

Illustration 3 – Question 1

```
. keep if YEAR==1999
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----						
Model	10.6236513	5	2.12473026	F(5, 4686)	=	73.26
Residual	135.914959	4,686	.029004473	Prob > F	=	0.0000
-----+-----						
Total	146.53861	4,691	.031238246	R-squared	=	0.0725
-----+-----						
				Adj R-squared	=	0.0715
				Root MSE	=	.17031

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0227699	.0013931	16.34	0.000	.0200387	.025501
COLLAT	.0356729	.0112018	3.18	0.001	.013712	.0576337
PROF	-.1313503	.0278565	-4.72	0.000	-.1859621	-.0767386
GROWTH	5.98e-07	2.20e-06	0.27	0.786	-3.71e-06	4.91e-06
AGE	-.0003878	.0001676	-2.31	0.021	-.0007165	-.0000591
_cons	-.2246217	.018836	-11.93	0.000	-.2615491	-.1876942
-----+-----						

Illustration 3 – Question 2

Interpretation of partial effects. Ceteris paribus,

- if total of assets increases 1%, the proportion of long term debt increases, in average, $0.023/100=0.0002$.
- each additional year of the firm, reduces the proportion of debt, in average, 0.0004
- if the proportion of collateral increases 0.1 (10pp), the proportion of debt increases, in average, $0.036*0.1=0.0036$ (0.36pp in scale 100%)
- if profitability increases 0.1, the proportion of debt decreases, in average, $0.131*0.1=0.0131$
- if growth increases 0.1, the proportion of debt decreases, in average, $0.000*0.1=0.0000$
 - 3 last effects are response to 0.1 instead of 1, in order to give rise to a meaningful interpretation

Illustration 3 – Question 2

Intuition for effects:

Positive effects on the long term proportion:

- SIZE: larger firms are typically more diversified and have lower probability of failure. Therefore they are issued debt more easily than small firms.
- COLLAT: firms with a higher proportion of collateral are safer in cases of failure. Therefore they are issued debt more easily.
- GROWTH: firms with higher growth are more promising. Therefore they are issued debt more easily.

Negative effects on the long term proportion:

- PROF: more profitable firms are more likely to generate internal funding. Therefore they use less debt.
- AGE: firms that survived for a long time display less debt needs.

Illustration 3 – Question 3

At the 5% significance level all variables, except GROWTH, are statistically individually significant.

At the 5% significance level, all variables are jointly significant

Illustration 3 – Question 4

To test the joint significance of PROF and GROWTH:

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE
```

```
. test PROF GROWTH
```

```
( 1)  PROF = 0
```

```
( 2)  GROWTH = 0
```

```
F( 2, 4686) = 11.12  
Prob > F = 0.0000
```

At the 5% significance level PROF and GROWTH are statistically significant

Illustration 3 – Question 4

Repetition of the test, for illustrative purposes, without the automatic command

```
. regress LEV_LT SIZE COLLAT AGE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(3, 4688)	=	114.19
Model	9.97877613	3	3.32625871	Prob > F	=	0.0000
Residual	136.559834	4,688	.029129657	R-squared	=	0.0681
-----+-----				Adj R-squared	=	0.0675
Total	146.53861	4,691	.031238246	Root MSE	=	.17067

...

```
. display ((0.0725-0.0681)/2/((1-0.0725)/(4692-6)))
```

11.11504

Illustration 3 – Question 5

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(8, 4683)	=	49.76
Model	11.4798615	8	1.43498268	Prob > F	=	0.0000
Residual	135.058749	4,683	.02884022	R-squared	=	0.0783
-----+-----				Adj R-squared	=	0.0768
Total	146.53861	4,691	.031238246	Root MSE	=	.16982

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0301683	.0021456	14.06	0.000	.0259618	.0343747
COLLAT	.0407159	.0113232	3.60	0.000	.0185171	.0629146
PROF	-.1208179	.0278736	-4.33	0.000	-.1754634	-.0661725
GROWTH	2.28e-07	2.19e-06	0.10	0.917	-4.08e-06	4.53e-06
AGE	-.0002906	.0001691	-1.72	0.086	-.000622	.0000409
SE	-.0208001	.0068367	-3.04	0.002	-.0342032	-.0073969
MedE	-.028482	.0100863	-2.82	0.005	-.0482558	-.0087082
LE	-.0873071	.0163264	-5.35	0.000	-.1193145	-.0552997
_cons	-.3107104	.0268095	-11.59	0.000	-.3632698	-.2581511

Illustration 3 – Question 5

```
. test SE MedE LE
```

```
( 1)  SE = 0
```

```
( 2)  MedE = 0
```

```
( 3)  LE = 0
```

```
F( 3, 4683) = 9.90  
Prob > F = 0.0000
```

The new regressors are jointly significant.

Ceteris paribus:

- a small firm, relative to a micro firm, displays in average a long term proportion of debt smaller in 0.021
- a medium firm, relative to a micro firm, displays in average a long term proportion of debt smaller in 0.028
- a large firm, relative to a micro firm, displays in average a long term proportion of debt smaller in 0.087

Illustration 3 – Question 6.1

Model:

$$\begin{aligned} LEV_{LT} &= \beta_0 + \beta_1 SIZE + \beta_2 COLLAT + \beta_3 PROF + \beta_4 GROWTH + \beta_5 AGE \\ &+ \beta_6 (SE * PROF) + \beta_7 (MedE * PROF) + \beta_8 (LE * PROF) + v \end{aligned}$$

Effects of firm profitability:

- Micro firms: β_3
- Small firms: $\beta_3 + \beta_6$
- Medium firms: $\beta_3 + \beta_7$
- Large firms: $\beta_3 + \beta_8$

Null hypotheses:

	Small	Medium	Large
Micro	$\beta_6 = 0$	$\beta_7 = 0$	$\beta_8 = 0$
Small		$\beta_6 = \beta_7$	$\beta_6 = \beta_8$
Medium			$\beta_7 = \beta_8$

Illustration 3 – Question 6.1

