

# Illustration 3 – Question 1

```
. keep if YEAR==1999
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(5, 4686)	=	73.26
Model	10.6236513	5	2.12473026	Prob > F	=	0.0000
Residual	135.914959	4,686	.029004473	R-squared	=	0.0725
-----+-----				Adj R-squared	=	0.0715
Total	146.53861	4,691	.031238246	Root MSE	=	.17031

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0227699	.0013931	16.34	0.000	.0200387	.025501
COLLAT	.0356729	.0112018	3.18	0.001	.013712	.0576337
PROF	-.1313503	.0278565	-4.72	0.000	-.1859621	-.0767386
GROWTH	5.98e-07	2.20e-06	0.27	0.786	-3.71e-06	4.91e-06
AGE	-.0003878	.0001676	-2.31	0.021	-.0007165	-.0000591
_cons	-.2246217	.018836	-11.93	0.000	-.2615491	-.1876942
-----+-----						

$$\widehat{LEV\_LT}_i = -0.225 + 0.023SIZE_i + 0.036COLLAT_i - 0.131PROF_i + 0.000GROWTH_i + 0.000AGE_i$$

# Illustration 3 – Question 2

Interpretation of partial effects. Ceteris paribus,

- if total of assets increases 1%, the proportion of long term debt increases, in average,  $0.023/100=0.0002$ .
- each additional year of the firm, reduces the proportion of debt, in average, 0.0004
- if the proportion of collateral increases 0.1 (10pp), the proportion of debt increases, in average,  $0.036*0.1=0.0036$  (0.36pp in scale 100%)
- if profitability increases 0.1, the proportion of debt decreases, in average,  $0.131*0.1=0.0131$
- if growth increases 0.1, the proportion of debt decreases, in average,  $0.000*0.1=0.0000$ 
  - 3 last effects are response to 0.1 instead of 1, in order to give rise to a meaningful interpretation

# Illustration 3 – Question 2

Intuition for effects:

Positive effects on the long term proportion:

- SIZE: larger firms are typically more diversified and have lower probability of failure. Therefore they are issued debt more easily than small firms.
- COLLAT: firms with a higher proportion of collateral are safer in cases of failure. Therefore they are issued debt more easily.
- GROWTH: firms with higher growth are more promising. Therefore they are issued debt more easily.

Negative effects on the long term proportion:

- PROF: more profitable firms are more likely to generate internal funding. Therefore they use less debt.
- AGE: firms that survived for a long time display less debt needs.

## Illustration 3 – Question 3

At the 5% significance level all variables, except GROWTH, are statistically individually significant.

At the 5% significance level, all variables are jointly significant

# Illustration 3 – Question 4

To test the joint significance of PROF and GROWTH:

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE
```

```
. test PROF GROWTH
```

```
( 1)  PROF = 0
```

```
( 2)  GROWTH = 0
```

```
F( 2, 4686) = 11.12  
Prob > F = 0.0000
```

At the 5% significance level PROF and GROWTH are statistically significant

# Illustration 3 – Question 4

Repetition of the test, for illustrative purposes, without the automatic command

```
. regress LEV_LT SIZE COLLAT AGE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(3, 4688)	=	114.19
Model	9.97877613	3	3.32625871	Prob > F	=	0.0000
Residual	136.559834	4,688	.029129657	R-squared	=	0.0681
-----+-----				Adj R-squared	=	0.0675
Total	146.53861	4,691	.031238246	Root MSE	=	.17067

...

```
. display ((0.0725-0.0681)/2/((1-0.0725)/(4692-6)))
```

11.11504

# Illustration 3 – Question 5

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(8, 4683)	=	49.76
Model	11.4798615	8	1.43498268	Prob > F	=	0.0000
Residual	135.058749	4,683	.02884022	R-squared	=	0.0783
-----+-----				Adj R-squared	=	0.0768
Total	146.53861	4,691	.031238246	Root MSE	=	.16982

  

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0301683	.0021456	14.06	0.000	.0259618	.0343747
COLLAT	.0407159	.0113232	3.60	0.000	.0185171	.0629146
PROF	-.1208179	.0278736	-4.33	0.000	-.1754634	-.0661725
GROWTH	2.28e-07	2.19e-06	0.10	0.917	-4.08e-06	4.53e-06
AGE	-.0002906	.0001691	-1.72	0.086	-.000622	.0000409
SE	-.0208001	.0068367	-3.04	0.002	-.0342032	-.0073969
MedE	-.028482	.0100863	-2.82	0.005	-.0482558	-.0087082
LE	-.0873071	.0163264	-5.35	0.000	-.1193145	-.0552997
_cons	-.3107104	.0268095	-11.59	0.000	-.3632698	-.2581511

# Illustration 3 – Question 5

