda2  
.32082215
```

```
. display 1-0.10890092-0.24146784  
.64963124
```

```
. display 1-0.18352361-0.31114812  
.50532827
```

```
. display 1-0.19826418-0.32082215  
.48091367
```

Illustration 1 – Question 3.3 (cont.)

<i>coins:</i>	0	50	100
$E(mdu \dots)$	2.2	1.7	1.6
$Pr(mdu = 0 \dots)$	10.9	18.4	19.8
$Pr(mdu = 1 \dots)$	24.1	31.1	32.1
$Pr(mdu \geq 2 \dots)$	65.0	50.5	48.1

Illustration 1 – Question 4.1

```
. xtset id year
      panel variable:  id (unbalanced)
      time variable:  year, 1 to 5, but with gaps
                  delta:  1 unit
```

```
. xtdescribe
      id:  125024, 125025, ..., 632167          n =          5908
      year:  1, 2, ..., 5                      T =              5
      Delta(year) = 1 unit
      Span(year)  = 5 periods
      (id*year uniquely identifies each observation)
```

Distribution of T_i:							
	min	5%	25%	50%	75%	95%	max
	1	2	3	3	5	5	5

Freq.	Percent	Cum.	Pattern
3710	62.80	62.80	111..
1584	26.81	89.61	11111
156	2.64	92.25	1....
147	2.49	94.74	11...
79	1.34	96.07	..1..
66	1.12	97.19	.11..
33	0.56	97.75	..111
33	0.56	98.31	.1111
29	0.49	98.80	...11
71	1.20	100.00	(other patterns)
5908	100.00		XXXXX

Illustration 1 – Question 4.2.1

```
. poisson mdu lcoins ndisease female age lfam child, vce(cluster id)
```

```
Poisson regression                                Number of obs   =       20186
                                                    Wald chi2(6)    =       476.93
                                                    Prob > chi2     =       0.0000
Log pseudolikelihood = -62579.401                Pseudo R2      =       0.0609
                                                    (Std. Err. adjusted for 5908 clusters in id)
```

		Robust				
mdu		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
lcoins		-.0808023	.0080013	-10.10	0.000	-.0964846 -.0651199
ndisease		.0339334	.0026024	13.04	0.000	.0288328 .039034
female		.1717862	.0342551	5.01	0.000	.1046473 .2389251
age		.0040585	.0016891	2.40	0.016	.000748 .0073691
lfam		-.1481981	.0323434	-4.58	0.000	-.21159 -.0848062
child		.1030453	.0506901	2.03	0.042	.0036944 .2023961
_cons		.748789	.0785738	9.53	0.000	.5947872 .9027907

```
. estimates store pooled
```

Note that s.e. are almost 4 times larger than their standard version, that overlooks overdispersion

Illustration 1 – Question 4.2.2

```
. xtpoisson mdu lcoins ndisease female age lfam child, re
(...)
```

mdu	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
lcoins	-.0878258	.0068682	-12.79	0.000	-.1012873	-.0743642
ndisease	.0387629	.0022046	17.58	0.000	.034442	.0430839
female	.1667192	.0286298	5.82	0.000	.1106058	.2228325
age	.0019159	.0011134	1.72	0.085	-.0002663	.0040982
lfam	-.1351786	.0260022	-5.20	0.000	-.186142	-.0842152
child	.1082678	.0341477	3.17	0.002	.0413396	.1751961
_cons	.7574177	.0618346	12.25	0.000	.6362241	.8786112
/lnalpha	.0251256	.0209586			-.0159526	.0662038
alpha	1.025444	.0214919			.984174	1.068444

Likelihood-ratio test of alpha=0: $\chi^2(01) = 3.9e+04$ Prob>= $\chi^2 = 0.000$

```
. estimates store RE
```

Illustration 1 – Question 4.2.3

```
. xtpoisson mdu lcoins ndisease female age lfam child, fe
note: 265 groups (265 obs) dropped because of only one obs per group
note: 666 groups (2130 obs) dropped because of all zero outcomes
note: lcoins dropped because it is constant within group
note: ndisease dropped because it is constant within group
note: female dropped because it is constant within group
(...)
```