```

. gen PROF_SE=PROF*SE
. gen PROF_MedE=PROF*MedE
. gen PROF_LE=PROF*LE

. regress LEV_LT SIZE COLLAT PROF GROWTH AGE PROF_SE PROF_MedE PROF_LE

```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(8, 4683)	=	51.41
Model	11.8295674	8	1.47869593	Prob > F	=	0.0000
Residual	134.709043	4,683	.028765544	R-squared	=	0.0807
-----+-----				Adj R-squared	=	0.0792
Total	146.53861	4,691	.031238246	Root MSE	=	.1696

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0276292	.0015862	17.42	0.000	.0245194	.030739
COLLAT	.0363435	.0111628	3.26	0.001	.0144591	.0582279
PROF	.0095068	.0369042	0.26	0.797	-.0628428	.0818565
GROWTH	-2.68e-07	2.20e-06	-0.12	0.903	-4.57e-06	4.04e-06
AGE	-.0003792	.0001672	-2.27	0.023	-.000707	-.0000515
PROF_SE	-.2026651	.0517991	-3.91	0.000	-.3042157	-.1011145
PROF_MedE	-.343819	.0730633	-4.71	0.000	-.4870576	-.2005805
PROF_LE	-.7133022	.1360224	-5.24	0.000	-.9799701	-.4466343
_cons	-.2892868	.021412	-13.51	0.000	-.3312644	-.2473092

Illustration 3 – Question 6.1

Effects of firm profitability:

- When profitability increases 0.1 (10 pp for percentage), the proportion of long-term debt in the firm's capital structure decreases, on average:
 - 0.00095 (0.095 pp in percentage) - micro firms
 - -0.0193 - small firms → $\text{display_b[PROF]} + \text{b[PROF_SE]}$
 - -0.0334 - medium firms → $\text{display_b[PROF]} + \text{b[PROF_MedE]}$
 - -0.0704 - large firms → $\text{display_b[PROF]} + \text{b[PROF_LE]}$

Illustration 3 – Question 6.1

Profitability effects differ significantly across groups?

```
. test PROF_SE=PROF_MedE
( 1)  PROF_SE - PROF_MedE = 0
      F( 1, 4683) = 3.84
      Prob > F = 0.0502
```

```
. test PROF_SE=PROF_LE
( 1)  PROF_SE - PROF_LE = 0
      F( 1, 4683) = 14.66
      Prob > F = 0.0001
```

```
. test PROF_MedE=PROF_LE
( 1)  PROF_MedE - PROF_LE = 0
      F( 1, 4683) = 7.44
      Prob > F = 0.0064
```

(p-values)	Small	Medium	Large
Micro	0.000***	0.000***	0.000***
Small		0.0502*	0.000***
Medium			0.006***

Illustration 3 – Question 6.2

Model:

$$\begin{aligned} LEV_{LT1} &= \beta_0 + \beta_1 SIZE2 + \beta_2 COLLAT2 + \beta_3 PROF1 + \beta_4 GROWTH2 + \beta_5 AGE \\ &+ \beta_6 LE + \beta_7 (LE * SIZE2) + \beta_8 (LE * COLLAT2) + \beta_9 (LE * PROF1) \\ &+ \beta_{10} (LE * GROWTH2) + \beta_{11} (LE * AGE) + w \end{aligned}$$

Null hypotheses:

- $H_0: \beta_6 = \beta_7 = \beta_8 = \beta_9 = \beta_{10} = \beta_{11} = 0$ (no structural break)

```
. gen SIZE_LE=SIZE*LE
. gen COLLAT_LE=COLLAT*LE
. gen GROWTH_LE=GROWTH*LE
. gen AGE_LE=AGE*LE
```

Illustration 3 – Question 6.2

```
regress LEV_LT SIZE COLLAT PROF GROWTH AGE LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(11, 4680)	=	36.14
Model	11.4734211	11	1.04303829	Prob > F	=	0.0000
Residual	135.065189	4,680	.028860083	R-squared	=	0.0783
-----+-----				Adj R-squared	=	0.0761
Total	146.53861	4,691	.031238246	Root MSE	=	.16988