```
. test SE MedE LE
```

```
( 1)  SE = 0
```

```
( 2)  MedE = 0
```

```
( 3)  LE = 0
```

```
F( 3, 4683) = 9.90  
Prob > F = 0.0000
```

The new regressors are jointly significant.

Ceteris paribus:

- a small firm, relative to a micro firm, displays in average a long term proportion of debt smaller in 0.021
- a medium firm, relative to a micro firm, displays in average a long term proportion of debt smaller in 0.028
- a large firm, relative to a micro firm, displays in average a long term proportion of debt smaller in 0.087



# Illustration 3 – Question 6.1

Model:

$$\begin{aligned} LEV_{LT} &= \beta_0 + \beta_1 SIZE + \beta_2 COLLAT + \beta_3 PROF + \beta_4 GROWTH + \beta_5 AGE \\ &+ \beta_6 (SE * PROF) + \beta_7 (MedE * PROF) + \beta_8 (LE * PROF) + v \end{aligned}$$

Effects of firm profitability:

- Micro firms:  $\beta_3$
- Small firms:  $\beta_3 + \beta_6$
- Medium firms:  $\beta_3 + \beta_7$
- Large firms:  $\beta_3 + \beta_8$

Null hypotheses:

	Small	Medium	Large
Micro	$\beta_6 = 0$	$\beta_7 = 0$	$\beta_8 = 0$
Small		$\beta_6 = \beta_7$	$\beta_6 = \beta_8$
Medium			$\beta_7 = \beta_8$

# Illustration 3 – Question 6.1

```

. gen PROF_SE=PROF*SE
. gen PROF_MedE=PROF*MedE
. gen PROF_LE=PROF*LE

. regress LEV_LT SIZE COLLAT PROF GROWTH AGE PROF_SE PROF_MedE PROF_LE

```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(8, 4683)	=	51.41
Model	11.8295674	8	1.47869593	Prob > F	=	0.0000
Residual	134.709043	4,683	.028765544	R-squared	=	0.0807
-----+-----				Adj R-squared	=	0.0792
Total	146.53861	4,691	.031238246	Root MSE	=	.1696

  

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0276292	.0015862	17.42	0.000	.0245194	.030739
COLLAT	.0363435	.0111628	3.26	0.001	.0144591	.0582279
PROF	.0095068	.0369042	0.26	0.797	-.0628428	.0818565
GROWTH	-2.68e-07	2.20e-06	-0.12	0.903	-4.57e-06	4.04e-06
AGE	-.0003792	.0001672	-2.27	0.023	-.000707	-.0000515
PROF_SE	-.2026651	.0517991	-3.91	0.000	-.3042157	-.1011145
PROF_MedE	-.343819	.0730633	-4.71	0.000	-.4870576	-.2005805
PROF_LE	-.7133022	.1360224	-5.24	0.000	-.9799701	-.4466343
_cons	-.2892868	.021412	-13.51	0.000	-.3312644	-.2473092

# Illustration 3 – Question 6.1

## Effects of firm profitability:

- When profitability increases 0.1 (10 pp for percentage), the proportion of long-term debt in the firm's capital structure decreases, on average:
  - 0.00095 (0.095 pp in percentage) - micro firms
  - -0.0193 - small firms →  $\text{display\_b[PROF]} + \text{b[PROF\_SE]}$
  - -0.0334 - medium firms →  $\text{display\_b[PROF]} + \text{b[PROF\_MedE]}$
  - -0.0704 - large firms →  $\text{display\_b[PROF]} + \text{b[PROF\_LE]}$

# Illustration 3 – Question 6.1

Profitability effects differ significantly across groups?

```
. test PROF_SE=PROF_MedE
( 1) PROF_SE - PROF_MedE = 0
      F( 1, 4683) = 3.84
      Prob > F = 0.0502
```

```
. test PROF_SE=PROF_LE
( 1) PROF_SE - PROF_LE = 0
      F( 1, 4683) = 14.66
      Prob > F = 0.0001
```

```
. test PROF_MedE=PROF_LE
( 1) PROF_MedE - PROF_LE = 0
      F( 1, 4683) = 7.44
      Prob > F = 0.0064
```

(p-values)	Small	Medium	Large
Micro	0.000***	0.000***	0.000***
Small		0.0502*	0.000***
Medium			0.006***

# Illustration 3 – Question 6.2

Model:

$$\begin{aligned} &LEV\_LT \\ &= \beta_0 + \beta_1 SIZE + \beta_2 COLLAT + \beta_3 PROF + \beta_4 GROWTH + \beta_5 AGE + \beta_6 LE \\ &+ \beta_7 (LE * SIZE) + \beta_8 (LE * COLLAT) + \beta_9 (LE * PROF) \\ &+ \beta_{10} (LE * GROWTH) + \beta_{11} (LE * AGE) + w \end{aligned}$$

Null hypotheses:

- $H_0: \beta_6 = \beta_7 = \beta_8 = \beta_9 = \beta_{10} = \beta_{11} = 0$  (no structural break)

```
. gen SIZE_LE=SIZE*LE
. gen COLLAT_LE=COLLAT*LE
. gen GROWTH_LE=GROWTH*LE
. gen AGE_LE=AGE*LE
```

# Illustration 3 – Question 6.2