```
Log likelihood = -24173.211      Wald chi2(3)      =      19.24
                                Prob > chi2      =      0.0002
```

mdu	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
age	-.0112009	.0039024	-2.87	0.004	-.0188494	-.0035523
lfam	.0877134	.0554606	1.58	0.114	-.0209874	.1964141
child	.1059867	.0437744	2.42	0.015	.0201905	.1917829

```
. estimates store FE
```

Illustration 1 – Question 4.2 (summary)

```
. estimates table pooled RE FE, b star(0.1 0.05 0.01)
```

Variable	pooled	RE	FE
mdu			
lcoins	-.08080226***	-.08782576***	
ndisease	.0339334***	.03876295***	
female	.17178621***	.16671918***	
age	.00405853**	.00191594*	-.01120087***
lfam	-.1481981***	-.13517858***	.08771336
child	.10304526**	.10826785***	.10598669**
_cons	.74878896***	.7574177***	
lnalpha			
_cons		.02512562	

legend: * p<.1; ** p<.05; *** p<.01

Note that relative to pooled and RE results, FE:

- Coefficients differ substantially (allow correlation between individual heterogeneity and time-varying covariates)
- Tend to present larger s.e. for identified coefficients (less significance)

Illustration 1 – Question 4.3

```
. hausman FE RE
```

---- Coefficients ----				
	(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
	EF	EA	Difference	S.E.
age	-.0112009	.0019159	-.0131168	.0037402
lfam	.0877134	-.1351786	.2228919	.0489874
child	.1059867	.1082678	-.0022812	.0273886

b = consistent under Ho and Ha; obtained from xtpoisson
B = inconsistent under Ha, efficient under Ho; obtained from xtpoisson

Test: Ho: difference in coefficients not systematic

```
chi2(3) = (b-B)'[(V_b-V_B)^(-1)](b-B)
          = 29.73
Prob>chi2 = 0.0000
```

→ The hypothesis of random effects is rejected

Illustration 2 – Question 1

```
. drop if LE==1  
(1922 observations deleted)
```

```
. summarize LEV_LT1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
LEV_LT1	30304	.0734047	.1668595	0	.9982489

```
. count if LEV_LT==0  
23027
```

```
. display 23027/30304  
.75986668
```

Illustration 2 – Question 2

```
. glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE, family(binomial) link(logit)
vce(cluster id)
(...)
```

```

                                     AIC          =   .4259979
Log pseudolikelihood = -6448.71996      BIC          = -304709.7
```

(Std. Err. adjusted for 4995 clusters in id)

		Robust				
LEV_LT1	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
SIZE2	.3160809	.0150444	21.01	0.000	.2865946	.3455673
COLLAT2	.453255	.1394787	3.25	0.001	.1798818	.7266282
PROF1	-3.629699	.2667238	-13.61	0.000	-4.152469	-3.10693
GROWTH2	.0004846	.0002476	1.96	0.050	-6.89e-07	.0009699
AGE	-.0045993	.0017012	-2.70	0.007	-.0079335	-.0012651
_cons	-6.823869	.2264392	-30.14	0.000	-7.267681	-6.380056

```
. estimates store LOGIT
```


Illustration 2 – Question 3.1, 3.2 & 3.3

3.1 Logit model:

$$E(LEV_LT_i|x_i) = \frac{e^{x_i'\beta}}{1 + e^{x_i'\beta}}$$

```
. estimates restore LOGIT  
(results LOGIT are active now)
```

```
. scalar xb = _b[_cons] + _b[SIZE2]*13.54 + _b[COLLAT2]*0.41 + _b[PROF1]*0.07 +  
_b[GROWTH2]*15.03 + _b[AGE]*19
```

```
. display exp(xb)/(1+exp(xb))  
.06341858
```