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0262108	.0015979	16.40	0.000	.0230782	.0293433
COLLAT	.0298537	.0114426	2.61	0.009	.0074207	.0522867
PROF	-.1178161	.0283403	-4.16	0.000	-.1733765	-.0622557
GROWTH	4.15e-07	2.19e-06	0.19	0.850	-3.89e-06	4.72e-06
AGE	-.0003768	.0001822	-2.07	0.039	-.0007339	-.0000196
LE	-.0570063	.1411598	-0.40	0.686	-.3337461	.2197334
SIZE_LE	-.002168	.0082257	-0.26	0.792	-.0182943	.0139582
COLLAT_LE	.1055932	.053384	1.98	0.048	.0009355	.2102509
PROF_LE	-.2771129	.1595962	-1.74	0.083	-.5899966	.0357708
GROWTH_LE	.0010892	.0006263	1.74	0.082	-.0001386	.002317
AGE_LE	.0003647	.000469	0.78	0.437	-.0005547	.0012842
_cons	-.2680171	.0213989	-12.52	0.000	-.3099691	-.2260651

Illustration 3 – Question 6.2

```
. test LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE
```

- (1) LE = 0
- (2) SIZE_LE = 0
- (3) COLLAT_LE = 0
- (4) PROF_LE = 0
- (5) GROWTH_LE = 0
- (6) AGE_LE = 0

```
F( 6, 4680) = 4.91  
Prob > F = 0.0001
```

The null hypothesis is rejected, which implies that it is better to estimate separate models for each size-based group of firms or, equivalently, the model of the previous page, since:

(see the next page and compare with the interactions model)

Illustration 3 – Question 6.2

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE if LE==0  
(...)
```

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE	.0262108	.0016006	16.38	0.000	.0230728	.0293487
COLLAT	.0298537	.011462	2.60	0.009	.0073825	.0523249
PROF	-.1178161	.0283882	-4.15	0.000	-.1734712	-.062161
GROWTH	4.15e-07	2.20e-06	0.19	0.850	-3.89e-06	4.72e-06
AGE	-.0003768	.0001825	-2.06	0.039	-.0007345	-.000019
_cons	-.2680171	.0214351	-12.50	0.000	-.3100405	-.2259936

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE if LE==1  
(...)
```

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE	.0240427	.0078385	3.07	0.002	.008609	.0394765
COLLAT	.1354469	.0506539	2.67	0.008	.0357116	.2351822
PROF	-.394929	.1525738	-2.59	0.010	-.6953401	-.0945179
GROWTH	.0010896	.0006084	1.79	0.074	-.0001082	.0022875
AGE	-.000012	.0004198	-0.03	0.977	-.0008387	.0008146
_cons	-.3250234	.1355432	-2.40	0.017	-.591902	-.0581448

Illustration 3 – Question 7.1 & 7.2

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE
```

```
. ovtest
```

```
Ramsey RESET test using powers of the fitted values of LEV_LT
```

```
Ho: model has no omitted variables
```

```
F(3, 4680) = 10.62
```

```
Prob > F = 0.0000 → Unsuitable model functional form
```

```
. estat hettest, rhs fstat
```

```
Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
```

```
Ho: Constant variance
```

```
Variables: SIZE COLLAT PROF GROWTH AGE SE MedE LE
```

```
F(8 , 4683) = 20.76
```

```
Prob > F = 0.0000 → Heteroskedastic errors
```

Illustration 3 – Question 7.1 & 7.2

RESET test – manual implementation

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE
. predict XB
. gen XB2=XB^2
. gen XB3=XB^3
. gen XB4=XB^4
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE XB2 XB3 XB4

. test XB2 XB3 XB4

( 1)  XB2 = 0
( 2)  XB3 = 0
( 3)  XB4 = 0

F( 3, 4680) = 10.62
Prob > F = 0.0000
```

Illustration 3 – Question 7.1 & 7.2

BP test – manual implementation

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE

. predict uhat, resid

. gen uhat2=uhat^2

. quietly regress uhat2 SIZE COLLAT PROF GROWTH AGE SE MedE LE

. test SIZE COLLAT PROF GROWTH AGE SE MedE LE
( 1)  SIZE = 0
( 2)  COLLAT = 0
( 3)  PROF = 0
( 4)  GROWTH = 0
( 5)  AGE = 0
( 6)  SE = 0
( 7)  MedE = 0
( 8)  LE = 0

      F(  8,  4683) =    20.76
      Prob > F =    0.0000
```

Illustration 3 – Question 8

```
regress LEV_LT SIZE COLLAT PROF GROWTH AGE LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE,
robust
```

Linear regression

```
Number of obs      =      4,692
F(11, 4680)        =      47.01
Prob > F           =      0.0000
R-squared          =      0.0783
Root MSE          =      .16988
```

		Robust				[95% Conf. Interval]	
LEV_LT	Coef.	Std. Err.	t	P> t			
SIZE	.0262108	.0014557	18.01	0.000	.0233569	.0290646	
COLLAT	.0298537	.0125967	2.37	0.018	.0051583	.0545491	
PROF	-.1178161	.0215441	-5.47	0.000	-.1600527	-.0755795	
GROWTH	4.15e-07	1.06e-06	0.39	0.694	-1.65e-06	2.48e-06	
AGE	-.0003768	.0001775	-2.12	0.034	-.0007247	-.0000289	
LE	-.0570063	.1558992	-0.37	0.715	-.3626422	.2486295	
SIZE_LE	-.002168	.0091113	-0.24	0.812	-.0200304	.0156944	
COLLAT_LE	.1055932	.0590662	1.79	0.074	-.0102043	.2213907	
PROF_LE	-.2771129	.1219132	-2.27	0.023	-.5161203	-.0381055	
GROWTH_LE	.0010892	.0007304	1.49	0.136	-.0003428	.0025212	
AGE_LE	.0003647	.0004279	0.85	0.394	-.0004742	.0012037	
_cons	-.2680171	.0173169	-15.48	0.000	-.3019663	-.2340678	

Stata – Panel data

Commands for panel data:

- Prefix **xt**

Initial command:

- **xtset** *identvar tempvar*

Examples:

- **xtdescribe** [*varlist*]
- **xtsum** [*varlist*]
- **xttab** *varname*

Illustration 4

Model:

$$\begin{aligned} & \text{Log}(Wage) \\ & = \beta_0 + \beta_1 \text{Schooling} + \beta_2 \text{Exper} + \beta_3 \text{Exper}^2 + \beta_4 \text{Black} \\ & + \beta_5 \text{Union} + \beta_6 \text{South} + \beta_7 \text{Public} + u \end{aligned}$$

Aim:

- Estimating β_5 (impact of collective bargaining on wages)

Sample:

- 545 full-time working males who completed their schooling by 1980 and were observed over the period 1980-1987

Details:

- Verbeek (2008), pp. 375-377

Illustration 4 – Question 1