```
regress LEV_LT SIZE COLLAT PROF GROWTH AGE LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE
```

Source	SS	df	MS	Number of obs	=	4,692
-----+-----				F(11, 4680)	=	36.14
Model	11.4734211	11	1.04303829	Prob > F	=	0.0000
Residual	135.065189	4,680	.028860083	R-squared	=	0.0783
-----+-----				Adj R-squared	=	0.0761
Total	146.53861	4,691	.031238246	Root MSE	=	.16988
-----						
LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE	.0262108	.0015979	16.40	0.000	.0230782	.0293433
COLLAT	.0298537	.0114426	2.61	0.009	.0074207	.0522867
PROF	-.1178161	.0283403	-4.16	0.000	-.1733765	-.0622557
GROWTH	4.15e-07	2.19e-06	0.19	0.850	-3.89e-06	4.72e-06
AGE	-.0003768	.0001822	-2.07	0.039	-.0007339	-.0000196
LE	-.0570063	.1411598	-0.40	0.686	-.3337461	.2197334
SIZE_LE	-.002168	.0082257	-0.26	0.792	-.0182943	.0139582
COLLAT_LE	.1055932	.053384	1.98	0.048	.0009355	.2102509
PROF_LE	-.2771129	.1595962	-1.74	0.083	-.5899966	.0357708
GROWTH_LE	.0010892	.0006263	1.74	0.082	-.0001386	.002317
AGE_LE	.0003647	.000469	0.78	0.437	-.0005547	.0012842
_cons	-.2680171	.0213989	-12.52	0.000	-.3099691	-.2260651
-----						

# Illustration 3 – Question 6.2

```
. test LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE
```

- ( 1) LE = 0
- ( 2) SIZE\_LE = 0
- ( 3) COLLAT\_LE = 0
- ( 4) PROF\_LE = 0
- ( 5) GROWTH\_LE = 0
- ( 6) AGE\_LE = 0

```
F( 6, 4680) = 4.91  
Prob > F = 0.0001
```

The null hypothesis is rejected, which implies that it is better to estimate separate models for each size-based group of firms or, equivalently, the model of the previous page, since:

(see the next page and compare with the interactions model)

# Illustration 3 – Question 6.2

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE if LE==0  
(...)
```

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE	.0262108	.0016006	16.38	0.000	.0230728	.0293487
COLLAT	.0298537	.011462	2.60	0.009	.0073825	.0523249
PROF	-.1178161	.0283882	-4.15	0.000	-.1734712	-.062161
GROWTH	4.15e-07	2.20e-06	0.19	0.850	-3.89e-06	4.72e-06
AGE	-.0003768	.0001825	-2.06	0.039	-.0007345	-.000019
_cons	-.2680171	.0214351	-12.50	0.000	-.3100405	-.2259936

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE if LE==1  
(...)
```

LEV_LT	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE	.0240427	.0078385	3.07	0.002	.008609	.0394765
COLLAT	.1354469	.0506539	2.67	0.008	.0357116	.2351822
PROF	-.394929	.1525738	-2.59	0.010	-.6953401	-.0945179
GROWTH	.0010896	.0006084	1.79	0.074	-.0001082	.0022875
AGE	-.000012	.0004198	-0.03	0.977	-.0008387	.0008146
_cons	-.3250234	.1355432	-2.40	0.017	-.591902	-.0581448



# Illustration 3 – Question 7.1 & 7.2

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE
```

```
. ovtest
```

```
Ramsey RESET test using powers of the fitted values of LEV_LT
```

```
Ho: model has no omitted variables
```

```
F(3, 4680) = 10.62
```

```
Prob > F = 0.0000 → Unsuitable model functional form
```

```
. estat hettest, rhs fstat
```

```
Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
```

```
Ho: Constant variance
```

```
Variables: SIZE COLLAT PROF GROWTH AGE SE MedE LE
```

```
F(8 , 4683) = 20.76
```

```
Prob > F = 0.0000 → Heteroskedastic errors
```

# Illustration 3 – Question 7.1 & 7.2

## RESET test – manual implementation

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE
. predict XB
. gen XB2=XB^2
. gen XB3=XB^3
. gen XB4=XB^4
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE XB2 XB3 XB4

. test XB2 XB3 XB4

( 1)  XB2 = 0
( 2)  XB3 = 0
( 3)  XB4 = 0

F( 3, 4680) = 10.62
Prob > F = 0.0000
```

# Illustration 3 – Question 7.1 & 7.2

## BP test – manual implementation

```
. quietly regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE

. predict uhat, resid

. gen uhat2=uhat^2

. quietly regress uhat2 SIZE COLLAT PROF GROWTH AGE SE MedE LE

. test SIZE COLLAT PROF GROWTH AGE SE MedE LE
( 1)  SIZE = 0
( 2)  COLLAT = 0
( 3)  PROF = 0
( 4)  GROWTH = 0
( 5)  AGE = 0
( 6)  SE = 0
( 7)  MedE = 0
( 8)  LE = 0