3.2 By definition, it is approximately equal to one

3.3 By definition, $E(LEV_LT_i|x_i, Y_i > 0) \approx E(LEV_LT_i|x_i)$

Illustration 2 – Question 4

```
. quietly glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE, family(binomial)  
link(logit) vce(cluster id)
```

```
. predict xbl, xb
```

```
. gen xbl2=xbl^2
```

```
. gen xbl3=xbl^3
```

```
. quietly glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE xbl2 xbl3 ,  
family(binomial) link(logit) vce(cluster id)
```

```
. test xbl2 xbl3
```

```
( 1)  [LEV_LT1]xbl2 = 0
```

```
( 2)  [LEV_LT1]xbl3 = 0
```

```
chi2( 2) = 131.99
```

```
Prob > chi2 = 0.0000
```

Illustration 2 – Question 4

```
. quietly glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE, family(binomial)  
link(probit) vce(cluster id)
```

```
. predict xbP, xb
```

```
. gen xbP2=xbP^2
```

```
. gen xbP3=xbP^3
```

```
. quietly glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE xbP2 xbP3 ,  
family(binomial) link(probit) vce(cluster id)
```

```
. test xbP2 xbP3
```

```
( 1)  [LEV_LT1]xbP2 = 0
```

```
( 2)  [LEV_LT1]xbP3 = 0
```

```
chi2( 2) = 120.35
```

```
Prob > chi2 = 0.0000
```

Illustration 2 – Question 4

```
. quietly glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE, family(binomial)
link(logit) vce(cluster id)
. margins, dydx (SIZE2 COLLAT2 PROF1 GROWTH2 AGE) atmean
Conditional marginal effects          Number of obs      =       30,304
Model VCE      : Robust
Expression     : Predicted mean LEV_LT1, predict()
dy/dx w.r.t.   : SIZE2 COLLAT2 PROF1 GROWTH2 AGE
at
      : SIZE2          =      13.53998 (mean)
      : COLLAT2        =       .4105925 (mean)
      : PROF1          =       .0678864 (mean)
      : GROWTH2        =       15.034 (mean)
      : AGE            =      19.38932 (mean)
```

		Delta-method				
		dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]
-----+						
SIZE2		.0188751	.0008384	22.51	0.000	.0172319 .0205184
COLLAT2		.0270666	.0080893	3.35	0.001	.011212 .0429213
PROF1		-.2167517	.0160444	-13.51	0.000	-.2481981 -.1853052
GROWTH2		.0000289	.0000148	1.95	0.051	-9.42e-08 .000058
AGE		-.0002747	.0001021	-2.69	0.007	-.0004747 -.0000746

Illustration 2 – Question 4

```
. margins, dydx (SIZE2 COLLAT2 PROF1 GROWTH2 AGE)
```

Average marginal effects

Number of obs = 30,304

Model VCE : Robust

Expression : Predicted mean LEV_LT1, predict()

dy/dx w.r.t. : SIZE2 COLLAT2 PROF1 GROWTH2 AGE

		Delta-method					
		dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
SIZE2		.0210336	.0010403	20.22	0.000	.0189947	.0230724
COLLAT2		.0301618	.0091286	3.30	0.001	.01227	.0480536
PROF1		-.2415379	.0185096	-13.05	0.000	-.277816	-.2052598
GROWTH2		.0000322	.0000165	1.95	0.051	-8.76e-08	.0000646
AGE		-.0003061	.0001139	-2.69	0.007	-.0005292	-.0000829

Illustration 2 – Question 4

```
. quietly glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE, family(binomial)
link(probit) vce(cluster id)
. margins, dydx (SIZE2 COLLAT2 PROF1 GROWTH2 AGE) atmean
Conditional marginal effects          Number of obs      =       30,304
Model VCE      : Robust
Expression     : Predicted mean LEV_LT1, predict()
dy/dx w.r.t.   : SIZE2 COLLAT2 PROF1 GROWTH2 AGE
at
      : SIZE2          =      13.53998 (mean)
      : COLLAT2        =       .4105925 (mean)
      : PROF1          =       .0678864 (mean)
      : GROWTH2        =       15.034 (mean)
      : AGE            =      19.38932 (mean)
```

		Delta-method				
		dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]
-----+						
SIZE2		.0201057	.0009364	21.47	0.000	.0182704 .0219409
COLLAT2		.0245778	.008263	2.97	0.003	.0083827 .040773
PROF1		-.2265281	.017486	-12.95	0.000	-.2608001 -.1922561
GROWTH2		.0000398	.0000184	2.17	0.030	3.79e-06 .0000758
AGE		-.0002937	.0001078	-2.72	0.006	-.0005049 -.0000824