```
. xtset NR YEAR
      panel variable:  NR (strongly balanced)
      time variable:  YEAR, 1980 to 1987
      delta: 1 unit
```

```
. xtdescribe
      NR: 13, 17, ..., 12548          n =          545
      YEAR: 1980, 1981, ..., 1987    T =           8
      Delta(YEAR) = 1 unit
      Span(YEAR) = 8 periods
      (NR*YEAR uniquely identifies each observation)
```

```
Distribution of T_i:  min      5%      25%      50%      75%      95%      max
                    8        8        8        8        8        8        8
```

```
      Freq.  Percent  Cum. | Pattern
-----+-----
      545    100.00  100.00 | 11111111
-----+-----
      545    100.00      | XXXXXXXXX
```

• Balanced Panel

Illustration 4 – Question 2

```
. xtsum WAGE SCHOOLING EXPER BLACK UNION SOUTH PUBLIC
```

Variable		Mean	Std. Dev.	Min	Max	Observations
WAGE	overall	5.919175	3.202225	.0279014	57.50431	N = 4360
	between		2.455819	1.503564	28.35696	n = 545
	within		2.057397	-16.42349	35.06652	T = 8
SCHOOL~G	overall	11.76697	1.746181	3	16	N = 4360
	between		1.747585	3	16	n = 545
	within		0	11.76697	11.76697	T = 8
EXPER	overall	6.514679	2.825873	0	18	N = 4360
	between		1.654918	3.5	14.5	n = 545
	within		2.291551	3.014679	10.01468	T = 8
BLACK	overall	.1155963	.3197769	0	1	N = 4360
	between		.320034	0	1	n = 545
	within		0	.1155963	.1155963	T = 8

Illustration 4 – Question 2 (cont.)

```
. xtsum WAGE SCHOOLING EXPER BLACK UNION SOUTH PUBLIC
```

Variable		Mean	Std. Dev.	Min	Max	Observations
UNION	overall	.2440367	.4295639	0	1	N = 4360
	between		.3294467	0	1	n = 545
	within		.2759787	-.6309633	1.119037	T = 8
SOUTH	overall	.3506881	.4772402	0	1	N = 4360
	between		.4644838	0	1	n = 545
	within		.1111732	-.5243119	1.225688	T = 8
PUBLIC	overall	.0401376	.1963044	0	1	N = 4360
	between		.1388064	0	1	n = 545
	within		.1389214	-.8348624	.9151376	T = 8

Illustration 4 – Question 3

```
. xttab UNION
```

UNION	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	3296	75.60	511	93.76	80.63
1	1064	24.40	280	51.38	47.50
Total	4360	100.00	791	145.14	68.90

(n = 545)

- Overall, 24,4% of wages resulted from collective bargaining
- For 51,38% of workers (280 out of 545), in at least one year their wages resulted from collective bargaining
- For 47,5% of those 280 workers, wages resulted always from collective bargaining

Illustration 4 – Question 3 (cont.)

```
. xttab SOUTH
```

SOUTH	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	2831	64.93	374	68.62	94.62
1	1529	35.07	212	38.90	90.15
Total	4360	100.00	586	107.52	93.00

(n = 545)

- Overall, 35,1% of observations concern cases of individuals living in the South
- 38,9% of workers (212 out of 545) lived at least one year in the South
- 90,15% of those 212 workers lived always in the South

Illustration 4 – Question 3 (cont.)

```
. xttab PUBLIC
```

PUBLIC	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	4185	95.99	543	99.63	96.34
1	175	4.01	62	11.38	35.28
Total	4360	100.00	605	111.01	90.08

(n = 545)

- Overall, 4,01 % of observations concern cases of individuals working in Public Administration
- 11,38% of workers (62 out of 545) worked at least one year in Public Administration
- 35,28% of those 62 workers worked always in Public Administration

Illustration 4 – Question 4.1 - Pooled

```
. gen LWAGE=log(WAGE)
. regress LWAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, vce(cluster NR)
```

Linear regression