      F( 8, 4683) = 20.76
      Prob > F = 0.0000
```

# Illustration 3 – Question 8

```
. regress LEV_LT SIZE COLLAT PROF GROWTH AGE SE MedE LE, robust
```

```
Linear regression                Number of obs    =    4,692
                                F(8, 4683)        =    60.36
                                Prob > F            =    0.0000
                                R-squared           =    0.0783
                                Root MSE        =    .16982
```

---

LEV_LT	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE	.0301683	.0024476	12.33	0.000	.0253699	.0349667
COLLAT	.0407159	.0120525	3.38	0.001	.0170872	.0643445
PROF	-.1208179	.0214538	-5.63	0.000	-.1628775	-.0787583
GROWTH	2.28e-07	1.00e-06	0.23	0.821	-1.74e-06	2.20e-06
AGE	-.0002906	.0001641	-1.77	0.077	-.0006123	.0000312
SE	-.0208001	.0075488	-2.76	0.006	-.0355994	-.0060008
MedE	-.028482	.0118234	-2.41	0.016	-.0516614	-.0053026
LE	-.0873071	.0180288	-4.84	0.000	-.122652	-.0519622
_cons	-.3107104	.0273125	-11.38	0.000	-.3642558	-.257165

---

# Stata – Panel data

Commands for panel data:

- Prefix **xt**

Initial command:

- **xtset** *identvar tempvar*

Examples:

- **xtdescribe** [*varlist*]
- **xtsum** [*varlist*]
- **xttab** *varname*

# Illustration 4

Model:

$$\begin{aligned} & \text{Log}(Wage_{it}) \\ &= \beta_0 + \beta_1 \text{Schooling}_{it} + \beta_2 \text{Exper}_{it} + \beta_3 \text{Exper}^2_{it} + \beta_4 \text{Black}_i \\ &+ \beta_5 \text{Union}_{it} + \beta_6 \text{South}_{it} + \beta_7 \text{Public}_{it} + u \end{aligned}$$

Aim:

- Estimating  $\beta_5$  (impact of collective bargaining on wages)

Sample:

- 545 full-time working males who completed their schooling by 1980 and were observed over the period 1980-1987

Details:

- Verbeek (2008), pp. 375-377

# Illustration 4 – Question 1

```
. xtset NR YEAR
      panel variable:  NR (strongly balanced)
      time variable:  YEAR, 1980 to 1987
                  delta:  1 unit
```

```
. xtdescribe
      NR:  13, 17, ..., 12548                n =          545
      YEAR: 1980, 1981, ..., 1987           T =           8
      Delta(YEAR) = 1 unit
      Span(YEAR)  = 8 periods
      (NR*YEAR uniquely identifies each observation)
```

```
Distribution of T_i:  min      5%      25%      50%      75%      95%      max
                    8         8         8         8         8         8         8
```

```
      Freq.  Percent  Cum.  |  Pattern
-----+-----
      545    100.00  100.00 | 11111111
-----+-----
      545    100.00      |  XXXXXXXX
```

## • Balanced Panel

# Illustration 4 – Question 2

```
. xtsum WAGE SCHOOLING EXPER BLACK UNION SOUTH PUBLIC
```

Variable		Mean	Std. Dev.	Min	Max	Observations
WAGE	overall	5.919175	3.202225	.0279014	57.50431	N = 4360
	between		2.455819	1.503564	28.35696	n = 545
	within		2.057397	-16.42349	35.06652	T = 8
SCHOOL~G	overall	11.76697	1.746181	3	16	N = 4360
	between		1.747585	3	16	n = 545
	within		0	11.76697	11.76697	T = 8
EXPER	overall	6.514679	2.825873	0	18	N = 4360
	between		1.654918	3.5	14.5	n = 545
	within		2.291551	3.014679	10.01468	T = 8
BLACK	overall	.1155963	.3197769	0	1	N = 4360
	between		.320034	0	1	n = 545
	within		0	.1155963	.1155963	T = 8



# Illustration 4 – Question 2 (cont.)

```
. xtsum WAGE SCHOOLING EXPER BLACK UNION SOUTH PUBLIC
```

Variable		Mean	Std. Dev.	Min	Max	Observations
UNION	overall	.2440367	.4295639	0	1	N = 4360
	between		.3294467	0	1	n = 545
	within		.2759787	-.6309633	1.119037	T = 8
SOUTH	overall	.3506881	.4772402	0	1	N = 4360
	between		.4644838	0	1	n = 545
	within		.1111732	-.5243119	1.225688	T = 8
PUBLIC	overall	.0401376	.1963044	0	1	N = 4360
	between		.1388064	0	1	n = 545
	within		.1389214	-.8348624	.9151376	T = 8

# Illustration 4 – Question 3

```
. xttab UNION
```

UNION	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	3296	75.60	511	93.76	80.63
1	1064	24.40	280	51.38	47.50
Total	4360	100.00	791	145.14	68.90

(n = 545)

- Overall, 24,4% of wages resulted from collective bargaining
- For 51,38% of workers (280 out of 545), in at least one year their wages resulted from collective bargaining
- For 47,5% of those 280 workers, wages resulted always from collective bargaining

# Illustration 4 – Question 3 (cont.)

```
. xttab SOUTH
```

SOUTH	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	2831	64.93	374	68.62	94.62
1	1529	35.07	212	38.90	90.15
Total	4360	100.00	586	107.52	93.00

(n = 545)

- Overall, 35,1% of observations concern cases of individuals living in the South
- 38,9% of workers (212 out of 545) lived at least one year in the South
- 90,15% of those 212 workers lived always in the South

# Illustration 4 – Question 3 (cont.)

```
. xttab PUBLIC
```

PUBLIC	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	4185	95.99	543	99.63	96.34
1	175	4.01	62	11.38	35.28
Total	4360	100.00	605	111.01	90.08

(n = 545)

- Overall, 4,01 % of observations concern cases of individuals working in Public Administration
- 11,38% of workers (62 out of 545) worked at least one year in Public Administration
- 35,28% of those 62 workers worked always in Public Administration

# Illustration 4 – Question 4.1 - Pooled

```
. regress WAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, vce(cluster NR)
```