Illustration 2 – Question 4

```
. margins, dydx (SIZE2 COLLAT2 PROF1 GROWTH2 AGE)
```

Average marginal effects

Number of obs = 30,304

Model VCE : Robust

Expression : Predicted mean LEV_LT1, predict()

dy/dx w.r.t. : SIZE2 COLLAT2 PROF1 GROWTH2 AGE

		Delta-method					
		dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
	+						
SIZE2		.0213241	.0010521	20.27	0.000	.019262	.0233862
COLLAT2		.0260673	.0088125	2.96	0.003	.008795	.0433395
PROF1		-.2402559	.0189853	-12.65	0.000	-.2774664	-.2030453
GROWTH2		.0000422	.0000195	2.17	0.030	4.04e-06	.0000804
AGE		-.0003115	.0001144	-2.72	0.006	-.0005357	-.0000873

•

Illustration 2 – Question 5

```
. xtset id YEAR
      panel variable:  id (unbalanced)
      time variable:  YEAR, 1995 to 2001, but with gaps
                delta:  1 unit

. gen LEVt= LEV_LT1/(1- LEV_LT1)

. xtpoisson LEVt SIZE2 COLLAT2 PROF1 GROWTH2 AGE, fe vce(bootstrap)
(...)
Conditional fixed-effects Poisson regression      Number of obs      =      13245
Group variable: id                               Number of groups    =      2048
(...)
```

(Replications based on 2048 clusters in id)

		Observed	Bootstrap			Normal-based
LEVt		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
SIZE2		-.2563746	.1857317	-1.38	0.167	-.620402 .1076529
COLLAT2		.7169194	.4518515	1.59	0.113	-.1686932 1.602532
PROF1		-1.205703	2.287238	-0.53	0.598	-5.688608 3.277201
GROWTH2		.0024675	.0016065	1.54	0.125	-.0006813 .0056162
AGE		.0402399	.0242907	1.66	0.098	-.007369 .0878487

Illustration 2 – Question 5

```
. gen LEVtlink=ln(LEV_LT1/(1- LEV_LT1))
(23027 missing values generated)
```

```
. xtreg LEVtlink SIZE2 COLLAT2 PROF1 GROWTH2 AGE, fe vce(bootstrap)
(running xtreg on estimation sample)
(...)
```

```
Fixed-effects (within) regression          Number of obs      =          7277
Group variable: id                        Number of groups   =          2064
(...)
```

(Replications based on 2064 clusters in id)

		Observed	Bootstrap	Normal-based		
LEVtlink		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
SIZE2		-.0741736	.0541096	-1.37	0.170	-.1802265 .0318793
COLLAT2		-.1839106	.1936327	-0.95	0.342	-.5634237 .1956025
PROF1		-1.928813	.4180013	-4.61	0.000	-2.74808 -1.109545
GROWTH2		.0017424	.0005736	3.04	0.002	.0006181 .0028667
AGE		-.0132349	.0073683	-1.80	0.072	-.0276765 .0012067
_cons		.4652777	.8023379	0.58	0.562	-1.107276 2.037831
-----+-----						

(...)

Illustration 3 – Question 1.1

```
. gen DIVIDA=LEV_LT1>0
```

```
. probit DIVIDA SIZE2 COLLAT2 PROF1 GROWTH2 AGE, vce(cluster id)
```

```
(...)
```

```
Probit regression                                Number of obs    =      30304
                                                Wald chi2(5)      =      1125.30
                                                Prob > chi2       =       0.0000
Log pseudolikelihood = -14363.976              Pseudo R2        =       0.1401
```

(Std. Err. adjusted for 4995 clusters in id)

		Robust				
DIVIDA		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
SIZE2		.3309445	.011001	30.08	0.000	.309383 .3525061
COLLAT2		.754961	.07028	10.74	0.000	.6172147 .8927074
PROF1		-2.114841	.176704	-11.97	0.000	-2.461175 -1.768508
GROWTH2		.000226	.0001309	1.73	0.084	-.0000307 .0004826
AGE		.0011396	.0011065	1.03	0.303	-.001029 .0033082
_cons		-5.510938	.1548431	-35.59	0.000	-5.814425 -5.207451

```
. estimates store DP1
```