```
Number of obs = 4360
F( 7, 544) = 56.52
Prob > F = 0.0000
R-squared = 0.1792
Root MSE = .48292
```

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
SCHOOLING	.1012109	.0090121	11.23	0.000	.0835082	.1189137
EXPER	.1007275	.0120267	8.38	0.000	.0771029	.124352
EXPER2	-.003228	.0008458	-3.82	0.000	-.0048895	-.0015665
BLACK	-.1602247	.0488069	-3.28	0.001	-.2560979	-.0643516
UNION	.1842702	.0279656	6.59	0.000	.1293365	.239204
SOUTH	-.0437693	.0328437	-1.33	0.183	-.1082852	.0207466
PUBLIC	.0105751	.0486814	0.22	0.828	-.0850514	.1062016
_cons	-.0467585	.1147854	-0.41	0.684	-.2722354	.1787184

```
. estimates store POOLED
```

Illustration 4 – Question 4.1 – Random Effects

```
. xtreg LWAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, vce(cluster NR)
```

```
Random-effects GLS regression                Number of obs      =       4360
Group variable: NR                          Number of groups   =       545

R-sq:   within  = 0.1764                    Obs per group: min =         8
        between = 0.1711                    avg   =       8.0
        overall  = 0.1735                    max   =         8

                                           Wald chi2(7)       =       569.04
corr(u_i, X)  = 0 (assumed)                 Prob > chi2        =       0.0000
```

(continues in the next slide)

Illustration 4 – Question 4.1 – Random Effects (cont.)

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
LWAGE						
SCHOOLING	.1019444	.0086285	11.81	0.000	.0850329	.1188558
EXPER	.1187674	.010413	11.41	0.000	.0983582	.1391766
EXPER2	-.0043029	.000674	-6.38	0.000	-.0056239	-.0029818
BLACK	-.160817	.049919	-3.22	0.001	-.2586564	-.0629776
UNION	.1083223	.0210233	5.15	0.000	.0671174	.1495272
SOUTH	-.0040207	.0317309	-0.13	0.899	-.0662121	.0581707
PUBLIC	.0326192	.0335003	0.97	0.330	-.0330401	.0982784
_cons	-.1149344	.1107496	-1.04	0.299	-.3319997	.1021309
sigma_u	.3278637					
sigma_e	.35132742					
rho	.46549462	(fraction of variance due to u_i)				

. estimates store RE

Illustration 4 – Question 4.1 – Fixed Effects

```
. xtreg LWAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, fe vce(cluster NR)
```

```
note: SCHOOLING omitted because of collinearity
```

```
note: BLACK omitted because of collinearity
```

```
Fixed-effects (within) regression
```

```
Group variable: NR
```

```
Number of obs      =      4360
```

```
Number of groups   =      545
```

```
R-sq:  within  = 0.1779
```

```
        between = 0.0039
```

```
        overall = 0.0435
```

```
Obs per group: min =      8
```

```
                avg =     8.0
```

```
                max =      8
```

```
corr(u_i, Xb) = -0.1837
```

```
F(5, 544) =      87.62
```

```
Prob > F =      0.0000
```

(continues in the next slide)

Illustration 4 – Question 4.1 – Fixed Effects (cont.)

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
LWAGE						
SCHOOLING	0	(omitted)				
EXPER	.1206647	.0105609	11.43	0.000	.0999194	.1414099
EXPER2	-.004429	.0006836	-6.48	0.000	-.0057717	-.0030863
BLACK	0	(omitted)				
UNION	.0833501	.0229404	3.63	0.000	.0382875	.1284127
SOUTH	.1038881	.0657487	1.58	0.115	-.0252644	.2330406
PUBLIC	.0363333	.0373874	0.97	0.332	-.0371081	.1097748
_cons	1.028155	.0425644	24.16	0.000	.9445442	1.111765
sigma_u	.41161426					
sigma_e	.35132742					
rho	.57852861	(fraction of variance due to u_i)				

Ignore

. estimates store FE

Illustration 4 – Question 4.1 – Table

. estimates table POOLED RE FE, b se

Variable	POOLED	RE	FE
SCHOOLING	.10121091	.10194436	(omitted)
	.00901209	.00862846	
EXPER	.10072746	.11876742	.12066465
	.01202673	.01041304	.01056093
EXPER2	-.00322797	-.00430286	-.00442898
	.00084583	.00067402	.00068355
BLACK	-.16022473	-.16081702	(omitted)
	.04880694	.04991897	
UNION	.18427022	.1083223	.08335012
	.02796559	.02102328	.02294039
SOUTH	-.04376929	-.00402067	.1038881
	.03284366	.03173089	.06574874
PUBLIC	.01057512	.03261915	.03633333
	.04868137	.03350025	.03738745
_cons	-.0467585	-.1149344	1.0281548
	.1147854	.11074962	.04256437

legend: b/se

Illustration 4 – Question 4.1 – Table

```
. estimates table POOLED RE FE, b star(0.1 0.05 0.01)
```

Variable	POOLED	RE	FE
SCHOOLING	.10121091***	.10194436***	(omitted)
EXPER	.10072746***	.11876742***	.12066465***
EXPER2	-.00322797***	-.00430286***	-.00442898***
BLACK	-.16022473***	-.16081702***	(omitted)
UNION	.18427022***	.1083223***	.08335012***
SOUTH	-.04376929	-.00402067	.1038881
PUBLIC	.01057512	.03261915	.03633333
_cons	-.0467585	-.1149344	1.0281548***

legend: * p<.1; ** p<.05; *** p<.01

Illustration 4 – Question 4.2