```
Linear regression                Number of obs    =    4,360
                                F(7, 544)        =    41.86
                                Prob > F              =    0.0000
                                R-squared              =    0.1574
                                Root MSE           =    2.9418
```

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
SCHOOLING	.6151076	.0607653	10.12	0.000	.4957443	.7344709
EXPER	.5637173	.062699	8.99	0.000	.4405556	.6868791
EXPER2	-.0190516	.0046374	-4.11	0.000	-.028161	-.0099423
BLACK	-.8307319	.2819856	-2.95	0.003	-1.384646	-.2768179
UNION	.9102328	.1801955	5.05	0.000	.5562686	1.264197
SOUTH	-.153659	.227973	-0.67	0.501	-.6014743	.2941563
PUBLIC	-.2967572	.2922075	-1.02	0.310	-.8707505	.2772361
_cons	-4.090845	.7595731	-5.39	0.000	-5.5829	-2.598789

```
. estimates store POOLED
```

# Illustration 4 – Question 4.1 – Random Effects

```
. xtreg WAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, vce(cluster NR)
```

Random-effects GLS regression

Group variable: NR

Number of obs = 4,360

Number of groups = 545

R-sq:

within = 0.1792

between = 0.1362

overall = 0.1536

Obs per group:

min = 8

avg = 8.0

max = 8

corr(u\_i, X) = 0 (assumed)

Wald chi2(7) = 444.91

Prob > chi2 = 0.0000

(continues in the next slide)

# Illustration 4 – Question 4.1 – Random Effects (cont.)

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
WAGE						
SCHOOLING	.628984	.0564469	11.14	0.000	.5183502	.7396178
EXPER	.6888584	.0675657	10.20	0.000	.5564322	.8212847
EXPER2	-.0252594	.0043403	-5.82	0.000	-.0337663	-.0167526
BLACK	-.8210685	.2853966	-2.88	0.004	-1.380435	-.2617014
UNION	.5444092	.1352284	4.03	0.000	.2793664	.809452
SOUTH	-.0413904	.2199264	-0.19	0.851	-.4724382	.3896574
PUBLIC	-.2358804	.255716	-0.92	0.356	-.7370746	.2653137
_cons	-4.710013	.7186157	-6.55	0.000	-6.118474	-3.301553
sigma_u	2.141408					
sigma_e	1.9932793					
rho	.53577993	(fraction of variance due to u_i)				

. estimates store RE

# Illustration 4 – Question 4.1 – Fixed Effects

```
. xtreg WAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, fe vce(cluster NR)
note: SCHOOLING omitted because of collinearity
note: BLACK omitted because of collinearity
```

```
Fixed-effects (within) regression
Group variable: NR
```

```
Number of obs      =      4,360
Number of groups   =       545
```

```
R-sq:
```

```
    within = 0.1796
    between = 0.0090
    overall = 0.0365
```

```
Obs per group:
```

```
    min =      8
    avg =     8.0
    max =      8
```

```
corr(u_i, Xb) = -0.1634
```

```
F(5,544) = 77.00
Prob > F = 0.0000
```

(continues in the next slide)



# Illustration 4 – Question 4.1 – Fixed Effects (cont.)

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
WAGE						
SCHOOLING	0	(omitted)				
EXPER	.7007947	.0708028	9.90	0.000	.5617143	.8398751
EXPER2	-.0257947	.0044667	-5.77	0.000	-.0345688	-.0170205
BLACK	0	(omitted)				
UNION	.4541499	.1456477	3.12	0.002	.1680491	.7402507
SOUTH	.1959752	.4142672	0.47	0.636	-.6177842	1.009734
PUBLIC	-.2332215	.2943376	-0.79	0.428	-.8113989	.344956
_cons	2.484219	.2793021	8.89	0.000	1.935577	3.032862
sigma_u	2.5677161					
sigma_e	1.9932793					
rho	.62397889	(fraction of variance due to u_i)				

. estimates store FE

# Illustration 4 – Question 4.1 – Table

. estimates table POOLED RE FE, b se

Variable	POOLED	RE	FE
SCHOOLING	.61510759	.628984	(omitted)
	.06076525	.05644686	
EXPER	.56371732	.68885842	.70079466
	.06269899	.06756565	.07080282
EXPER2	-.01905164	-.02525943	-.02579468
	.00463736	.00434032	.00446673
BLACK	-.83073193	-.82106847	(omitted)
	.28198564	.28539658	
UNION	.91023283	.54440916	.45414989
	.18019551	.13522841	.14564771
SOUTH	-.15365902	-.0413904	.19597517
	.22797304	.2199264	.41426722
PUBLIC	-.29675722	-.23588045	-.23322146
	.29220754	.25571599	.29433762
_cons	-4.0908448	-4.7100134	2.4842194
	.75957307	.71861566	.27930209

legend: b/se

# Illustration 4 – Question 4.1 – Table

```
. estimates table POOLED RE FE, b star(0.1 0.05 0.01)
```

Variable	POOLED	RE	FE
SCHOOLING	.61510759***	.628984***	(omitted)
EXPER	.56371732***	.68885842***	.70079466***
EXPER2	-.01905164***	-.02525943***	-.02579468***
BLACK	-.83073193***	-.82106847***	(omitted)
UNION	.91023283***	.54440916***	.45414989***
SOUTH	-.15365902	-.0413904	.19597517
PUBLIC	-.29675722	-.23588045	-.23322146
_cons	-4.0908448***	-4.7100134***	2.4842194***

legend: \* p<.1; \*\* p<.05; \*\*\* p<.01

# Illustration 4 – Question 4.2