Illustration 3 – Question 1.1 (cont.)

```
. glm LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE if LEV_LT>0, family(binomial)
link(logit) vce(cluster id)
(...)
```

```
Log pseudolikelihood = -3187.347033
```

AIC	=	.8776548
BIC	=	-63243.49

(Std. Err. adjusted for 2064 clusters in id)

		Robust					
LEV_LT1		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----							
SIZE2		-.185588	.0140524	-13.21	0.000	-.2131301	-.1580459
COLLAT2		-.7261906	.0998168	-7.28	0.000	-.9218279	-.5305533
PROF1		-1.523694	.2894609	-5.26	0.000	-2.091027	-.9563615
GROWTH2		.0030029	.0004916	6.11	0.000	.0020395	.0039663
AGE		-.0071666	.0013872	-5.17	0.000	-.0098856	-.0044477
_cons		2.389845	.2057352	11.62	0.000	1.986611	2.793078

```
. estimates store DP2
```

Illustration 3 – Question 1.2

```
. tobit LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE, ll(0) vce(cluster id)
(...)
```

(Std. Err. adjusted for 4995 clusters in id)

		Robust				
LEV_LT1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
SIZE2		.1228862	.0040375	30.44	0.000	.1149725 .1308
COLLAT2		.2341081	.0318929	7.34	0.000	.1715967 .2966195
PROF1		-.9789402	.0758608	-12.90	0.000	-1.127631 -.8302498
GROWTH2		.0001581	.0000812	1.95	0.052	-1.13e-06 .0003173
AGE		-.0003276	.0004359	-0.75	0.452	-.001182 .0005267
_cons		-2.064098	.0565249	-36.52	0.000	-2.174889 -1.953306
-----+-----						
/sigma		.4669837	.0080894			.4511282 .4828392

```
Obs. summary:      23027 left-censored observations at LEV_LT1<=0
                  7277  uncensored observations
                  0 right-censored observations
```

```
. estimates store TOBIT
```

Illustration 3 – Question 2.1

Two-part model:

$$E(LEV_LT_i|x_i) = \Phi(x_i'\theta) \frac{e^{x_i'\beta}}{1 + e^{x_i'\beta}}$$

```
. estimates restore DP1  
(results DP1 are active now)
```

```
. scalar xb1 = _b[_cons] + _b[SIZE2]*13.54 + _b[COLLAT2]*0.41 + _b[PROF1]*0.07  
+ _b[GROWTH2]*15.03 + _b[AGE]*19
```

```
. estimates restore DP2  
(results DP2 are active now)
```

```
. scalar xb2 = _b[_cons] + _b[SIZE2]*13.54 + _b[COLLAT2]*0.41 + _b[PROF1]*0.07  
+ _b[GROWTH2]*15.03 + _b[AGE]*19
```

```
. display normal(xb1)*exp(xb2)/(1+exp(xb2))  
.06985232
```

Illustration 3 – Question 2.1 (cont.)

Tobit model:

$$E(LEV_LT_i|x_i) = \Phi\left(\frac{x_i'\beta}{\sigma}\right) x_i'\beta + \sigma\phi\left(\frac{x_i'\beta}{\sigma}\right)$$

```
. estimates restore TOBIT
(results TOBIT are active now)

. scalar xbt = _b[_cons] + _b[SIZE2]*13.54 + _b[COLLAT2]*0.41 + _b[PROF1]*0.07
+ _b[GROWTH2]*15.03 + _b[AGE]*19

. display normal(xbt/_b[/sigma])*xbt+_b[/sigma]*normalden(xbt/_b[/sigma])
.05549759
```

Illustration 3 – Question 2.2

Two-part model:

$$Pr(LEV_LT_i > 0|x_i) = \Phi(x_i'\theta)$$

```
. display normal(xb1)  
.19950111
```