```
. quietly xtreg LWAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, fe
. estimates store FEh
. quietly xtreg LWAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC
. estimates store REh
. hausman FEh REh
```

	---- Coefficients ----			
	(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
	FEh	REh	Difference	S.E.
EXPER	.1206647	.1187674	.0018972	.0014118
EXPER2	-.004429	-.0043029	-.0001261	.0001223
UNION	.0833501	.1083223	-.0249722	.0072895
SOUTH	.1038881	-.0040207	.1079088	.0395125
PUBLIC	.0363333	.0326192	.0037142	.0125681

b = consistent under Ho and Ha; obtained from xtreg
 B = inconsistent under Ha, efficient under Ho; obtained from xtreg

Test: Ho: difference in coefficients not systematic

```
chi2(5) = (b-B)'[(V_b-V_B)^(-1)](b-B)
          = 34.93
Prob>chi2 = 0.0000
```

→ The hypothesis of random effects is rejected.

Illustration 4 – Question 4.3

```
. regress D.LWAGE D.SCHOOLING D.EXPER D.EXPER2 D.BLACK D.UNION D.SOUTH D.PUBLIC, vce(  
cluster NR) nocons  
note: _delete omitted because of collinearity  
note: _delete omitted because of collinearity
```

Linear regression

```
Number of obs =    3815  
F( 5,    544) =    76.05  
Prob > F      =    0.0000  
R-squared     =    0.0269  
Root MSE     =    .44307
```

(continues in the next slide)

Illustration 4 – Question 4.3

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
D.LWAGE						
SCHOOLING						
D1.	0	(omitted)				
EXPER						
D1.	.119242	.0143533	8.31	0.000	.0910474	.1474366
EXPER2						
D1.	-.0040136	.0009438	-4.25	0.000	-.0058676	-.0021596
BLACK						
D1.	0	(omitted)				
UNION						
D1.	.0436369	.022141	1.97	0.049	.0001446	.0871292
SOUTH						
D1.	.0811916	.0782312	1.04	0.300	-.0724807	.2348638
PUBLIC						
D1.	.0425582	.0356367	1.19	0.233	-.0274442	.1125606

Illustration 4 – Question 4.4

```
. xtreg LWAGE SCHOOLING EXPER EXPER2 c.BLACK##i.YEAR UNION SOUTH PUBLIC, re vce(cluster NR)
(...)

```

Search help for
fvvarlist

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
SCHOOLING	.0928925	.0109548	8.48	0.000	.0714214	.1143636
EXPER	.1105721	.016444	6.72	0.000	.0783424	.1428018
EXPER2	-.0048039	.0007875	-6.10	0.000	-.0063473	-.0032605
BLACK	-.0888817	.0714913	-1.24	0.214	-.229002	.0512386
YEAR						
1981	.0396178	.0289286	1.37	0.171	-.0170812	.0963169
1982	.0461313	.0350288	1.32	0.188	-.0225238	.1147864
1983	.0321001	.0442559	0.73	0.468	-.0546398	.11884
1984	.0526727	.0565513	0.93	0.352	-.0581658	.1635111
1985	.0778807	.0654343	1.19	0.234	-.0503683	.2061296
1986	.0969979	.0753684	1.29	0.198	-.0507214	.2447171
1987	.1510222	.0856029	1.76	0.078	-.0167565	.3188008

(continues in the next slide)

Illustration 4 – Question 4.4 (cont.)

```

-----
                |
                |           Robust
LWAGE |           Coef.   Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
YEAR#c.BLACK |
1981 |      .0272047   .0719625     0.38   0.705   - .1138393   .1682486
1982 |     -.1121196   .0918991    -1.22   0.222   - .2922386   .0679993
1983 |     -.0685286   .0846309    -0.81   0.418   - .2344021   .097345
1984 |     -.0541201   .0818181    -0.66   0.508   - .2144806   .1062405
1985 |     -.1522521   .0853343    -1.78   0.074   - .3195042   .0150001
1986 |     -.0376476   .0877731    -0.43   0.668   - .2096796   .1343844
1987 |     -.1414707   .0807227    -1.75   0.080   - .2996843   .016743
                |
UNION |      .1084207   .0208784     5.19   0.000    .0674999   .1493416
SOUTH |     -.0061861   .0317151    -0.20   0.845   - .0683466   .0559745
PUBLIC |      .0311455   .033775     0.92   0.356   - .0350523   .0973433
_cons |      .008449    .1567611     0.05   0.957   - .2987971   .3156951
-----+-----
sigma_u |      .32792944
sigma_e |      .3508362
rho |      .46629073   (fraction of variance due to u_i)
-----

```


Illustration 5 – Model

- Trade-Off theory – optimal / target proportion of debt:

$$MDR_{it}^* = x'_{i,t-1}\beta + \eta_{it}$$

- Target adjustment model:

$$MDR_{it} - MDR_{i,t-1} = (1 - \gamma)(MDR_{it}^* - MDR_{i,t-1})$$

- $0 \leq \gamma \leq 1$
- Adjustment speed: $1 - \gamma$
 - $\gamma = 0 \Rightarrow$ Firms adjust immediately and completely
 - $\gamma = 1 \Rightarrow$ No adjustment

- Econometric model:

$$MDR_{it} = MDR_{i,t-1} + (1 - \gamma)(x'_{i,t-1}\beta + \eta_{it} - MDR_{i,t-1})$$
$$\vdots$$
$$MDR_{it} = \gamma MDR_{i,t-1} + x'_{i,t-1}\theta + \alpha_i + u_{it}$$

Illustration 5 – Model (cont.)

Aim:

- Estimate γ
- Test whether $H_0: \gamma = 1$ (trade-off theory not valid)

Sample:

- 5449 firms observed over the period 1986-2001
- Unbalanced panel
- Source: Compustat Industrial Annual Tapes

Details:

- Verbeek (2008), pp. 383-388