```
. quietly xtreg WAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC, fe
. estimates store FEh
. quietly xtreg WAGE SCHOOLING EXPER EXPER2 BLACK UNION SOUTH PUBLIC
. estimates store REh
. hausman FEh REh
```

	---- Coefficients ----			
	(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
	FEh	REh	Difference	S.E.
EXPER	.7007947	.6888584	.0119362	.0071876
EXPER2	-.0257947	-.0252594	-.0005353	.0006201
UNION	.4541499	.5444092	-.0902593	.0371575
SOUTH	.1959752	-.0413904	.2373656	.2146316
PUBLIC	-.2332215	-.2358804	.002659	.0638298

b = consistent under Ho and Ha; obtained from xtreg  
 B = inconsistent under Ha, efficient under Ho; obtained from xtreg

Test: Ho: difference in coefficients not systematic

```
chi2(5) = (b-B)' [(V_b-V_B)^(-1)] (b-B)
          = 28.86
Prob>chi2 = 0.0000
```

→ The hypothesis of random effects is rejected.

# Illustration 4 – Question 4.3

```
. regress D.WAGE D.SCHOOLING D.EXPER D.EXPER2 D.BLACK D.UNION D.SOUTH D.PUBLIC, vce(cluster NR) nocons
```

note: D.SCHOOLING omitted because of collinearity

note: D.BLACK omitted because of collinearity

Linear regression	Number of obs	=	3,815
	F(5, 544)	=	64.45
	Prob > F	=	0.0000
	R-squared	=	0.0288
	Root MSE	=	2.3181

(continues in the next slide)

# Illustration 4 – Question 4.3

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
SCHOOLING						
D1.	0	(omitted)				
EXPER						
D1.	.6019491	.072804	8.27	0.000	.4589378	.7449605
EXPER2						
D1.	-.0175666	.0050963	-3.45	0.001	-.0275774	-.0075558
BLACK						
D1.	0	(omitted)				
UNION						
D1.	.292697	.1263287	2.32	0.021	.0445452	.5408488
SOUTH						
D1.	.0849769	.3899608	0.22	0.828	-.6810365	.8509903
PUBLIC						
D1.	-.0037897	.1961057	-0.02	0.985	-.3890068	.3814274

# Illustration 4 – Question 4.4

```
. xtreg LWAGE SCHOOLING EXPER EXPER2 c.BLACK##i.YEAR UNION SOUTH PUBLIC, re vce(cluster NR)
```

...

(Std. Err. adjusted for 545 clusters in NR)

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
SCHOOLING	.0928925	.0109548	8.48	0.000	.0714214	.1143636
EXPER	.1105721	.016444	6.72	0.000	.0783424	.1428018
EXPER2	-.0048039	.0007875	-6.10	0.000	-.0063473	-.0032605
BLACK	-.0888817	.0714913	-1.24	0.214	-.229002	.0512386
YEAR						
1981	.0396178	.0289286	1.37	0.171	-.0170812	.0963169
1982	.0461313	.0350288	1.32	0.188	-.0225238	.1147864
1983	.0321001	.0442559	0.73	0.468	-.0546398	.11884
1984	.0526727	.0565513	0.93	0.352	-.0581658	.1635111
1985	.0778807	.0654343	1.19	0.234	-.0503683	.2061296
1986	.0969979	.0753684	1.29	0.198	-.0507214	.2447171
1987	.1510222	.0856029	1.76	0.078	-.0167565	.3188008

(continues in the next slide)

# Illustration 4 – Question 4.4 (cont.)

```
YEAR#c.BLACK |
  1981 | .0272047 .0719625 0.38 0.705 -.1138393 .1682486
  1982 | -.1121196 .0918991 -1.22 0.222 -.2922386 .0679993
  1983 | -.0685286 .0846309 -0.81 0.418 -.2344021 .097345
  1984 | -.0541201 .0818181 -0.66 0.508 -.2144806 .1062405
  1985 | -.1522521 .0853343 -1.78 0.074 -.3195042 .0150001
  1986 | -.0376476 .0877731 -0.43 0.668 -.2096796 .1343844
  1987 | -.1414707 .0807227 -1.75 0.080 -.2996843 .016743
      |
  UNION | .1084207 .0208784 5.19 0.000 .0674999 .1493416
  SOUTH | -.0061861 .0317151 -0.20 0.845 -.0683466 .0559745
  PUBLIC | .0311455 .033775 0.92 0.356 -.0350523 .0973433
  _cons | .008449 .1567611 0.05 0.957 -.2987971 .3156951
-----+-----
sigma_u | .32792944
sigma_e | .3508362
rho | .46629073 (fraction of variance due to u_i)
-----+-----
```



# Illustration 5 – Model

- Trade-Off theory – optimal / target proportion of debt:

$$MDR_{it}^* = x'_{i,t-1}\beta + \eta_{it}$$

- Target adjustment model:

$$MDR_{it} - MDR_{i,t-1} = (1 - \gamma)(MDR_{it}^* - MDR_{i,t-1})$$

- $0 \leq \gamma \leq 1$
- Adjustment speed:  $1 - \gamma$ 
  - $\gamma = 0 \Rightarrow$  Firms adjust immediately and completely
  - $\gamma = 1 \Rightarrow$  No adjustment

- Econometric model:

$$MDR_{it} = MDR_{i,t-1} + (1 - \gamma)(x'_{i,t-1}\beta + \eta_{it} - MDR_{i,t-1})$$
$$\vdots$$
$$MDR_{it} = \gamma MDR_{i,t-1} + x'_{i,t-1}\theta + \alpha_i + u_{it}$$

# Illustration 5 – Model (cont.)

## Aim:

- Estimate  $\gamma$
- Test whether  $H_0: \gamma = 1$  (trade-off theory not valid)

## Sample:

- 5449 firms observed over the period 1986-2001
- Unbalanced panel
- Source: Compustat Industrial Annual Tapes

## Details:

- Verbeek (2008), pp. 383-388