Tobit model:

$$Pr(LEV_LT_i > 0|x_i) = \Phi\left(\frac{x_i'\beta}{\sigma}\right)$$

```
. estimates restore TOBIT  
(results TOBIT are active now)  
  
. display normal(xbt/_b[/sigma])  
.20998609
```

Illustration 3 – Question 2.3

Two-part model:

$$E(LEV_LT_i | x_i, Y_i > 0) = \frac{e^{x_i' \beta}}{1 + e^{x_i' \beta}}$$

```
. display exp(xb2)/(1+exp(xb2))  
.35013498
```

Tobit model:

$$E(Y_i | x_i, Y_i > 0) = x_i' \beta + \sigma \lambda \left(\frac{x_i' \beta}{\sigma} \right)$$

```
. estimates restore TOBIT  
(results TOBIT are active now)  
  
. display xbt+_b[/sigma]*normalden(xbt/_b[/sigma])/normal(xbt/_b[/sigma])  
.26429175
```


Illustration 4

Description of share2:

```
. summarize share2
```

Variable	Obs	Mean	Std. Dev.	Min	Max
share2	2724	.0122429	.024919	0	.1927609

```
. count if share2==0
```

```
1688
```

```
. display 1688/2724
```

```
.61967695
```

Illustration 4 – Question 1

```
. tobit share2 age nadults nkids nkids2 lnx agelnx nadlnx, ll(0)
```

```
Tobit regression                                Number of obs   =       2724
                                                LR chi2(7)      =       170.18
                                                Prob > chi2     =       0.0000
Log likelihood = 758.70053                    Pseudo R2       =      -0.1263
```

share2	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	-.1258528	.0241782	-5.21	0.000	-.1732624	-.0784432
nadults	.01537	.0380475	0.40	0.686	-.0592349	.089975
nkids	.0042697	.0013247	3.22	0.001	.0016723	.0068671
nkids2	-.0099719	.0054713	-1.82	0.068	-.0207002	.0007565
lnx	-.0444314	.0068893	-6.45	0.000	-.0579402	-.0309225
agelnx	.0088221	.0017832	4.95	0.000	.0053256	.0123187
nadlnx	-.0006007	.0027501	-0.22	0.827	-.0059933	.0047918
_cons	.5899797	.0934268	6.31	0.000	.4067849	.7731745
/sigma	.0479951	.0011832			.0456751	.0503151

```
Obs. summary:      1688  left-censored observations at share2<=0
                   1036  uncensored observations
                   0    right-censored observations
```

Illustration 4 – Question 2

```
. probit d2 age nadults nkids nkids2 lnx agelnx nadlnx  
(...)
```

Probit regression

Number of obs = 2724

LR chi2(7) = 101.68

Prob > chi2 = 0.0000

Log likelihood = -1758.5011

Pseudo R2 = 0.0281

d2	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	-2.479476	.5573451	-4.45	0.000	-3.571853	-1.3871
nadults	.6532982	.8664414	0.75	0.451	-1.044896	2.351492
nkids	.0878242	.0305731	2.87	0.004	.027902	.1477463
nkids2	-.2184835	.1224591	-1.78	0.074	-.4584988	.0215319
lnx	-.6162177	.1615998	-3.81	0.000	-.9329476	-.2994879
agelnx	.1732215	.0410448	4.22	0.000	.0927751	.2536678
nadlnx	-.0365886	.0625613	-0.58	0.559	-.1592064	.0860293
_cons	8.077725	2.197003	3.68	0.000	3.771679	12.38377

Illustration 4 – Question 2

```
. reg share2 age nadults nkids nkids2 lnx agelnx nadlnx if share2>0
```

Source	SS	df	MS	Number of obs = 1036		
Model	.15794176	7	.022563109	F(7, 1028) = 26.73		
Residual	.867670386	1028	.000844037	Prob > F = 0.0000		
				R-squared = 0.1540		
				Adj R-squared = 0.1482		
Total	1.02561215	1035	.00099093	Root MSE = .02905		
share2	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	-.0314662	.0205632	-1.53	0.126	-.0718168	.0088844
nadults	-.0130266	.0324149	-0.40	0.688	-.0766334	.0505803
nkids	.0012847	.0010541	1.22	0.223	-.0007837	.0033531
nkids2	-.0034369	.004556	-0.75	0.451	-.0123771	.0055033
lnx	-.0335767	.0054672	-6.14	0.000	-.0443049	-.0228484
agelnx	.0022097	.001516	1.46	0.145	-.000765	.0051844
nadlnx	.0011125	.002345	0.47	0.635	-.003489	.0057141
_cons	.4896596	.0740595	6.61	0.000	.3443345	.6349847