```
. xtset gvkey yeara
      panel variable:  gvkey (unbalanced)
      time variable:  yeara, 1986 to 2001, but with gaps
                   delta: 1 unit
```

Illustration 5 – Question 1

```
. xtdescribe
   gvkey: 1003, 1004, ..., 233397          n =          5449
   yeara: 1986, 1987, ..., 2001          T =           16
           Delta(yeara) = 1 unit
           Span(yeara)  = 16 periods
           (gvkey*yeara uniquely identifies each observation)
Distribution of T_i:  min      5%      25%      50%      75%      95%      max
                    1        1        1        3        8        14        16
   Freq.  Percent  Cum.  |  Pattern
-----+-----
   384    7.05    7.05 | .....1
   176    3.23   10.28 | 1111111111111111
   167    3.06   13.34 | 1.....
   149    2.73   16.08 | .....1.
   136    2.50   18.57 | 11.....
   132    2.42   20.99 | .....1...
   131    2.40   23.40 | .....1..
   113    2.07   25.47 | 111.....
    98    1.80   27.27 | .....1....
  3963   72.73  100.00 | (other patterns)
-----+-----
  5449  100.00          | XXXXXXXXXXXXXXXXX
```

Illustration 5 – Question 2

```
. quietly regress mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, vce(cluster gvkey)
```

```
. estimates store POOLED
```

```
. quietly xtreg mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, vce(cluster gvkey)
```

```
. estimates store RE
```

```
. quietly xtreg mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, fe vce(cluster gvkey)
```

```
. estimates store FE
```

Illustration 5 – Question 2 (cont.)

```
. estimates table POOLED RE FE, b star
```

Variable	POOLED	RE	FE
mdr			
L1	.8835036***	.78734111***	.53498254***
ebit_ta	-.03233775***	-.03455703***	-.05003294***
mb	.0016432*	.00069052	.00227756*
dep_ta	-.26051795***	-.30137705***	-.12395444
lnta	-.00067042	.00183328*	.03803015***
fa_ta	.02012146***	.02919991***	.05934357***
rd_dum	.00688957**	.0096552***	.00005977
rd_ta	-.12020508***	-.14520778***	-.06567621*
indmedian	.03212249**	.06029162***	.16721793***
rated	.00713406*	.01103821**	.02058981***
_cons	.05818177***	.03798629**	-.60083475***

legend: * p<0.05; ** p<0.01; *** p<0.001

Illustration 5 – Question 3.1

```
. xtivreg mdr (L.mdr=L2.mdr) ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta
indmedian rated, fd
(...)
```

D.mdr	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
LD.	7.03303	5.494343	1.28	0.201	-3.735684	17.80174
ebit_ta						
D1.	1.207597	.9705551	1.24	0.213	-.6946564	3.10985
mb						
D1.	.244267	.1853757	1.32	0.188	-.1190627	.6075966
dep_ta						
D1.	-1.858345	1.577202	-1.18	0.239	-4.949603	1.232914
lnta						
D1.	-.5214084	.4557998	-1.14	0.253	-1.41476	.3719429

(...)

```
Instrumented:  L.mdr
Instruments:  ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated
               L2.mdr
```

Illustration 5 – Question 3.2

```
. xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep vce(robust)
```

```
Arellano-Bond dynamic panel-data estimation   Number of obs       =       15039  
Group variable: gvkey                         Number of groups    =        2996  
Time variable: yeara  
  
Obs per group:   min =           1  
                 avg =   5.019693  
                 max =           14  
  
Number of instruments =       115           Wald chi2(10)       =       212.72  
                                     Prob > chi2         =       0.0000
```

Two-step results

(continues in the next slide)

Illustration 5 – Question 3.2 (cont.)

(Std. Err. adjusted for clustering on gvkey)

	Coef.	WC-Robust Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
L1.	.3819695	.0731919	5.22	0.000	.238516	.525423
ebit_ta	.035684	.0173522	2.06	0.040	.0016744	.0696936
mb	.0147128	.0027217	5.41	0.000	.0093784	.0200472
dep_ta	.0648811	.109432	0.59	0.553	-.1496016	.2793639
lnta	.030107	.0083243	3.62	0.000	.0137916	.0464224
fa_ta	.0150317	.0286987	0.52	0.600	-.0412168	.0712801
rd_dum	-.0178784	.0122991	-1.45	0.146	-.0419841	.0062273
rd_ta	.001471	.035963	0.04	0.967	-.0690151	.0719571
indmedian	.0919917	.0453824	2.03	0.043	.0030439	.1809395
rated	-.0066174	.0099448	-0.67	0.506	-.0261088	.0128741
_cons	-.4235622	.1467658	-2.89	0.004	-.7112179	-.1359065

Instruments for differenced equation

GMM-type: L(2/.) .mdr

Standard: D.ebit_ta D.mb D.dep_ta D.lnta D.fa_ta D.rd_dum D.rd_ta D.indmedian
D.rated

Instruments for level equation

Standard: _cons

Illustration 5 – Question 3.3