```
. xtset gvkey yeara
      panel variable:  gvkey (unbalanced)
      time variable:   yeara, 1986 to 2001, but with gaps
                      delta: 1 unit
```

# Illustration 5 – Question 1

```
. xtdescribe
   gvkey: 1003, 1004, ..., 233397          n =          5449
   yeara: 1986, 1987, ..., 2001          T =           16
           Delta(yeara) = 1 unit
           Span(yeara)  = 16 periods
           (gvkey*yeara uniquely identifies each observation)
Distribution of T_i:  min      5%      25%      50%      75%      95%      max
                    1        1        1        3        8        14        16
   Freq.  Percent  Cum.  |  Pattern
-----+-----
   384    7.05    7.05 | .....1
   176    3.23   10.28 | 1111111111111111
   167    3.06   13.34 | 1.....
   149    2.73   16.08 | .....1.
   136    2.50   18.57 | 11.....
   132    2.42   20.99 | .....1...
   131    2.40   23.40 | .....1..
   113    2.07   25.47 | 111.....
    98    1.80   27.27 | .....1....
  3963   72.73  100.00 | (other patterns)
-----+-----
  5449  100.00          | XXXXXXXXXXXXXXXXXXXX
```

# Illustration 5 – Question 2

```
. quietly regress mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, vce(cluster gvkey)
```

```
. estimates store POOLED
```

```
. quietly xtreg mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, vce(cluster gvkey)
```

```
. estimates store RE
```

```
. quietly xtreg mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, fe vce(cluster gvkey)
```

```
. estimates store FE
```

# Illustration 5 – Question 2 (cont.)

```
. estimates table POOLED RE FE, b star
```

Variable	POOLED	RE	FE
mdr			
L1.	.8835036***	.78734111***	.53498254***
ebit_ta	-.03233775***	-.03455703***	-.05003294***
mb	.0016432*	.00069052	.00227756*
dep_ta	-.26051795***	-.30137705***	-.12395444
lnta	-.00067042	.00183328*	.03803015***
fa_ta	.02012146***	.02919991***	.05934357***
rd_dum	.00688957**	.0096552***	.00005977
rd_ta	-.12020508***	-.14520778***	-.06567621*
indmedian	.03212249**	.06029162***	.16721793***
rated	.00713406*	.01103821**	.02058981***
_cons	.05818177***	.03798629**	-.60083475***

legend: \* p<0.05; \*\* p<0.01; \*\*\* p<0.001

# Illustration 5 – Question 3.1

```
. xtivreg mdr (L.mdr=L2.mdr) ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta
indmedian rated, fd
(...)
```

D.mdr	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
LD.	7.03303	5.494343	1.28	0.201	-3.735684	17.80174
ebit_ta						
D1.	1.207597	.9705551	1.24	0.213	-.6946564	3.10985
mb						
D1.	.244267	.1853757	1.32	0.188	-.1190627	.6075966
dep_ta						
D1.	-1.858345	1.577202	-1.18	0.239	-4.949603	1.232914
lnta						
D1.	-.5214084	.4557998	-1.14	0.253	-1.41476	.3719429

(...)

```
Instrumented:  L.mdr
Instruments:  ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated
               L2.mdr
```

# Illustration 5 – Question 3.2

```
. xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep vce(robust)
```

```
Arellano-Bond dynamic panel-data estimation   Number of obs       =       15039  
Group variable: gvkey                         Number of groups    =        2996  
Time variable: yeara  
  
Obs per group:   min =           1  
                 avg =    5.019693  
                 max =           14  
  
Number of instruments =       115             Wald chi2(10)       =       212.72  
Prob > chi2        =       0.0000
```

Two-step results

(continues in the next slide)

# Illustration 5 – Question 3.2 (cont.)

(Std. Err. adjusted for clustering on gvkey)

	Coef.	WC-Robust Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
L1.	.3819695	.0731919	5.22	0.000	.238516	.525423
ebit_ta	.035684	.0173522	2.06	0.040	.0016744	.0696936
mb	.0147128	.0027217	5.41	0.000	.0093784	.0200472
dep_ta	.0648811	.109432	0.59	0.553	-.1496016	.2793639
lnta	.030107	.0083243	3.62	0.000	.0137916	.0464224
fa_ta	.0150317	.0286987	0.52	0.600	-.0412168	.0712801
rd_dum	-.0178784	.0122991	-1.45	0.146	-.0419841	.0062273
rd_ta	.001471	.035963	0.04	0.967	-.0690151	.0719571
indmedian	.0919917	.0453824	2.03	0.043	.0030439	.1809395
rated	-.0066174	.0099448	-0.67	0.506	-.0261088	.0128741
_cons	-.4235622	.1467658	-2.89	0.004	-.7112179	-.1359065

Instruments for differenced equation

GMM-type: L(2/.) .mdr

Standard: D.ebit\_ta D.mb D.dep\_ta D.lnta D.fa\_ta D.rd\_dum D.rd\_ta D.indmedian  
D.rated

Instruments for level equation

Standard: \_cons



# Illustration 5 – Question 3.3