Illustration 4 – Question 3

```
. probit d2 age nadults nkids nkids2 lnx agelnx nadlnx bluecol whitecol
(...)
```

```
Probit regression                                Number of obs   =       2724
                                                LR chi2(9)      =       108.91
                                                Prob > chi2     =       0.0000
Log likelihood =  -1754.886                    Pseudo R2      =       0.0301
```

d2	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	-2.482997	.5596009	-4.44	0.000	-3.579794	-1.386199
nadults	.4852072	.8717378	0.56	0.578	-1.223367	2.193782
nkids	.0812827	.0308297	2.64	0.008	.0208575	.1417078
nkids2	-.2116622	.1230496	-1.72	0.085	-.452835	.0295106
lnx	-.6320797	.1631954	-3.87	0.000	-.9519368	-.3122225
agelnx	.1747353	.0413054	4.23	0.000	.0937781	.2556924
nadlnx	-.0253409	.0629228	-0.40	0.687	-.1486674	.0979856
bluecol	.2064198	.0834317	2.47	0.013	.0428967	.3699428
whitecol	.0215348	.069428	0.31	0.756	-.1145416	.1576111
_cons	8.244464	2.211077	3.73	0.000	3.910833	12.57809

Illustration 4 – Question 3

```
. predict xb, xb
. gen Mills=normalden(xb)/normal(xb)
. reg share2 age nadults nkids nkids2 lnx agelnx nadlnx Mills if share2>0
```

Source	SS	df	MS	Number of obs = 1036		
Model	.158145344	8	.019768168	F(8, 1027)	=	23.40
Residual	.867466802	1027	.000844661	Prob > F	=	0.0000
				R-squared	=	0.1542
				Adj R-squared	=	0.1476
Total	1.02561215	1035	.00099093	Root MSE	=	.02906

share2	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	-.0172991	.0354385	-0.49	0.626	-.0868392	.052241
nadults	-.0174379	.0336488	-0.52	0.604	-.0834661	.0485903
nkids	.0007643	.0014951	0.51	0.609	-.0021695	.0036982
nkids2	-.0020755	.005335	-0.39	0.697	-.0125443	.0083932
lnx	-.0301094	.0089325	-3.37	0.001	-.0476375	-.0125814
agelnx	.0012243	.0025157	0.49	0.627	-.0037122	.0061608
nadlnx	.001365	.0024016	0.57	0.570	-.0033476	.0060775
Mills	-.0090179	.0183686	-0.49	0.624	-.0450622	.0270263
_cons	.4515814	.1072597	4.21	0.000	.2411081	.6620547

Illustration 5 – Question 1

```
. gen y=log( hrrate)>log(3.6)
. gen socd1234=socd1+socd2+socd3+socd4
. gen empmont= empmo/10
. logit y empmont part socd1234 hqd6 married size25 lppltw
(...)
```

```
Logistic regression                                Number of obs   =          4095
                                                    LR chi2(7)      =          421.39
                                                    Prob > chi2     =           0.0000
Log likelihood = -1223.9981                        Pseudo R2       =           0.1469
```

	y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
empmont		.0266874	.0079159	3.37	0.001	.0111725	.0422024
part		-1.178631	.1154354	-10.21	0.000	-1.404881	-.9523821
socd1234		1.142527	.1865971	6.12	0.000	.7768034	1.508251
hqd6		-.6375414	.1134435	-5.62	0.000	-.8598866	-.4151962
married		.3557298	.1093082	3.25	0.001	.1414897	.5699699
size25		1.023301	.1089946	9.39	0.000	.8096759	1.236927
lppltw		-1.127114	.4766336	-2.36	0.018	-2.061298	-.1929288
_cons		1.847127	.1300033	14.21	0.000	1.592325	2.101929