```
. xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep maxldep(2) vce(robust)
```

```
Arellano-Bond dynamic panel-data estimation   Number of obs       =       15039  
Group variable: gvkey                         Number of groups    =        2996  
Time variable: yeara  
  
Obs per group:   min =           1  
                 avg =   5.019693  
                 max =           14  
  
Number of instruments =          37           Wald chi2(10)       =       179.25  
                                     Prob > chi2         =       0.0000
```

Two-step results

(continues in the next slide)

Illustration 5 – Question 3.3 (cont.)

(Std. Err. adjusted for clustering on gvkey)

	Coef.	WC-Robust Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
L1.	.4005973	.0886448	4.52	0.000	.2268567	.574338
ebit_ta	.0445636	.0197946	2.25	0.024	.005767	.0833603
mb	.0165177	.003106	5.32	0.000	.0104301	.0226053
dep_ta	.0832717	.1094867	0.76	0.447	-.1313182	.2978616
lnta	.0252836	.0090865	2.78	0.005	.0074744	.0430928
fa_ta	.0065411	.0306739	0.21	0.831	-.0535787	.0666608
rd_dum	-.0135049	.0122092	-1.11	0.269	-.0374345	.0104248
rd_ta	.0015315	.0361632	0.04	0.966	-.0693471	.07241
indmedian	.0670826	.0493198	1.36	0.174	-.0295825	.1637477
rated	-.0098843	.0101263	-0.98	0.329	-.0297315	.009963
_cons	-.3367156	.1575827	-2.14	0.033	-.6455719	-.0278593

Instruments for differenced equation

GMM-type: L(2/3).mdr

Standard: D.ebit_ta D.mb D.dep_ta D.lnta D.fa_ta D.rd_dum D.rd_ta
D.indmedian D.rated

Instruments for level equation

Standard: _cons

Illustration 5 – Question 3.4

```
. xtdpdsys mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep vce(robust)
```

```
System dynamic panel-data estimation      Number of obs      =      19573  
Group variable: gvkey                    Number of groups   =      3777  
Time variable: yeara  
  
Obs per group:      min =           1  
                   avg =    5.182155  
                   max =           15  
  
Number of instruments =      129          Wald chi2(10)     =    2185.15  
                   Prob > chi2       =      0.0000  
  
Two-step results
```

(continues in the next slide)

Illustration 5 – Question 3.4 (cont.)

```

-----
                |
                |          WC-Robust
mdr |          Coef.   Std. Err.   z   P>|z|   [95% Conf. Interval]
-----+-----
mdr |
L1. |    1.030104   .0238744   43.15   0.000   .9833112   1.076897
    |
ebit_ta |    .1285762   .018583   6.92   0.000   .0921541   .1649983
mb |    .0312589   .0027109   11.53   0.000   .0259457   .0365722
dep_ta |   -.1036924   .1374032   -0.75   0.450   -.3729978   .165613
lnta |   -.0127849   .0068955   -1.85   0.064   -.0262999   .0007301
fa_ta |   -.1085659   .0327925   -3.31   0.001   -.172838   -.0442939
rd_dum |   -.0061031   .0153821   -0.40   0.692   -.0362514   .0240452
rd_ta |    .0759739   .0441006   1.72   0.085   -.0104617   .1624096
indmedian |  -.2372958   .0400293   -5.93   0.000   -.3157518   -.1588397
rated |   -.0302876   .0120363   -2.52   0.012   -.0538783   -.006697
_cons |    .273204   .1269921   2.15   0.031   .0243041   .522104
-----

```

Instruments for differenced equation

GMM-type: L(2/.)mdr

Standard: D.ebit_ta D.mb D.dep_ta D.lnta D.fa_ta D.rd_dum D.rd_ta D.indmedian
D.rated

Instruments for level equation

GMM-type: LD.mdr

Standard: _cons

Illustration 5 – Question 4.1

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep vce(robust)
```

```
. estat abond, artests(3)
```

(...)

Arellano-Bond test for zero autocorrelation in first-differenced errors

```
+-----+  
|Order | z      Prob > z|  
+-----+-----+  
|  1  | -6.206  0.0000 |  
|  2  | -3.587  0.0003 |  
|  3  | -3.3325 0.0009 |  
+-----+  
H0: no autocorrelation
```

There is autocorrelation of order higher than 1: the estimators are not consistent.

Illustration 5 – Question 4.2

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep
```

```
. estat sargan
```

Sargan test of overidentifying restrictions

H0: overidentifying restrictions are valid

```
chi2(104)      = 436.3939
```

```
Prob > chi2    = 0.0000
```

The hypothesis of exogenous
instruments is rejected: the estimators
are not consistent.

Illustration 5 – Question 4.3

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep vce(robust)
```

```
. test L.mdr=1
```

```
( 1)  L.mdr = 1
```

```
      chi2( 1) =    71.30  
Prob > chi2 =    0.0000
```

The hypothesis that $\gamma = 1$ is rejected, which implies that firms seem to adjust their debt ratios: there is no evidence against the trade-off theory.