```
. xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep maxldep(2) vce(robust)
```

```
Arellano-Bond dynamic panel-data estimation   Number of obs       =       15039  
Group variable: gvkey                         Number of groups    =        2996  
Time variable: yeara  
  
Obs per group:   min =           1  
                 avg =    5.019693  
                 max =           14  
  
Number of instruments =          37           Wald chi2(10)       =       179.25  
                                     Prob > chi2         =        0.0000
```

Two-step results

(continues in the next slide)

# Illustration 5 – Question 3.3 (cont.)

(Std. Err. adjusted for clustering on gvkey)

	Coef.	WC-Robust Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
L1.	.4005973	.0886448	4.52	0.000	.2268567	.574338
ebit_ta	.0445636	.0197946	2.25	0.024	.005767	.0833603
mb	.0165177	.003106	5.32	0.000	.0104301	.0226053
dep_ta	.0832717	.1094867	0.76	0.447	-.1313182	.2978616
lnta	.0252836	.0090865	2.78	0.005	.0074744	.0430928
fa_ta	.0065411	.0306739	0.21	0.831	-.0535787	.0666608
rd_dum	-.0135049	.0122092	-1.11	0.269	-.0374345	.0104248
rd_ta	.0015315	.0361632	0.04	0.966	-.0693471	.07241
indmedian	.0670826	.0493198	1.36	0.174	-.0295825	.1637477
rated	-.0098843	.0101263	-0.98	0.329	-.0297315	.009963
_cons	-.3367156	.1575827	-2.14	0.033	-.6455719	-.0278593

Instruments for differenced equation

GMM-type: L(2/3).mdr

Standard: D.ebit\_ta D.mb D.dep\_ta D.lnta D.fa\_ta D.rd\_dum D.rd\_ta  
D.indmedian D.rated

Instruments for level equation

Standard: \_cons

# Illustration 5 – Question 3.4

```
. xtdpdsys mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep vce(robust)
```

```
System dynamic panel-data estimation      Number of obs      =      19573  
Group variable: gvkey                    Number of groups   =      3777  
Time variable: yeara  
  
Obs per group:      min =           1  
                   avg =    5.182155  
                   max =           15  
  
Number of instruments =      129          Wald chi2(10)      =      2185.15  
                   Prob > chi2        =      0.0000  
  
Two-step results
```

(continues in the next slide)

# Illustration 5 – Question 3.4 (cont.)

```

-----
                |
                |          WC-Robust
mdr |          Coef.   Std. Err.   z   P>|z|   [95% Conf. Interval]
-----+-----
mdr |
L1. |    1.030104   .0238744   43.15   0.000   .9833112   1.076897
    |
ebit_ta |    .1285762   .018583   6.92   0.000   .0921541   .1649983
mb |    .0312589   .0027109   11.53   0.000   .0259457   .0365722
dep_ta |   -.1036924   .1374032   -0.75   0.450   -.3729978   .165613
lnta |   -.0127849   .0068955   -1.85   0.064   -.0262999   .0007301
fa_ta |   -.1085659   .0327925   -3.31   0.001   -.172838   -.0442939
rd_dum |   -.0061031   .0153821   -0.40   0.692   -.0362514   .0240452
rd_ta |    .0759739   .0441006   1.72   0.085   -.0104617   .1624096
indmedian |  -.2372958   .0400293   -5.93   0.000   -.3157518   -.1588397
rated |   -.0302876   .0120363   -2.52   0.012   -.0538783   -.006697
_cons |    .273204   .1269921   2.15   0.031   .0243041   .522104
-----

```

Instruments for differenced equation

GMM-type: L(2/.)mdr

Standard: D.ebit\_ta D.mb D.dep\_ta D.lnta D.fa\_ta D.rd\_dum D.rd\_ta D.indmedian  
D.rated

Instruments for level equation

GMM-type: LD.mdr

Standard: \_cons

# Illustration 5 – Question 4.1

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep vce(robust)
```

```
. estat abond, artests(3)
```

(...)

Arellano-Bond test for zero autocorrelation in first-differenced errors

```
+-----+  
|Order | z      Prob > z|  
+-----+-----+  
|  1  | -6.206  0.0000 |  
|  2  | -3.587  0.0003 |  
|  3  | -3.3325 0.0009 |  
+-----+  
H0: no autocorrelation
```

There is autocorrelation of order higher than 1: the estimators are not consistent.

# Illustration 5 – Question 4.2

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep
```

```
. estat sargan
```

Sargan test of overidentifying restrictions

H0: overidentifying restrictions are valid

```
chi2(104)      = 436.3939
```

```
Prob > chi2    = 0.0000
```

The hypothesis of exogenous  
instruments is rejected: the estimators  
are not consistent.

# Illustration 5 – Question 4.3

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep vce(robust)
```

```
. test L.mdr=1
```

```
( 1)  L.mdr = 1
```

```
      chi2( 1) =    71.30  
Prob > chi2 =    0.0000
```

The hypothesis that  $\gamma = 1$  is rejected, which implies that firms seem to adjust their debt ratios: there is no evidence against the trade-off theory.