Illustration 5 – Question 1

```
. predict xbLogit,xb
. gen xbLogit2= xbLogit^2
. gen xbLogit3= xbLogit^3

. quietly logit y empmont part socd1234 hqd6 married size25 lppltw xbLogit2
xbLogit3

. test xbLogit2 xbLogit3

( 1) [y]xbLogit2 = 0
( 2) [y]xbLogit3 = 0

             chi2( 2) =      5.72
Prob > chi2 =      0.0574
```


Illustration 5 – Question 1

```
. probit y empmont part socd1234 hqd6 married size25 lppltw  
(...)
```

Probit regression

Number of obs = 4095

LR chi2(7) = 413.14

Prob > chi2 = 0.0000

Log likelihood = -1228.1222

Pseudo R2 = 0.1440

y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
empmont	.0130406	.0038505	3.39	0.001	.0054937	.0205876
part	-.5965585	.0590365	-10.10	0.000	-.7122678	-.4808491
socd1234	.5665749	.088828	6.38	0.000	.3924751	.7406747
hqd6	-.3519214	.061924	-5.68	0.000	-.4732903	-.2305525
married	.1857742	.0584009	3.18	0.001	.0713105	.3002379
size25	.532082	.057272	9.29	0.000	.4198309	.6443332
lppltw	-.6578885	.2947191	-2.23	0.026	-1.235527	-.0802496
_cons	1.066556	.0682845	15.62	0.000	.932721	1.200391

Illustration 5 – Question 1 & 2.1

```
. predict xbProbit,xb

. gen xbProbit2= xbProbit^2

. gen xbProbit3= xbProbit^3

. quietly probit y empmont part socd1234 hqd6 married size25 lppltw xbProbit2
xbProbit3

. test xbProbit2 xbProbit3

( 1) [y]xbProbit2 = 0
( 2) [y]xbProbit3 = 0

             chi2( 2) =    13.60
Prob > chi2 =    0.0011
```

2.1

- Logit model: robust to CB sampling – slope parameter consistent
- Probit model: CB sampling renders standard ML estimators inconsistent
- It is likely that the results of the RESET test reflect these issues

Illustration 5 – Question 2.2

```
. gmm ((.055/.1118*(1-y)+(1-.055)/(1-.1118)*y)
*normalden({beta1}*empmont+{beta2}*part+{beta3}*socd1234+{beta4}
*hqd6+{beta5}*married+{beta6}*lppltw+{beta7}*size25+{beta0}
))* (y-
normprob({beta1}*empmont+{beta2}*part+{beta3}*socd1234+{beta4}
*hqd6+{beta5}*married+{beta6}*lppltw+{beta7}*size25+{beta0}))
)/(normprob({beta1}*empmont+{beta2}*part+{beta3}*socd1234+{beta4}
*hqd6+{beta5}*married+{beta6}*lppltw+{beta7}*size25+{beta0}))
*(1-
normprob({beta1}*empmont+{beta2}*part+{beta3}*socd1234+{beta4}
*hqd6+{beta5}*married+{beta6}*lppltw+{beta7}*size25+{beta0}))
))), instruments(empmont part socd1234 hqd6 married lppltw
size25)
```

Illustration 5 – Question 2.2

GMM estimation

Number of parameters = 8

Number of moments = 8

Initial weight matrix: Unadjusted

Number of obs = 4095

GMM weight matrix: Robust

		Robust				
		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
/beta1		.011869	.0041029	2.89	0.004	.0038274 .0199106
/beta2		-.5463279	.052858	-10.34	0.000	-.6499277 -.4427281
/beta3		.5181078	.0827541	6.26	0.000	.3559127 .6803029
/beta4		-.3183259	.0567416	-5.61	0.000	-.4295373 -.2071144
/beta5		.1730612	.0533727	3.24	0.001	.0684527 .2776697
/beta6		-.6414934	.25315	-2.53	0.011	-1.137658 -.1453286
/beta7		.4859684	.0504227	9.64	0.000	.3871417 .5847952
/beta0		1.458513	.0584018	24.97	0.000	1.344048 1.572978

Instruments for equation 1: empmont part socd1234 hqd6 married lppltw size25
_cons